



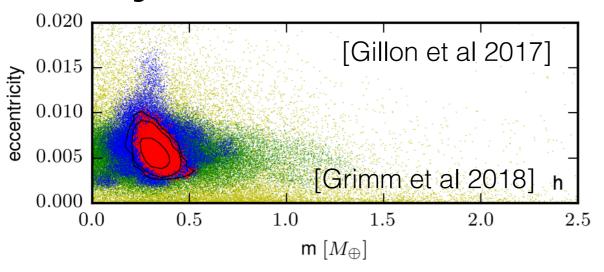


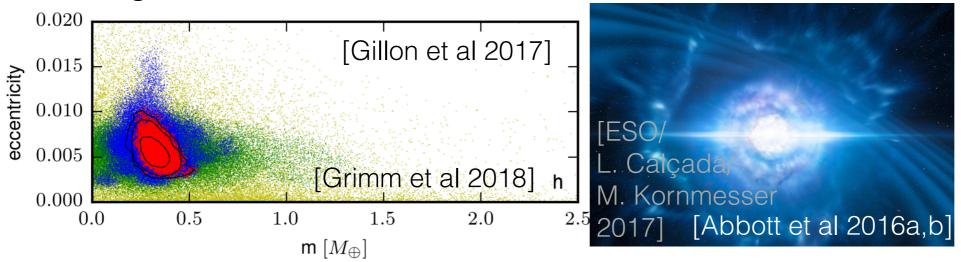
# Variational Bayes and beyond: Bayesian inference for big data

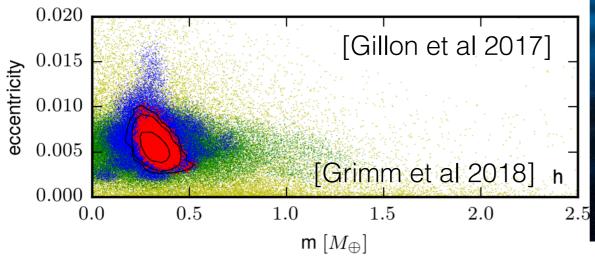
Tamara Broderick

Associate Professor,
Electrical Engineering & Computer Science
MIT

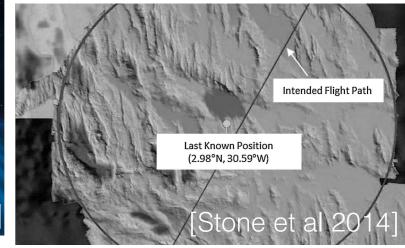
http://www.tamarabroderick.com/tutorials.html

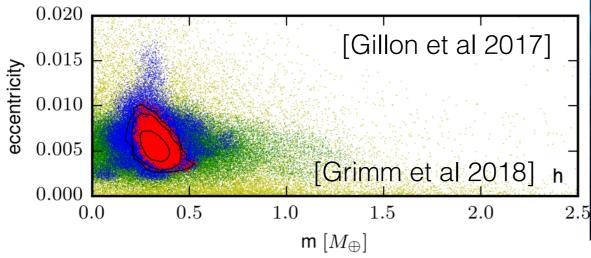




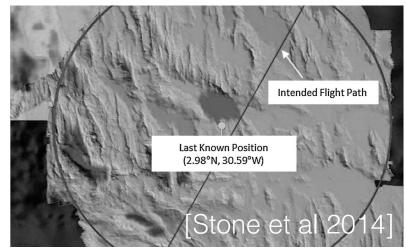


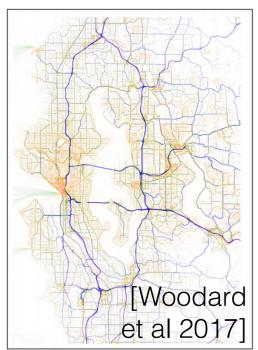


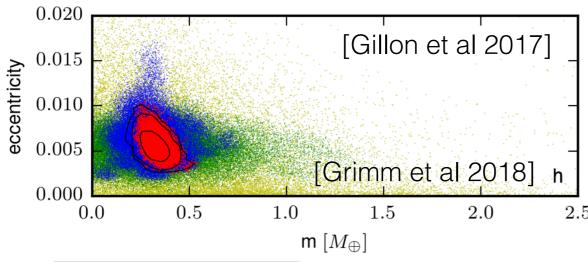




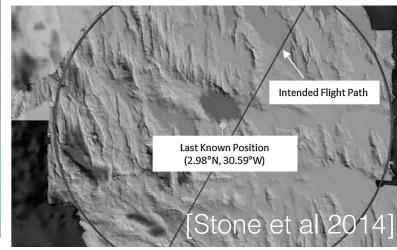


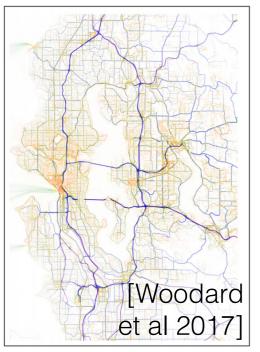




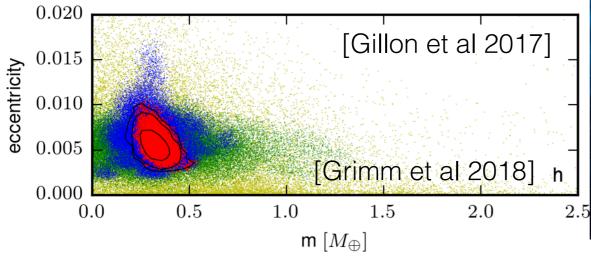




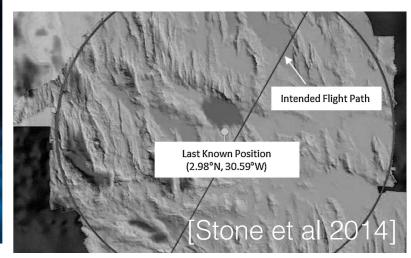


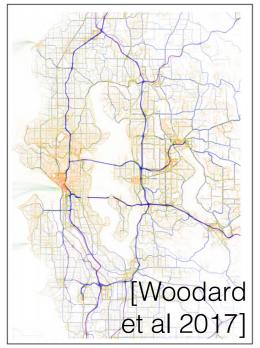


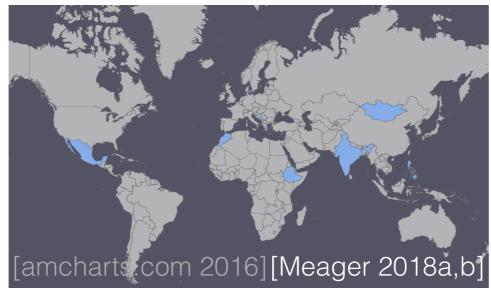




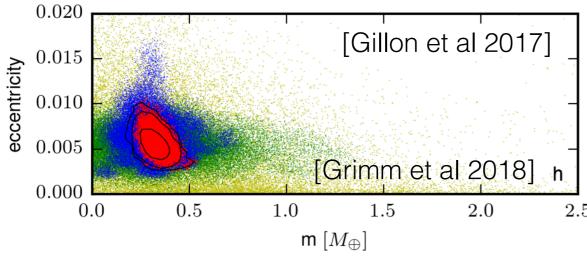




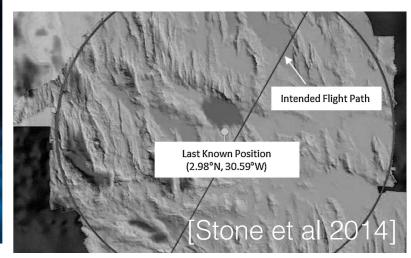


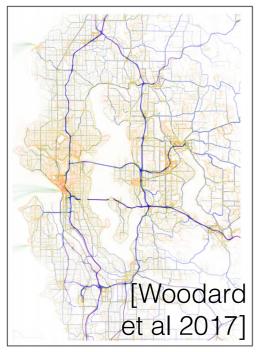








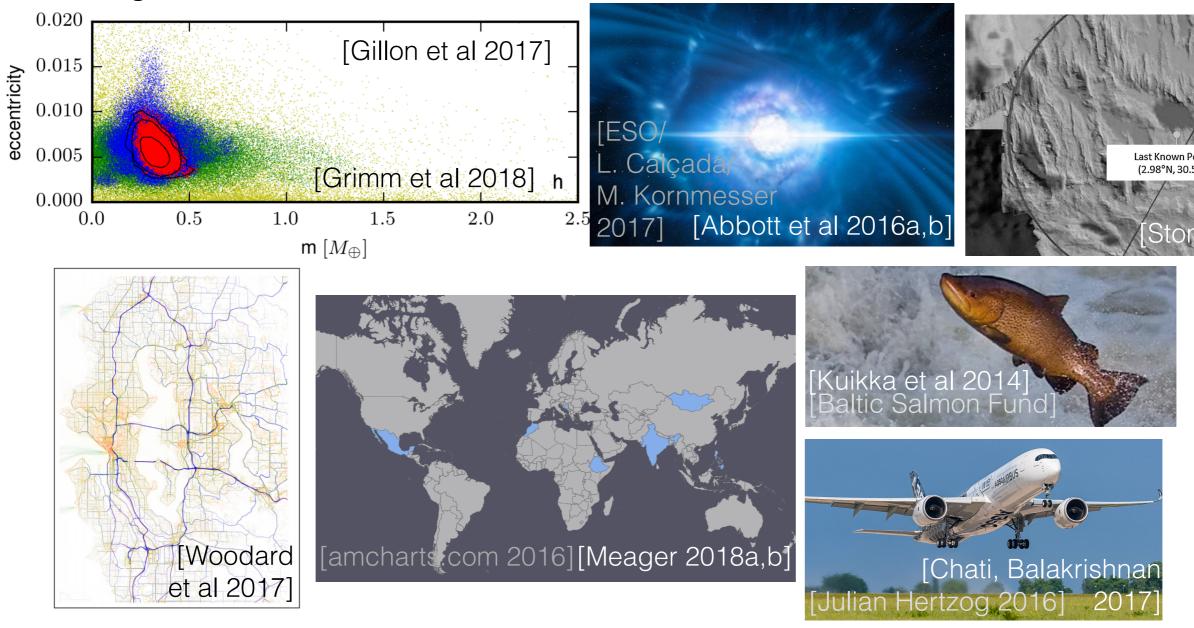






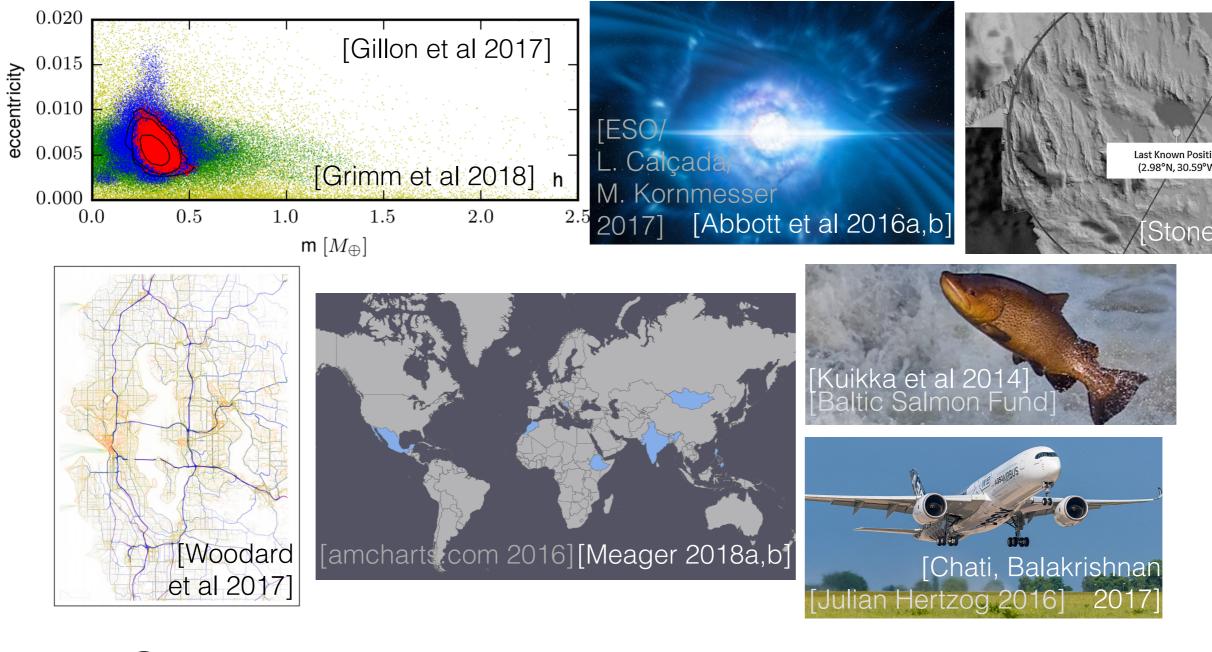






Intended Flight Path

Goals: good point estimates, uncertainty estimates



Intended Flight Path

- Goals: good point estimates, uncertainty estimates
  - More: interpretable, flexible, modular, expert info



- Goals: good point estimates, uncertainty estimates
  - More: interpretable, flexible, modular, expert info



- Goals: good point estimates, uncertainty estimates
  - More: interpretable, flexible, modular, expert info
- Challenge: speed (compute, user), reliable inference



- Goals: good point estimates, uncertainty estimates
  - More: interpretable, flexible, modular, expert info
- Challenge: speed (compute, user), reliable inference
- Uncertainty doesn't have to disappear in large data sets

• Modern problems: often large data, large dimensions

- Modern problems: often large data, large dimensions
- Variational Bayes can be very fast

- Modern problems: often large data, large dimensions
- Variational Bayes can be very fast

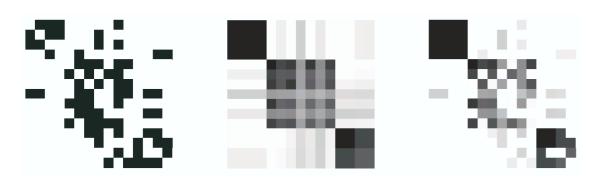
"Arts"	"Budgets"	"Children"	"Education"
"Arts"  NEW FILM SHOW MUSIC MOVIE PLAY MUSICAL BEST ACTOR FIRST YORK OPERA	"Budgets"  MILLION TAX PROGRAM BUDGET BILLION FEDERAL YEAR SPENDING NEW STATE PLAN MONEY	"Children"  CHILDREN WOMEN PEOPLE CHILD YEARS FAMILIES WORK PARENTS SAYS FAMILY WELFARE MEN	"Education"  SCHOOL [Blei et al STUDENTS SCHOOLS 2003] EDUCATION TEACHERS HIGH PUBLIC TEACHER BENNETT MANIGAT NAMPHY STATE
THEATER ACTRESS LOVE	PROGRAMS GOVERNMENT CONGRESS	PERCENT CARE LIFE	PRESIDENT ELEMENTARY HAITI

The William Randolph Hearst Foundation will give \$1.25 million to Lincoln Center, Metropolitan Opera Co., New York Philharmonic and Juilliard School. "Our board felt that we had a real opportunity to make a mark on the future of the performing arts with these grants an act every bit as important as our traditional areas of support in health, medical research, education and the social services," Hearst Foundation President Randolph A. Hearst said Monday in announcing the grants. Lincoln Center's share will be \$200,000 for its new building, which will house young artists and provide new public facilities. The Metropolitan Opera Co. and New York Philharmonic will receive \$400,000 each. The Juilliard School, where music and the performing arts are taught, will get \$250,000. The Hearst Foundation, a leading supporter of the Lincoln Center Consolidated Corporate Fund, will make its usual annual \$100,000 donation, too.

- Modern problems: often large data, large dimensions
- Variational Bayes can be very fast

"Arts"	"Budgets"	"Children"	"Education"
NEW	MILLION	CHILDREN	SCHOOL Blei et al STUDENTS SCHOOLS 2003
FILM	TAX	WOMEN	
SHOW	PROGRAM	PEOPLE	
MUSIC	BUDGET	CHILD	EDUCATION
MOVIE	BILLION	YEARS	TEACHERS
PLAY	FEDERAL	FAMILIES	HIGH PUBLIC TEACHER
MUSICAL	YEAR	WORK	
BEST	SPENDING	PARENTS	
ACTOR	NEW	SAYS	BENNETT
FIRST	STATE	FAMILY	MANIGAT
YORK	PLAN	WELFARE	NAMPHY
OPERA	MONEY	MEN	STATE
THEATER	PROGRAMS	PERCENT	PRESIDENT
ACTRESS	GOVERNMENT	CARE	ELEMENTARY
LOVE	CONGRESS	LIFE	HAITI

The William Randolph Hearst Foundation will give \$1.25 million to Lincoln Center, Metropolitan Opera Co., New York Philharmonic and Juilliard School. "Our board felt that we had a real opportunity to make a mark on the future of the performing arts with these grants an act every bit as important as our traditional areas of support in health, medical research, education and the social services," Hearst Foundation President Randolph A. Hearst said Monday in announcing the grants. Lincoln Center's share will be \$200,000 for its new building, which will house young artists and provide new public facilities. The Metropolitan Opera Co. and New York Philharmonic will receive \$400,000 each. The Juilliard School, where music and the performing arts are taught, will get \$250,000. The Hearst Foundation, a leading supporter of the Lincoln Center Consolidated Corporate Fund, will make its usual annual \$100,000 donation, too.

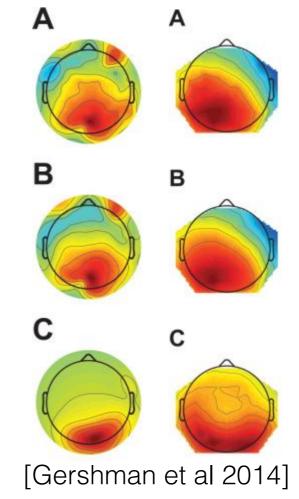


- Modern problems: often large data, large dimensions
- Variational Bayes can be very fast

"Arts"	"Budgets"	"Children"	"Education"
NEW	MILLION	CHILDREN	school Blei et al
$\operatorname{FILM}$	TAX	WOMEN	STUDENTS
SHOW	PROGRAM	PEOPLE	schools 2003
MUSIC	BUDGET	CHILD	EDUCATION
MOVIE	BILLION	YEARS	TEACHERS
PLAY	FEDERAL	FAMILIES	HIGH
MUSICAL	YEAR	WORK	PUBLIC
BEST	SPENDING	PARENTS	TEACHER
ACTOR	NEW	SAYS	BENNETT
FIRST	STATE	FAMILY	MANIGAT
YORK	PLAN	WELFARE	NAMPHY
OPERA	MONEY	MEN	STATE
THEATER	PROGRAMS	PERCENT	PRESIDENT
ACTRESS	GOVERNMENT	CARE	ELEMENTARY
LOVE	CONGRESS	LIFE	HAITI

The William Randolph Hearst Foundation will give \$1.25 million to Lincoln Center, Metropolitan Opera Co., New York Philharmonic and Juilliard School. "Our board felt that we had a real opportunity to make a mark on the future of the performing arts with these grants an act every bit as important as our traditional areas of support in health, medical research, education and the social services," Hearst Foundation President Randolph A. Hearst said Monday in announcing the grants. Lincoln Center's share will be \$200,000 for its new building, which will house young artists and provide new public facilities. The Metropolitan Opera Co. and New York Philharmonic will receive \$400,000 each. The Juilliard School, where music and the performing arts are taught, will get \$250,000. The Hearst Foundation, a leading supporter of the Lincoln Center Consolidated Corporate Fund, will make its usual annual \$100,000 donation, too.



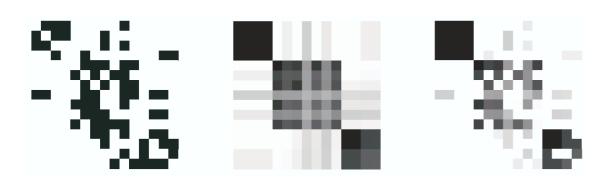


[Airoldi et al 2008]

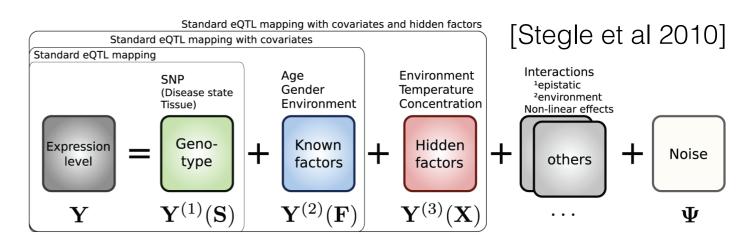
- Modern problems: often large data, large dimensions
- Variational Bayes can be very fast

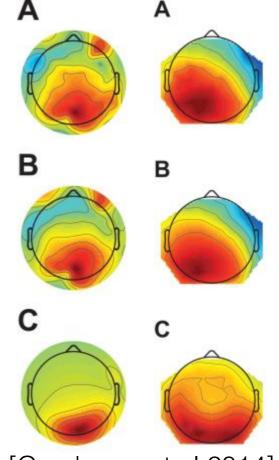
"Arts"	"Budgets"	"Children"	"Education"
			school [Blei et al
NEW	MILLION	CHILDREN	school [biei et ai
$\operatorname{FILM}$	TAX	WOMEN	STUDENTS
SHOW	PROGRAM	PEOPLE	schools 2003
MUSIC	BUDGET	CHILD	EDUCATION
MOVIE	BILLION	YEARS	TEACHERS
PLAY	FEDERAL	FAMILIES	HIGH
MUSICAL	YEAR	WORK	PUBLIC
BEST	SPENDING	PARENTS	TEACHER
ACTOR	NEW	SAYS	BENNETT
FIRST	STATE	FAMILY	MANIGAT
YORK	PLAN	WELFARE	NAMPHY
OPERA	MONEY	MEN	STATE
THEATER	PROGRAMS	PERCENT	PRESIDENT
ACTRESS	GOVERNMENT	CARE	ELEMENTARY
LOVE	CONGRESS	LIFE	HAITI

The William Randolph Hearst Foundation will give \$1.25 million to Lincoln Center, Metropolitan Opera Co., New York Philharmonic and Juilliard School. "Our board felt that we had a real opportunity to make a mark on the future of the performing arts with these grants an act every bit as important as our traditional areas of support in health, medical research, education and the social services," Hearst Foundation President Randolph A. Hearst said Monday in announcing the grants. Lincoln Center's share will be \$200,000 for its new building, which will house young artists and provide new public facilities. The Metropolitan Opera Co. and New York Philharmonic will receive \$400,000 each. The Juilliard School, where music and the performing arts are taught, will get \$250,000. The Hearst Foundation, a leading supporter of the Lincoln Center Consolidated Corporate Fund, will make its usual annual \$100,000 donation, too.



[Airoldi et al 2008]





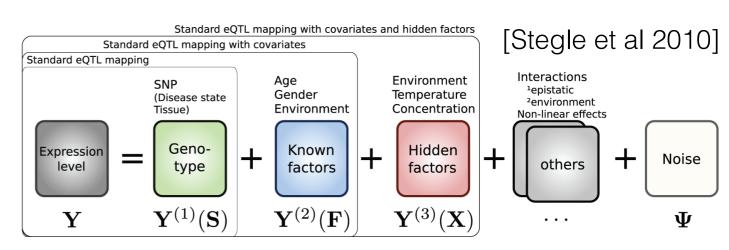
[Gershman et al 2014]

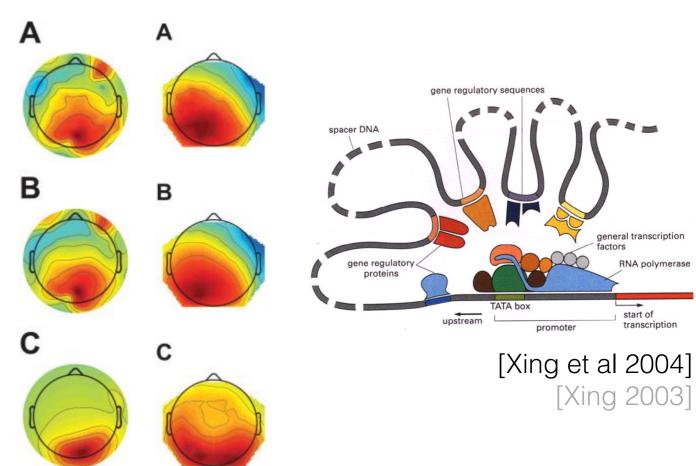
- Modern problems: often large data, large dimensions
- Variational Bayes can be very fast

"Arts"	"Budgets"	"Children"	"Education"
NEW	MILLION	CHILDREN	school   Blei et al
$\operatorname{FILM}$	TAX	WOMEN	STUDENTS
SHOW	PROGRAM	PEOPLE	SCHOOLS 2003
MUSIC	BUDGET	CHILD	EDUCATION
MOVIE	BILLION	YEARS	TEACHERS
PLAY	FEDERAL	FAMILIES	HIGH
MUSICAL	YEAR	WORK	PUBLIC
BEST	SPENDING	PARENTS	TEACHER
ACTOR	NEW	SAYS	BENNETT
FIRST	STATE	FAMILY	MANIGAT
YORK	PLAN	WELFARE	NAMPHY
OPERA	MONEY	MEN	STATE
THEATER	PROGRAMS	PERCENT	PRESIDENT
ACTRESS	GOVERNMENT	CARE	ELEMENTARY
LOVE	CONGRESS	LIFE	HAITI

The William Randolph Hearst Foundation will give \$1.25 million to Lincoln Center, Metropolitan Opera Co., New York Philharmonic and Juilliard School. "Our board felt that we had a real opportunity to make a mark on the future of the performing arts with these grants an act every bit as important as our traditional areas of support in health, medical research, education and the social services," Hearst Foundation President Randolph A. Hearst said Monday in announcing the grants. Lincoln Center's share will be \$200,000 for its new building, which will house young artists and provide new public facilities. The Metropolitan Opera Co. and New York Philharmonic will receive \$400,000 each. The Juilliard School, where music and the performing arts are taught, will get \$250,000. The Hearst Foundation, a leading supporter of the Lincoln Center Consolidated Corporate Fund, will make its usual annual \$100,000 donation, too.







[Gershman et al 2014]

general transcription

Bayes & Approximate Bayes review

- Bayes & Approximate Bayes review
- What is:
  - Variational Bayes (VB)
  - Mean-field variational Bayes (MFVB)

- Bayes & Approximate Bayes review
- What is:
  - Variational Bayes (VB)
  - Mean-field variational Bayes (MFVB)
- Why use MFVB?

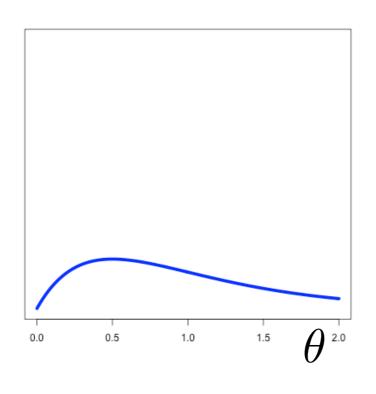
- Bayes & Approximate Bayes review
- What is:
  - Variational Bayes (VB)
  - Mean-field variational Bayes (MFVB)
- Why use MFVB?
- When can we trust MFVB?

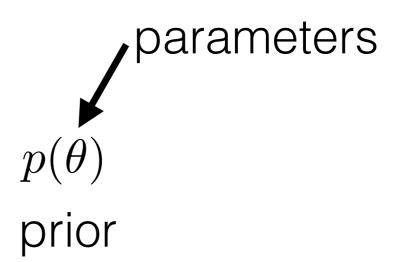
- Bayes & Approximate Bayes review
- What is:
  - Variational Bayes (VB)
  - Mean-field variational Bayes (MFVB)
- Why use MFVB?
- When can we trust MFVB?
- Where do we go from here?

- Bayes & Approximate Bayes review
- What is:
  - Variational Bayes (VB)
  - Mean-field variational Bayes (MFVB)
- Why use MFVB?
- When can we trust MFVB?
- Where do we go from here?



 $parameters \\ p(\theta)$  prior

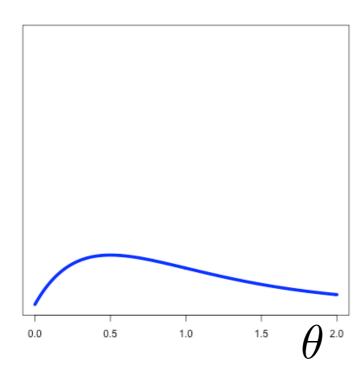




parameters

$$p(y_{1:N}|\theta)p(\theta)$$

likelihood prior

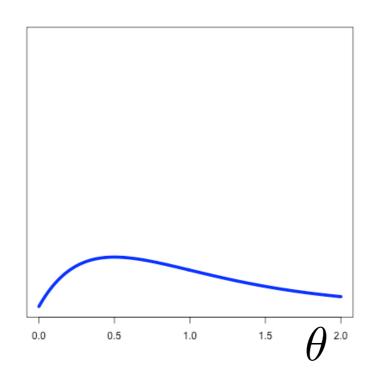


# Bayesian inference 1 data

parameters

$$p(y_{1:N}|\theta)p(\theta)$$

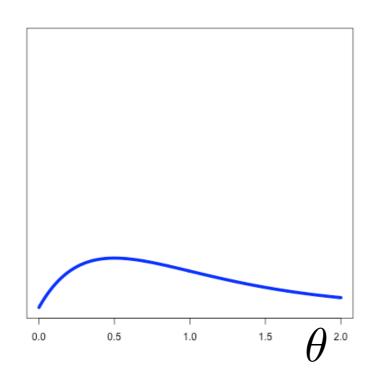
likelihood prior



# Bayesian inference / data / parameters

$$p(\theta|y_{1:N}) \propto_{\theta} p(y_{1:N}|\theta)p(\theta)$$

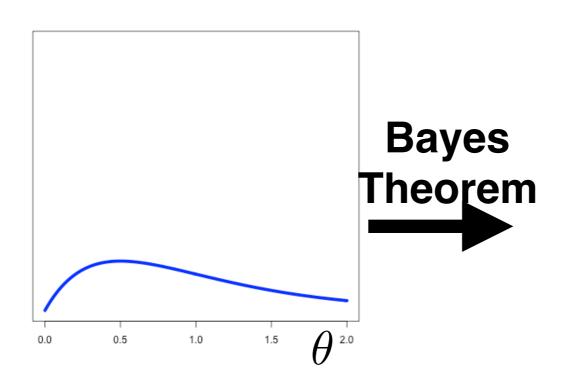
posterior likelihood prior



# Bayesian inference / data / parameters

 $p(\theta|y_{1:N}) \propto_{\theta} p(y_{1:N}|\theta)p(\theta)$ 

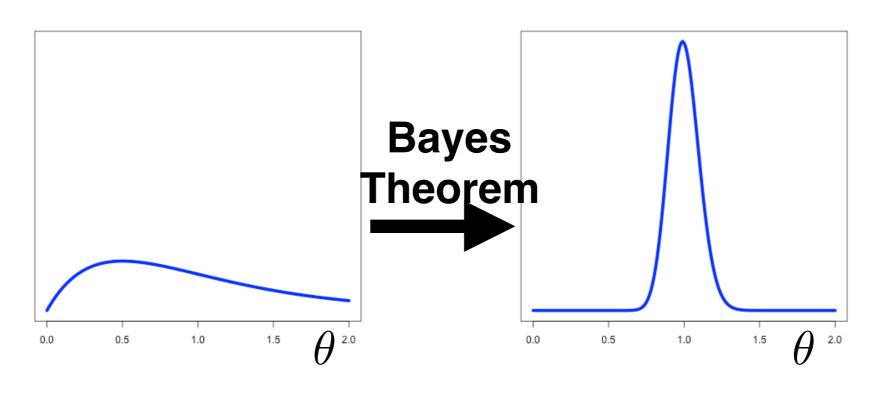
posterior likelihood prior



 $p(\theta|y_{1:N}) \propto_{\theta} p(y_{1:N}|\theta)p(\theta)$ 

posterior likelihood prior

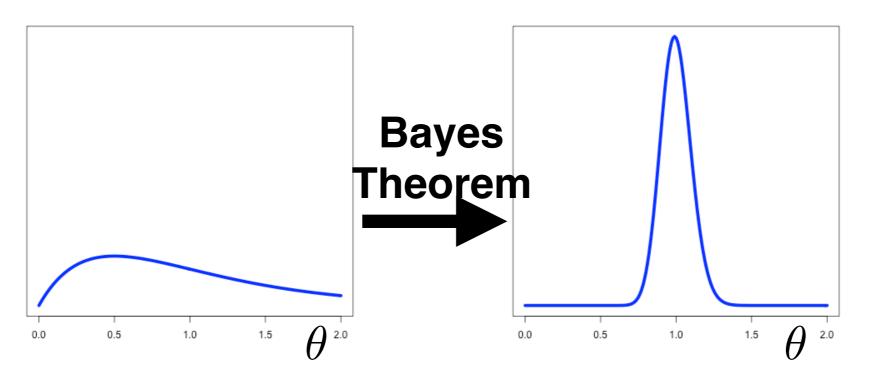
, parameters



 $p(\theta|y_{1:N}) \propto_{\theta} p(y_{1:N}|\theta)p(\theta)$ 

, parameters

posterior likelihood prior

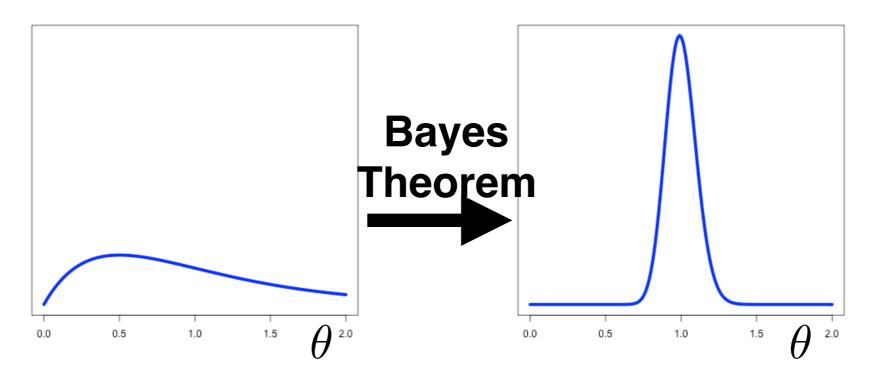


1. Build a model: choose prior & choose likelihood

$$p(\theta|y_{1:N}) \propto_{\theta} p(y_{1:N}|\theta)p(\theta)$$

parameters

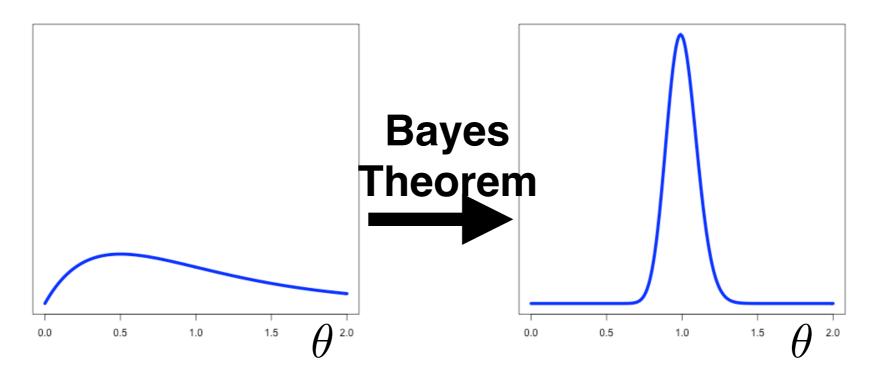
posterior likelihood prior



- 1. Build a model: choose prior & choose likelihood
- 2. Compute the posterior

# Bayesian inference ydata ypara

$$p(\theta|y_{1:N}) \propto_{\theta} p(y_{1:N}|\theta)p(\theta)$$
  
posterior likelihood prior

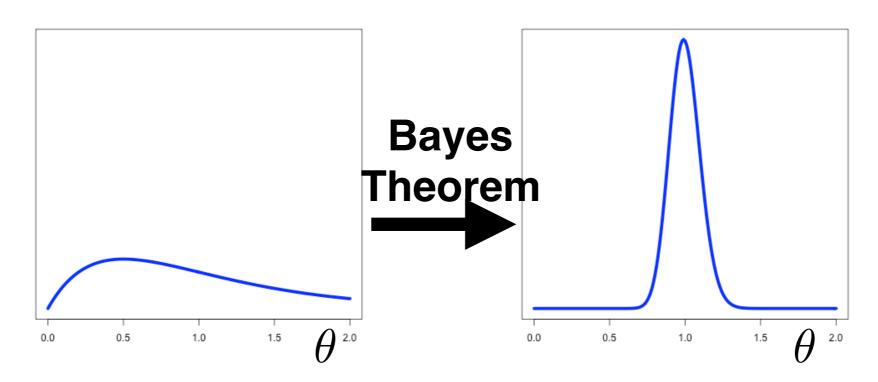


- 1. Build a model: choose prior & choose likelihood
- 2. Compute the posterior
- 3. Report a summary, e.g. posterior means and (co)variances

parameters

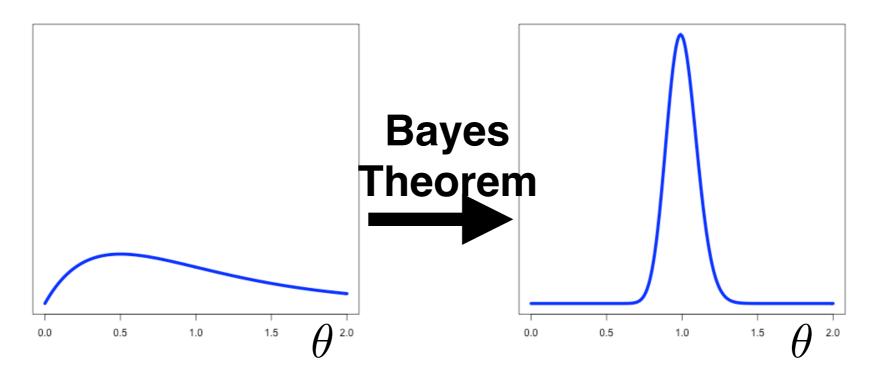
$$p(\theta|y_{1:N}) \propto_{\theta} p(y_{1:N}|\theta)p(\theta)$$

posterior likelihood prior



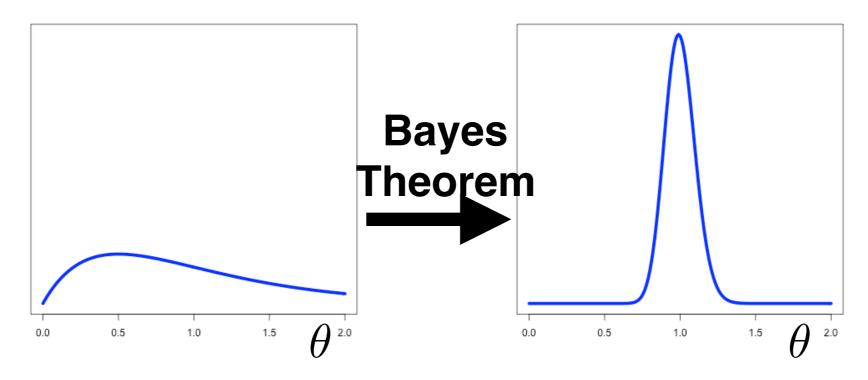
- 1. Build a model: choose prior & choose likelihood
- 2. Compute the posterior
- 3. Report a summary, e.g. posterior means and (co)variances
- Why are steps 2 and 3 hard?

 $p(\theta|y_{1:N}) \propto_{\theta} p(y_{1:N}|\theta)p(\theta)$ posterior likelihood prior



- 1. Build a model: choose prior & choose likelihood
- 2. Compute the posterior
- 3. Report a summary, e.g. posterior means and (co)variances
- Why are steps 2 and 3 hard?
  - Typically no closed form

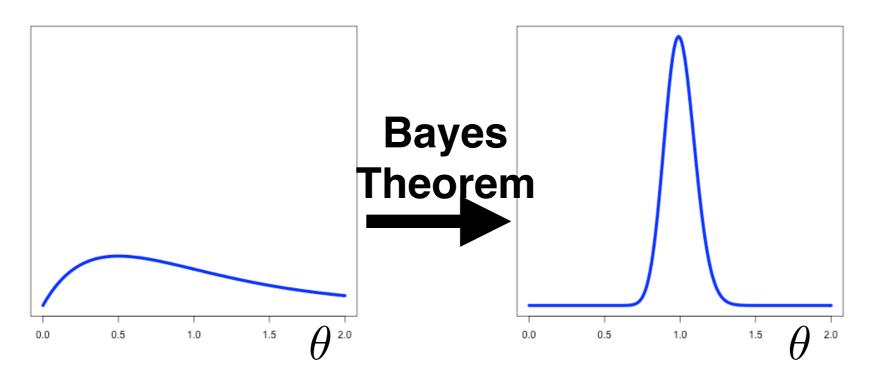
 $p(\theta|y_{1:N}) \propto_{\theta} p(y_{1:N}|\theta)p(\theta)$ posterior likelihood prior



- 1. Build a model: choose prior & choose likelihood
- 2. Compute the posterior
- 3. Report a summary, e.g. posterior means and (co)variances
- Why are steps 2 and 3 hard?
  - Typically no closed form, high-dimensional integration

## Bayesian inference Jata Jarameters

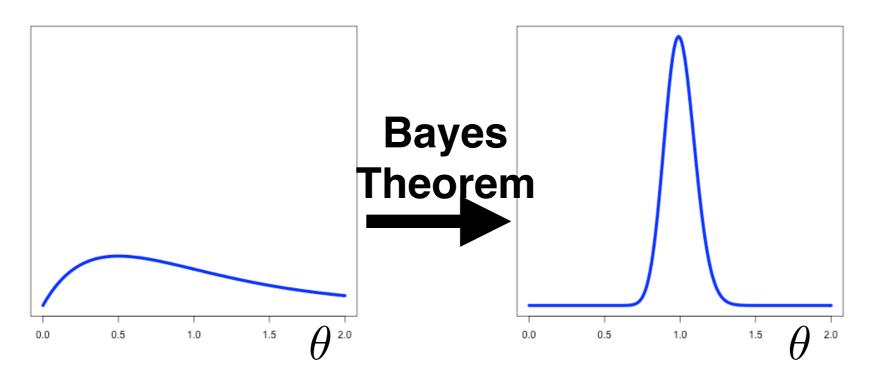
$$p(\theta|y_{1:N}) = p(y_{1:N}|\theta)p(\theta)/p(y_{1:N})$$
  
posterior likelihood prior



- 1. Build a model: choose prior & choose likelihood
- 2. Compute the posterior
- 3. Report a summary, e.g. posterior means and (co)variances
- Why are steps 2 and 3 hard?
  - Typically no closed form, high-dimensional integration

# Bayesian inference Jata Jarameters

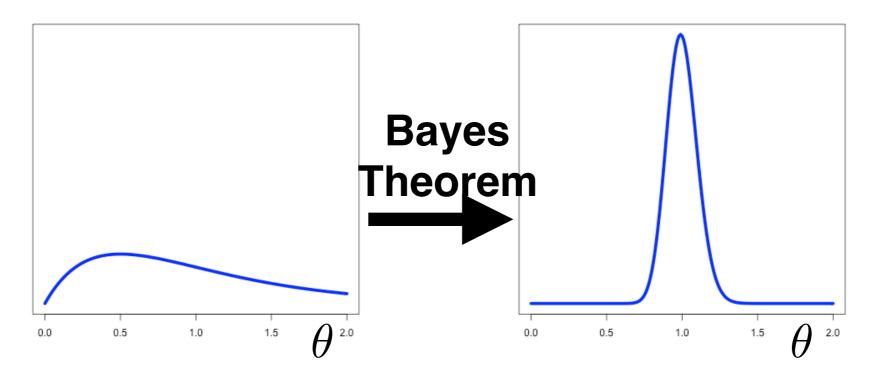
$$p(\theta|y_{1:N}) = p(y_{1:N}|\theta)p(\theta)/p(y_{1:N})$$
  
posterior likelihood prior evidence



- 1. Build a model: choose prior & choose likelihood
- 2. Compute the posterior
- 3. Report a summary, e.g. posterior means and (co)variances
- Why are steps 2 and 3 hard?
  - Typically no closed form, high-dimensional integration

# Bayesian inference 1 data 1 parameters

$$p(\theta|y_{1:N}) = p(y_{1:N}|\theta)p(\theta)/\int p(y_{1:N},\theta)d\theta$$
 posterior likelihood prior evidence



- 1. Build a model: choose prior & choose likelihood
- 2. Compute the posterior
- 3. Report a summary, e.g. posterior means and (co)variances
- Why are steps 2 and 3 hard?
  - Typically no closed form, high-dimensional integration

Gold standard: Markov Chain Monte Carlo (MCMC)

[Bardenet, Doucet, Holmes 2017]

- Gold standard: Markov Chain Monte Carlo (MCMC)
  - Eventually accurate but can be slow

[Bardenet, Doucet, Holmes 2017]

- Gold standard: Markov Chain Monte Carlo (MCMC)
- [Bardenet, Doucet, Holmes 2017]

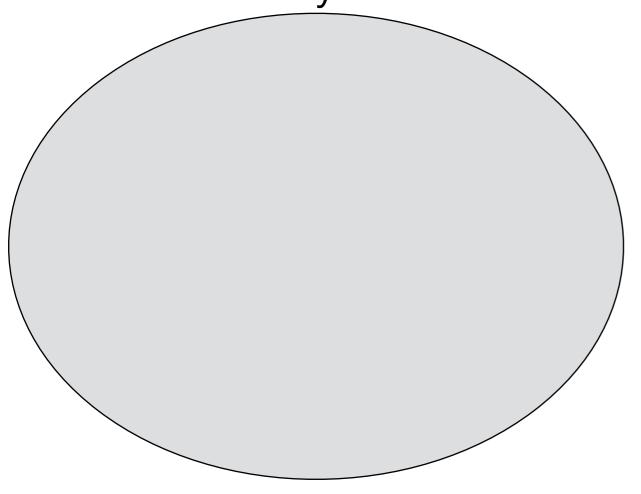
Eventually accurate but can be slow

Instead: an optimization approach

Gold standard: Markov Chain Monte Carlo (MCMC)

[Bardenet, Doucet, Holmes 2017]

Eventually accurate but can be slow

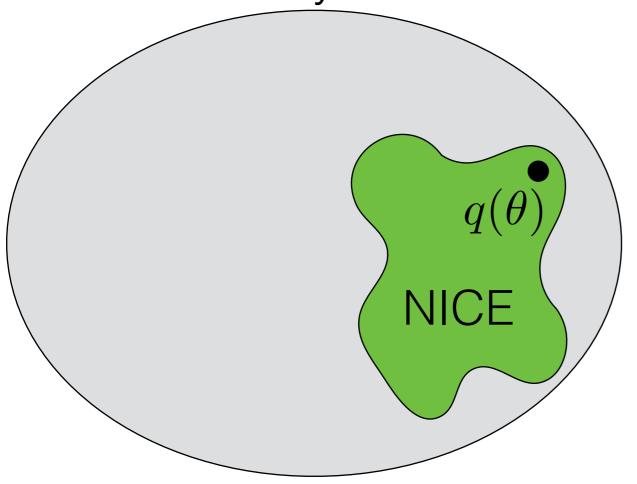


Instead: an optimization approach

Gold standard: Markov Chain Monte Carlo (MCMC)

[Bardenet, Doucet, Holmes 2017]

Eventually accurate but can be slow

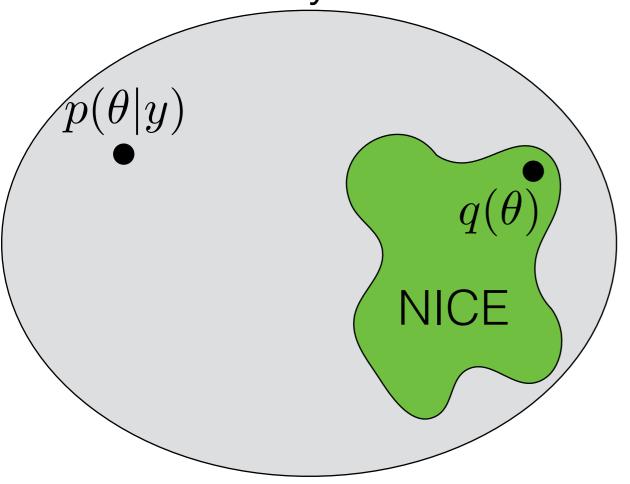


Instead: an optimization approach

Gold standard: Markov Chain Monte Carlo (MCMC)

[Bardenet, Doucet, Holmes 2017]

Eventually accurate but can be slow

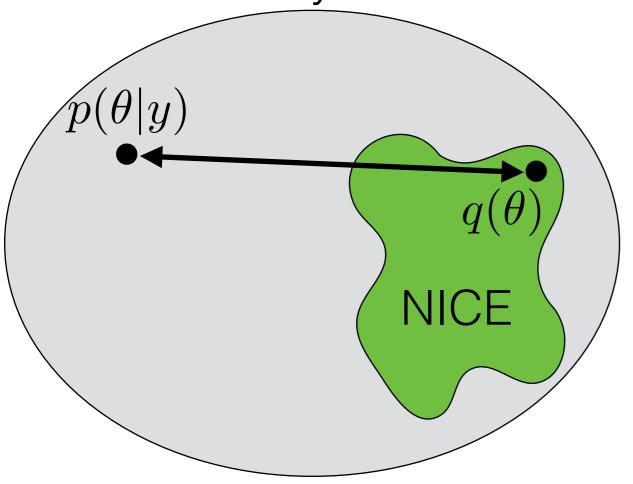


Instead: an optimization approach

Gold standard: Markov Chain Monte Carlo (MCMC)

[Bardenet, Doucet, Holmes 2017]

Eventually accurate but can be slow

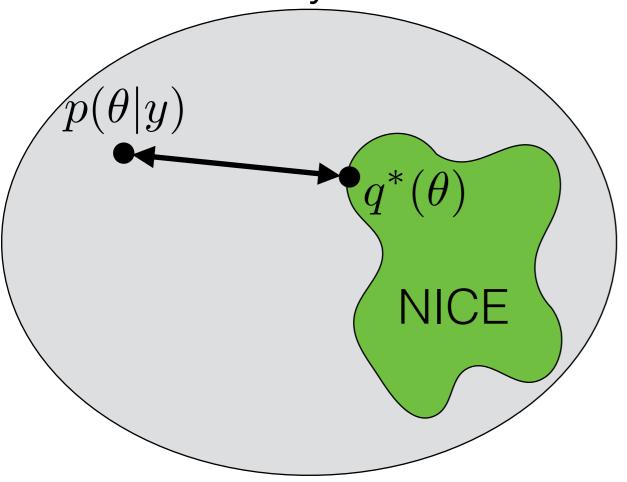


Instead: an optimization approach

Gold standard: Markov Chain Monte Carlo (MCMC)

[Bardenet, Doucet, Holmes 2017]

Eventually accurate but can be slow

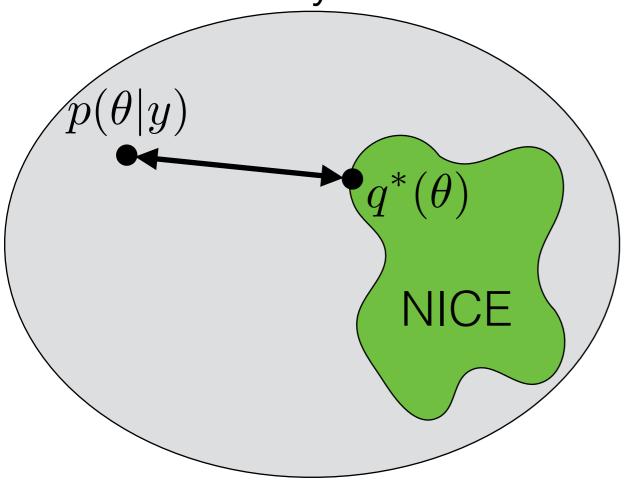


Instead: an optimization approach

Gold standard: Markov Chain Monte Carlo (MCMC)

[Bardenet, Doucet, Holmes 2017]

Eventually accurate but can be slow



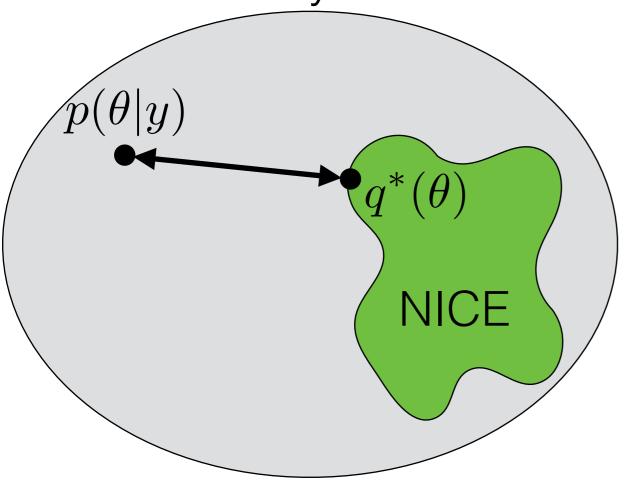
Instead: an optimization approach

$$q^* = \operatorname{argmin}_{q \in Q} f(q(\cdot), p(\cdot|y))$$

Gold standard: Markov Chain Monte Carlo (MCMC)

[Bardenet, Doucet, Holmes 2017]

Eventually accurate but can be slow



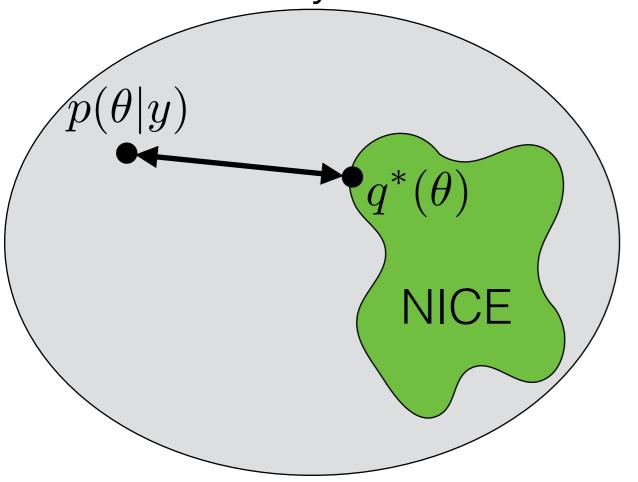
Instead: an optimization approach

$$q^* = \operatorname{argmin}_{q \in \mathbf{Q}} f(q(\cdot), p(\cdot|y))$$

Gold standard: Markov Chain Monte Carlo (MCMC)

[Bardenet, Doucet, Holmes 2017]

Eventually accurate but can be slow



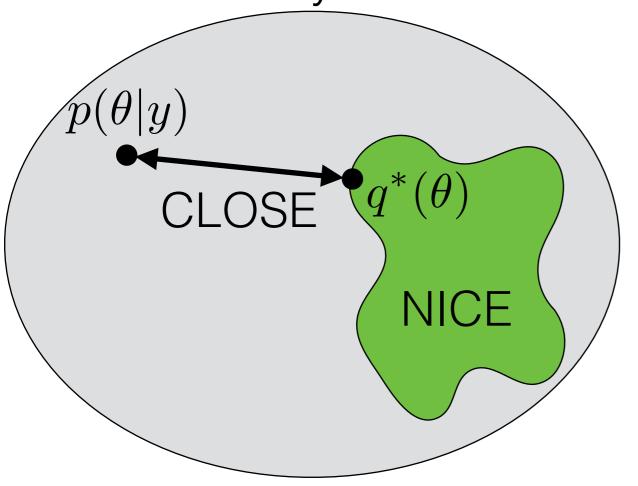
Instead: an optimization approach

$$q^* = \operatorname{argmin}_{q \in Q} f(q(\cdot), p(\cdot|y))$$

Gold standard: Markov Chain Monte Carlo (MCMC)

[Bardenet, Doucet, Holmes 2017]

Eventually accurate but can be slow



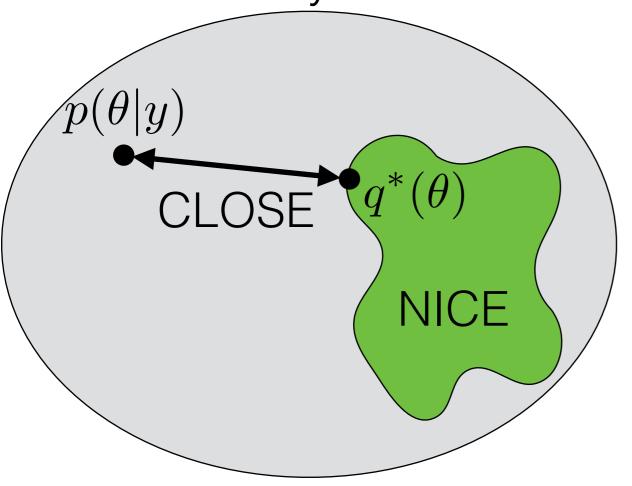
Instead: an optimization approach

$$q^* = \operatorname{argmin}_{q \in Q} f(q(\cdot), p(\cdot|y))$$

Gold standard: Markov Chain Monte Carlo (MCMC)

[Bardenet, Doucet, Holmes 2017]

Eventually accurate but can be slow



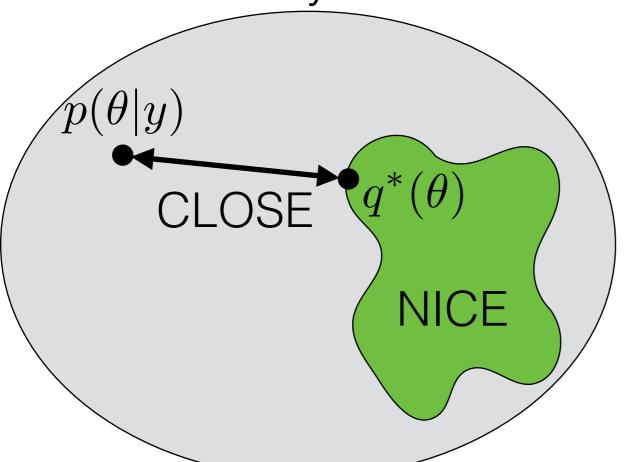
Instead: an optimization approach

$$q^* = \operatorname{argmin}_{q \in Q} f(q(\cdot), p(\cdot|y))$$

Gold standard: Markov Chain Monte Carlo (MCMC)

[Bardenet, Doucet, Holmes 2017]

Eventually accurate but can be slow



Instead: an optimization approach

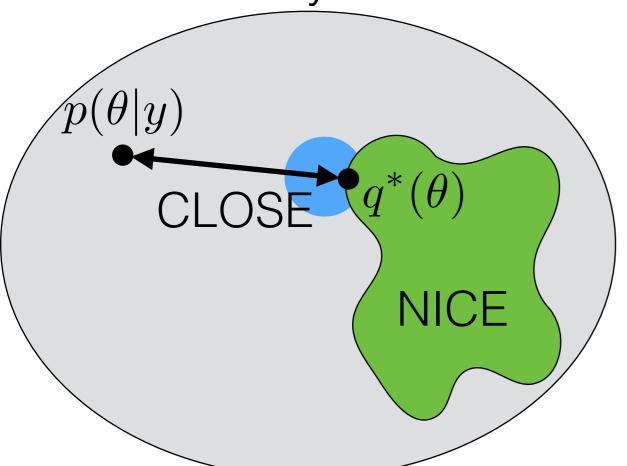
Approximate posterior with q\*

$$q^* = \operatorname{argmin}_{q \in Q} f(q(\cdot), p(\cdot|y))$$

Gold standard: Markov Chain Monte Carlo (MCMC)

[Bardenet, Doucet, Holmes 2017]

Eventually accurate but can be slow



Instead: an optimization approach

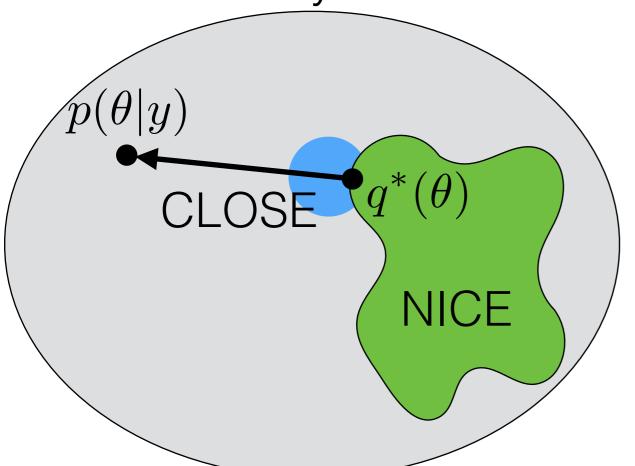
Approximate posterior with q\*

$$q^* = \operatorname{argmin}_{q \in Q} f(q(\cdot), p(\cdot|y))$$

Gold standard: Markov Chain Monte Carlo (MCMC)

[Bardenet, Doucet, Holmes 2017]

Eventually accurate but can be slow



Instead: an optimization approach

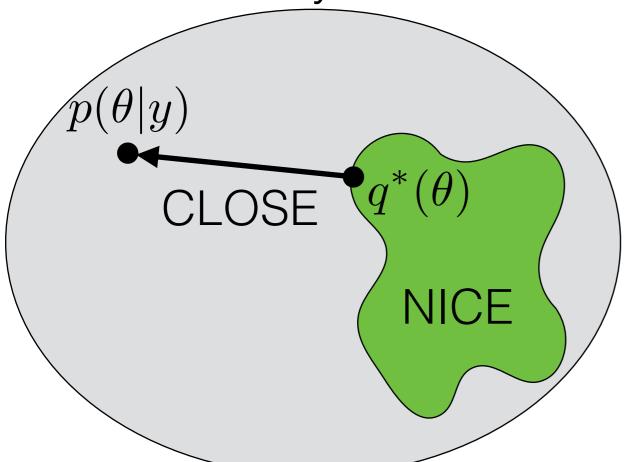
Approximate posterior with q\*

$$q^* = \operatorname{argmin}_{q \in Q} f(q(\cdot), p(\cdot|y))$$

Gold standard: Markov Chain Monte Carlo (MCMC)

[Bardenet, Doucet, Holmes 2017]

Eventually accurate but can be slow



Instead: an optimization approach

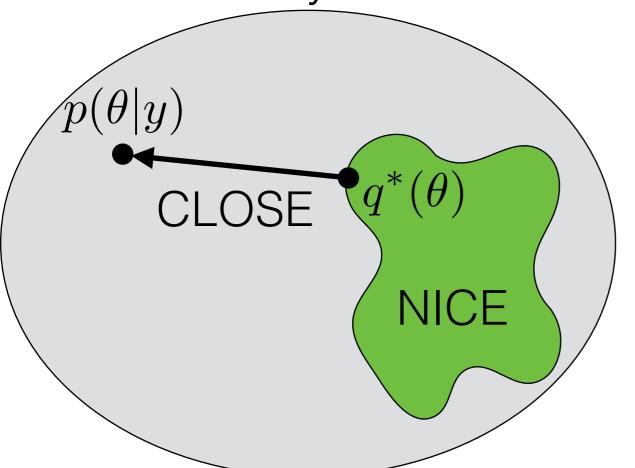
Approximate posterior with q\*

$$q^* = \operatorname{argmin}_{q \in Q} f(q(\cdot), p(\cdot|y))$$

Gold standard: Markov Chain Monte Carlo (MCMC)

[Bardenet, Doucet, Holmes 2017]

Eventually accurate but can be slow



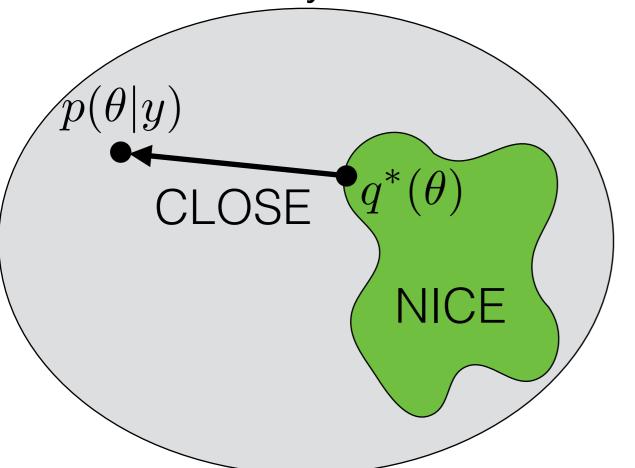
Instead: an optimization approach

$$q^* = \operatorname{argmin}_{q \in Q} f(q(\cdot), p(\cdot|y))$$

- Variational Bayes (VB): f is Kullback-Leibler divergence  $KL(q(\cdot)||p(\cdot|y))$
- VB practical success

- Gold standard: Markov Chain Monte Carlo (MCMC)
- [Bardenet, Doucet, Holmes 2017]

Eventually accurate but can be slow



Instead: an optimization approach

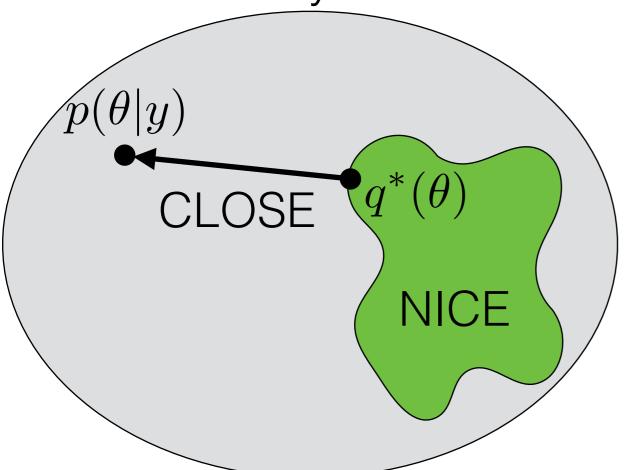
$$q^* = \operatorname{argmin}_{q \in Q} f(q(\cdot), p(\cdot|y))$$

- Variational Bayes (VB): f is Kullback-Leibler divergence  $KL(q(\cdot)||p(\cdot|y))$
- VB practical success: point estimates and prediction

Gold standard: Markov Chain Monte Carlo (MCMC)

[Bardenet, Doucet, Holmes 2017]

Eventually accurate but can be slow



Instead: an optimization approach

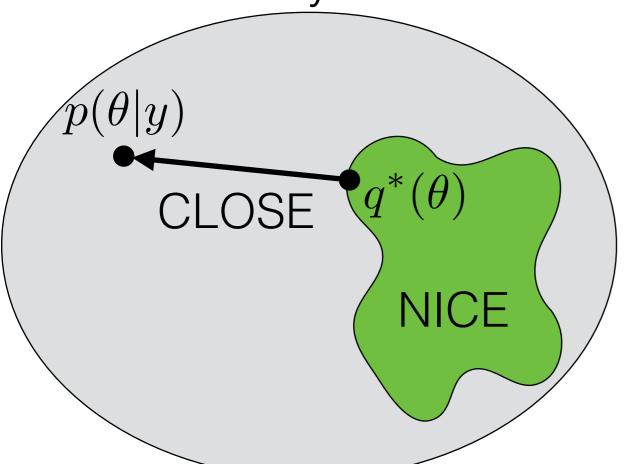
$$q^* = \operatorname{argmin}_{q \in Q} f(q(\cdot), p(\cdot|y))$$

- Variational Bayes (VB): f is Kullback-Leibler divergence  $KL(q(\cdot)||p(\cdot|y))$
- VB practical success: point estimates and prediction, fast

Gold standard: Markov Chain Monte Carlo (MCMC)

[Bardenet, Doucet, Holmes 2017]

Eventually accurate but can be slow

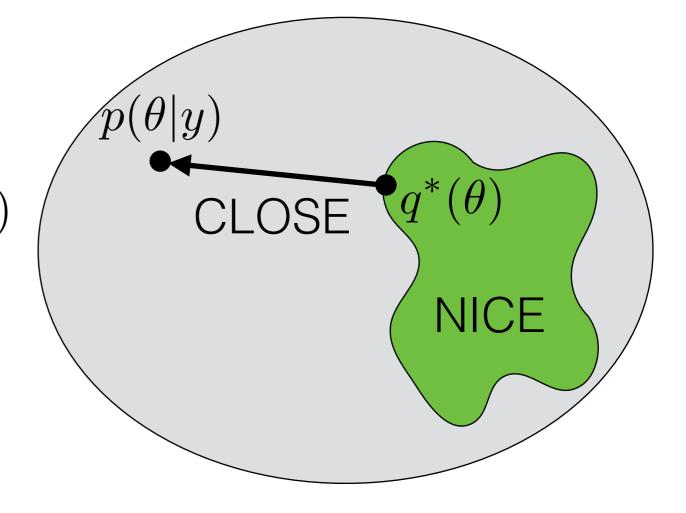


Instead: an optimization approach

$$q^* = \operatorname{argmin}_{q \in Q} f(q(\cdot), p(\cdot|y))$$

- Variational Bayes (VB): f is Kullback-Leibler divergence  $KL(q(\cdot)||p(\cdot|y))$
- VB practical success: point estimates and prediction, fast, streaming, distributed (3.6M Wikipedia, 350K Nature)

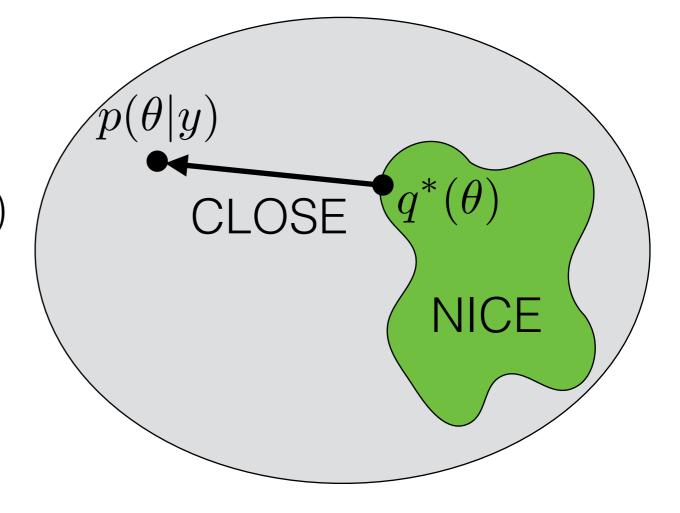
$$q^* = \operatorname{argmin}_{q \in Q} \operatorname{KL}\left(q(\cdot) || p(\cdot | y)\right)$$



$$q^* = \operatorname{argmin}_{q \in Q} \operatorname{KL}\left(q(\cdot) || p(\cdot | y)\right)$$

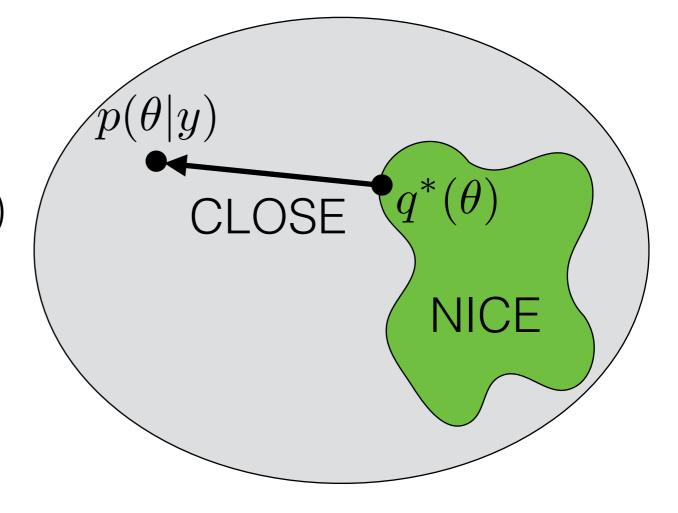
$$KL (q(\cdot)||p(\cdot|y))$$

$$:= \int q(\theta) \log \frac{q(\theta)}{p(\theta|y)} d\theta$$



$$q^* = \operatorname{argmin}_{q \in Q} \operatorname{KL}\left(q(\cdot) || p(\cdot | y)\right)$$

$$\begin{aligned} \mathrm{KL}\left(q(\cdot)||p(\cdot|y)\right) \\ &:= \int q(\theta) \log \frac{q(\theta)}{p(\theta|y)} d\theta \\ &= \int q(\theta) \log \frac{q(\theta)p(y)}{p(\theta,y)} d\theta \end{aligned}$$

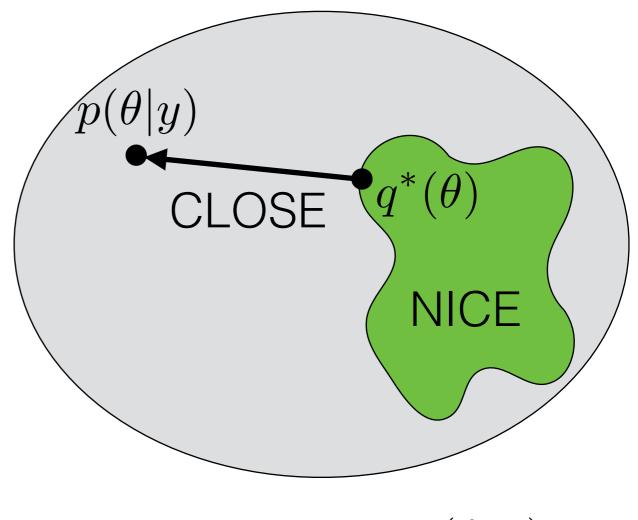


$$q^* = \operatorname{argmin}_{q \in Q} \operatorname{KL}\left(q(\cdot) || p(\cdot | y)\right)$$

$$\mathrm{KL}\left(q(\cdot)||p(\cdot|y)\right)$$

$$:= \int q(\theta) \log \frac{q(\theta)}{p(\theta|y)} d\theta$$

$$= \int q(\theta) \log \frac{q(\theta)p(y)}{p(\theta,y)} d\theta = \log p(y) - \int q(\theta) \log \frac{p(\theta,y)}{q(\theta)} d\theta$$



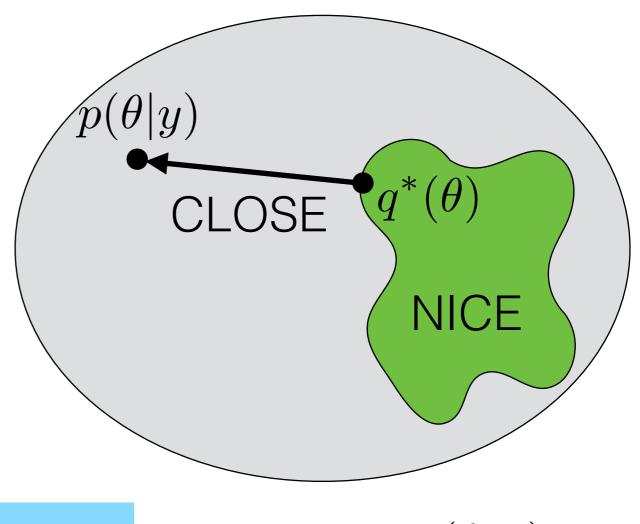
Variational Bayes

$$q^* = \operatorname{argmin}_{q \in Q} \operatorname{KL}\left(q(\cdot) || p(\cdot|y)\right)$$

$$\mathrm{KL}\left(q(\cdot)||p(\cdot|y)\right)$$

$$:= \int q(\theta) \log \frac{q(\theta)}{p(\theta|y)} d\theta$$

$$= \int q(\theta) \log \frac{q(\theta)p(y)}{p(\theta, y)} d\theta = \log p(y) - \int q(\theta) \log \frac{p(\theta, y)}{q(\theta)} d\theta$$



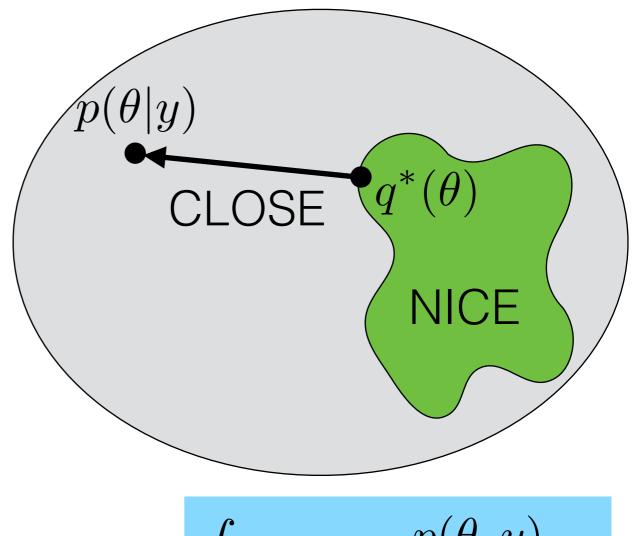
Variational Bayes

$$q^* = \operatorname{argmin}_{q \in Q} \operatorname{KL}\left(q(\cdot) || p(\cdot|y)\right)$$

$$\mathrm{KL}\left(q(\cdot)||p(\cdot|y)\right)$$

$$:= \int q(\theta) \log \frac{q(\theta)}{p(\theta|y)} d\theta$$

$$= \int q(\theta) \log \frac{q(\theta)p(y)}{p(\theta, y)} d\theta = \log p(y) - \int q(\theta) \log \frac{p(\theta, y)}{q(\theta)} d\theta$$



Variational Bayes

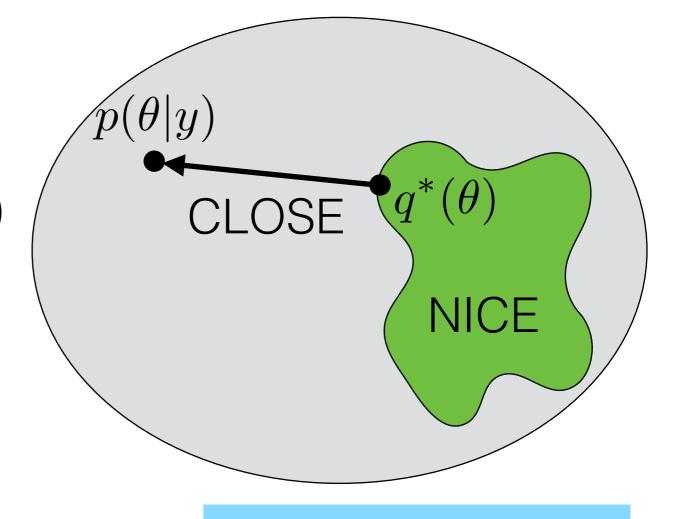
$$q^* = \operatorname{argmin}_{q \in Q} \operatorname{KL}\left(q(\cdot) || p(\cdot | y)\right)$$

$$KL (q(\cdot)||p(\cdot|y))$$

$$\int_{\alpha(\theta)} q(\theta)$$

$$:= \int q(\theta) \log \frac{q(\theta)}{p(\theta|y)} d\theta$$

$$= \int q(\theta) \log \frac{q(\theta)p(y)}{p(\theta, y)} d\theta = \log p(y) - \int q(\theta) \log \frac{p(\theta, y)}{q(\theta)} d\theta$$



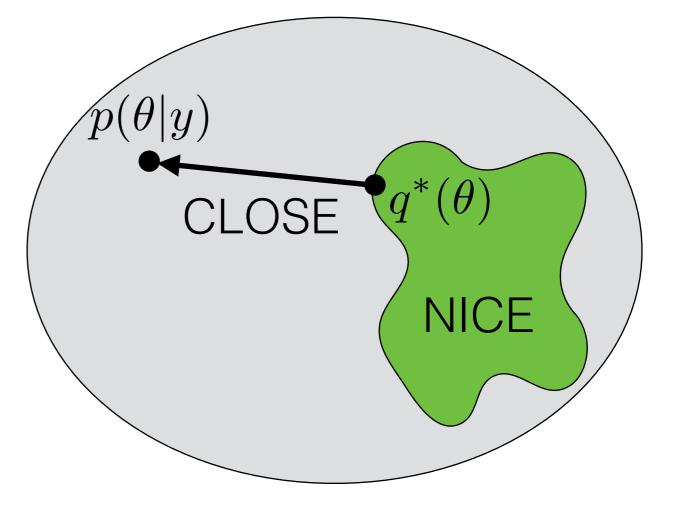
Variational Bayes

$$q^* = \operatorname{argmin}_{q \in Q} \operatorname{KL}\left(q(\cdot) || p(\cdot | y)\right)$$

$$\mathrm{KL}\left(q(\cdot)||p(\cdot|y)\right)$$

$$:= \int q(\theta) \log \frac{q(\theta)}{p(\theta|y)} d\theta$$

$$= \int q(\theta) \log \frac{q(\theta)p(y)}{p(\theta, y)} d\theta = \log p(y) - \int q(\theta) \log \frac{p(\theta, y)}{q(\theta)} d\theta$$



Variational Bayes

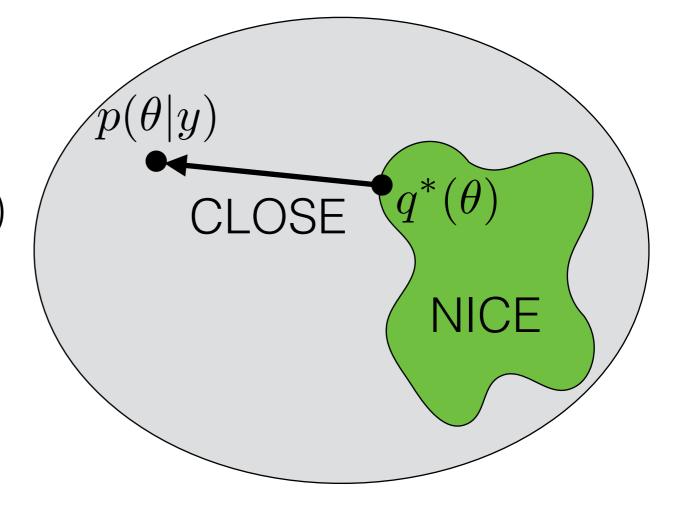
$$q^* = \operatorname{argmin}_{q \in Q} \operatorname{KL}\left(q(\cdot) || p(\cdot|y)\right)$$

$$KL (q(\cdot)||p(\cdot|y))$$

$$:= \int q(\theta) \log \frac{q(\theta)}{p(\theta|y)} d\theta$$

$$= \int q(\theta) \log \frac{q(\theta)p(y)}{p(\theta, y)} d\theta = \log p(y) - \int q(\theta) \log \frac{p(\theta, y)}{q(\theta)} d\theta$$

ullet Exercise: Show  $KL \geq 0$  [Bishop 2006, Sec 1.6.1]



Variational Bayes

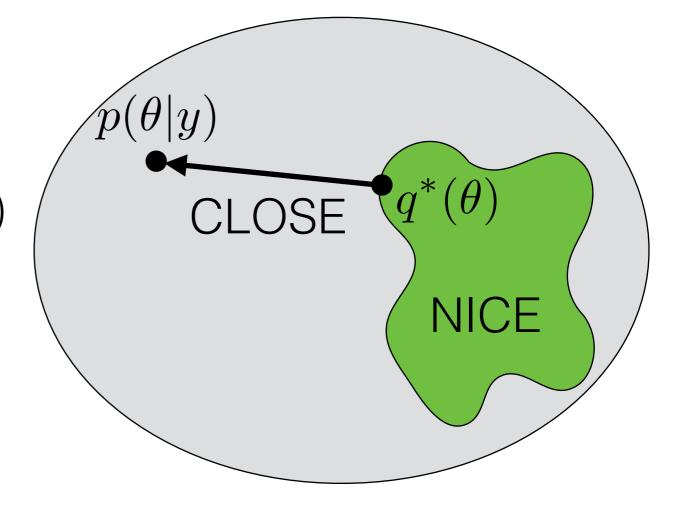
$$q^* = \operatorname{argmin}_{q \in Q} \operatorname{KL}\left(q(\cdot) || p(\cdot | y)\right)$$

$$\mathrm{KL}\left(q(\cdot)||p(\cdot|y)\right)$$

$$:= \int q(\theta) \log \frac{q(\theta)}{p(\theta|y)} d\theta$$

$$= \int q(\theta) \log \frac{q(\theta)p(y)}{p(\theta, y)} d\theta = \log p(y) - \int q(\theta) \log \frac{p(\theta, y)}{q(\theta)} d\theta$$

- ullet Exercise: Show  $\mathrm{KL} \geq 0$  [Bishop 2006, Sec 1.6.1]
- $KL \ge 0 \Rightarrow \log p(y) \ge ELBO$



Variational Bayes

$$q^* = \operatorname{argmin}_{q \in Q} \operatorname{KL}\left(q(\cdot) || p(\cdot | y)\right)$$

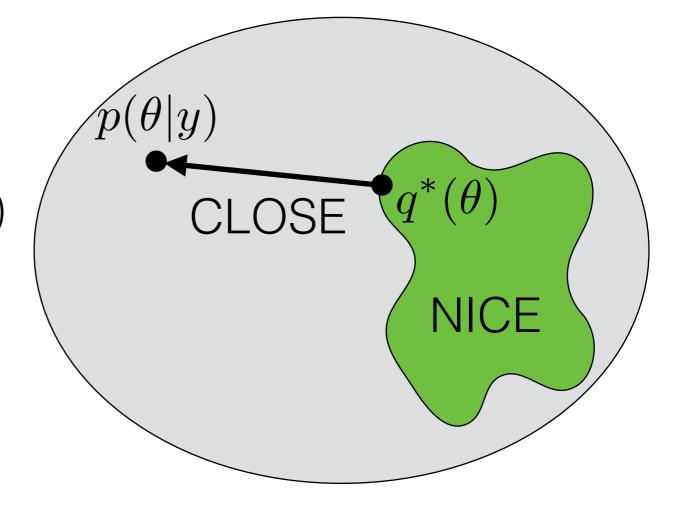
$$\mathrm{KL}\left(q(\cdot)||p(\cdot|y)\right)$$

$$:= \int q(\theta) \log \frac{q(\theta)}{p(\theta|y)} d\theta$$

$$= \int q(\theta) \log \frac{q(\theta)p(y)}{p(\theta, y)} d\theta = \log p(y) - \int q(\theta) \log \frac{p(\theta, y)}{q(\theta)} d\theta$$



- $KL \ge 0 \Rightarrow \log p(y) \ge ELBO$
- $q^* = \operatorname{argmax}_{q \in Q} \operatorname{ELBO}(q)$



Variational Bayes

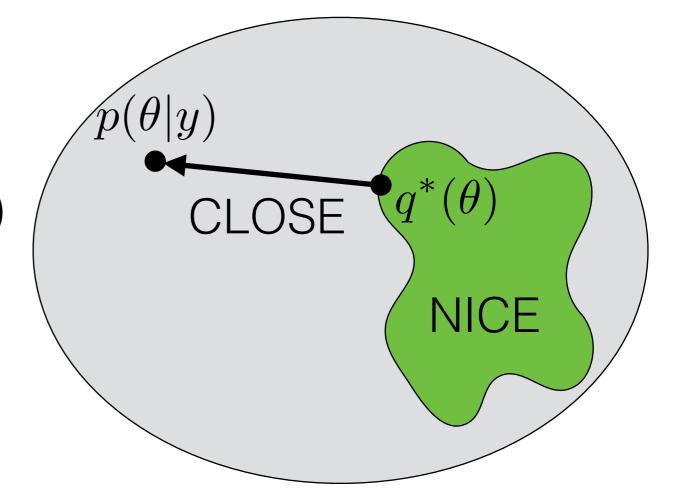
$$q^* = \operatorname{argmin}_{q \in Q} \operatorname{KL}\left(q(\cdot) || p(\cdot | y)\right)$$

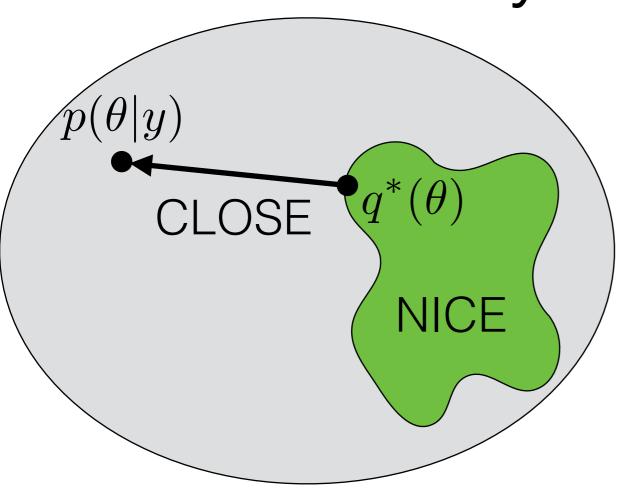
$$\mathrm{KL}\left(q(\cdot)||p(\cdot|y)\right)$$

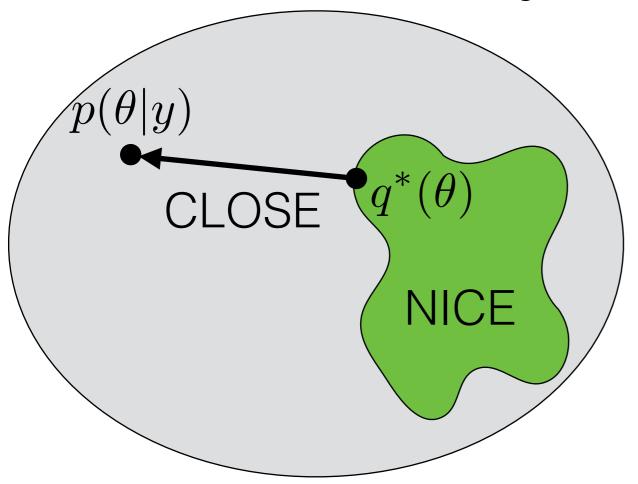
$$:= \int q(\theta) \log \frac{q(\theta)}{p(\theta|y)} d\theta$$

$$= \int q(\theta) \log \frac{q(\theta)p(y)}{p(\theta, y)} d\theta = \log p(y) - \int q(\theta) \log \frac{p(\theta, y)}{q(\theta)} d\theta$$

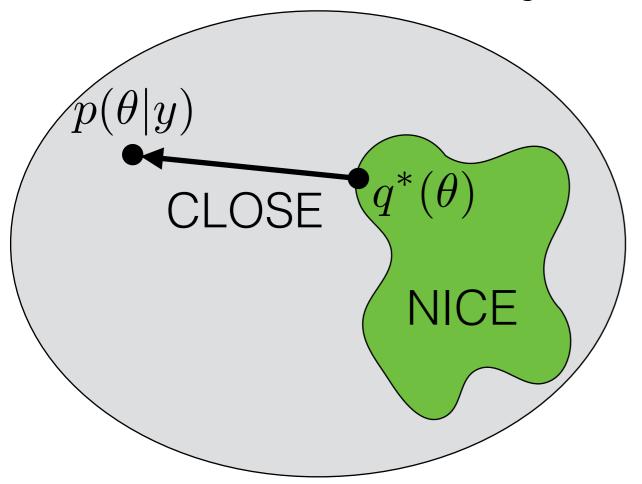
- ullet Exercise: Show  $\,\mathrm{KL} \geq 0\,$  [Bishop 2006, Sec 1.6.1]
- $KL \ge 0 \Rightarrow \log p(y) \ge ELBO$
- $q^* = \operatorname{argmax}_{q \in Q} \operatorname{ELBO}(q)$
- Why KL (in this direction)?



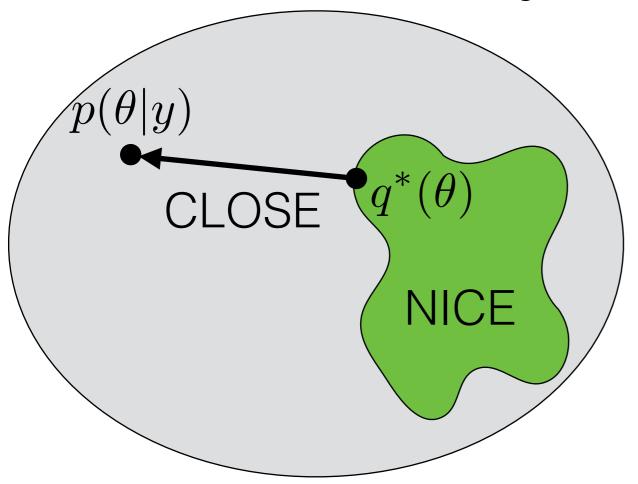




Choose "NICE" distributions

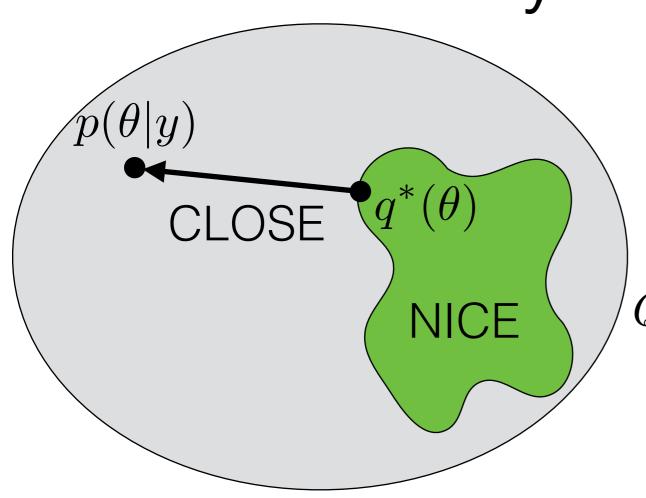


Choose "NICE" distributions



Choose "NICE" distributions

$$q^* = \operatorname{argmin}_{q \in Q} \operatorname{KL}(q(\cdot)||p(\cdot|y)|)$$

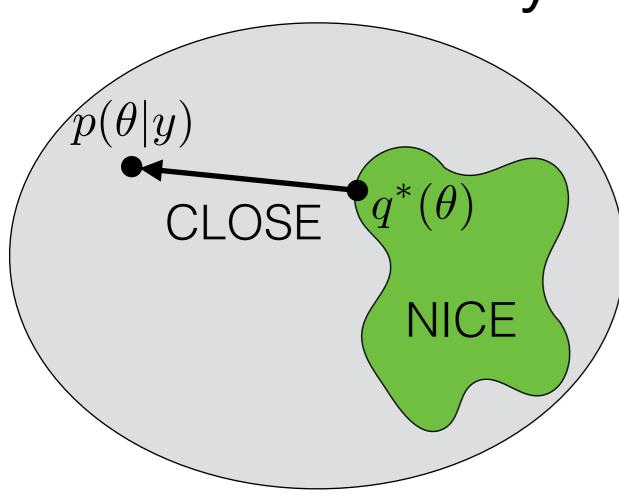


Choose "NICE" distributions

 Mean-field variational Bayes (MFVB)

$$Q_{MFVB} := \left\{ q : q(\theta) = \prod_{j=1}^{J} q_j(\theta_j) \right\}$$

$$q^* = \operatorname{argmin}_{q \in Q} \operatorname{KL}(q(\cdot)||p(\cdot|y))$$



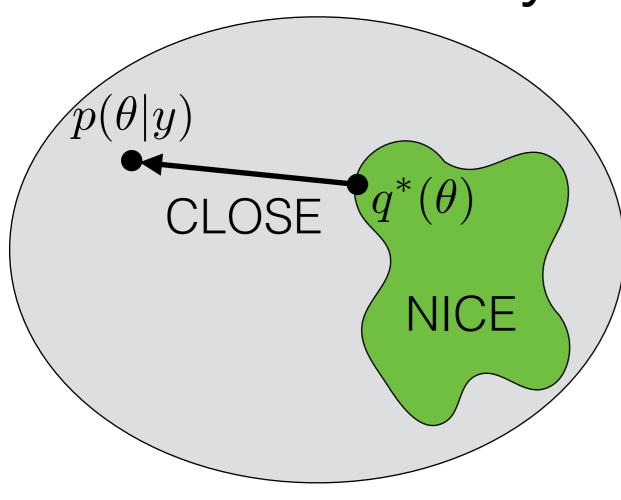
Choose "NICE" distributions

 Mean-field variational Bayes (MFVB)

$$Q_{MFVB} := \left\{ q : q(\theta) = \prod_{j=1}^{J} q_j(\theta_j) \right\}$$

Often also exponential family

$$q^* = \operatorname{argmin}_{q \in Q} \operatorname{KL}(q(\cdot)||p(\cdot|y)|)$$



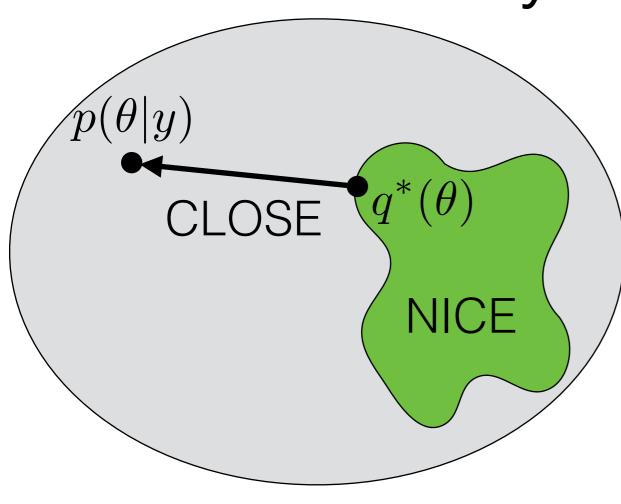
Choose "NICE" distributions

 Mean-field variational Bayes (MFVB)

$$Q_{MFVB} := \left\{ q : q(\theta) = \prod_{j=1}^{J} q_j(\theta_j) \right\}$$

- Often also exponential family
- Not a modeling assumption

$$q^* = \operatorname{argmin}_{q \in Q} \operatorname{KL}(q(\cdot)||p(\cdot|y))$$

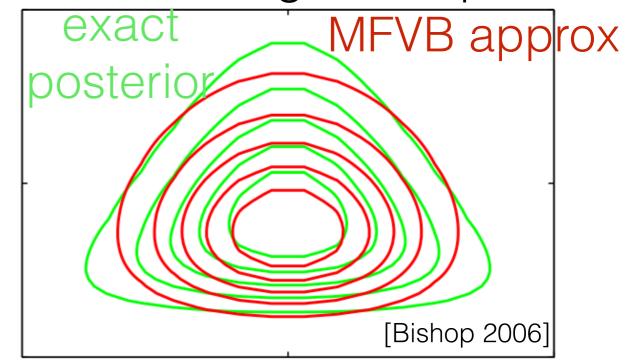


Choose "NICE" distributions

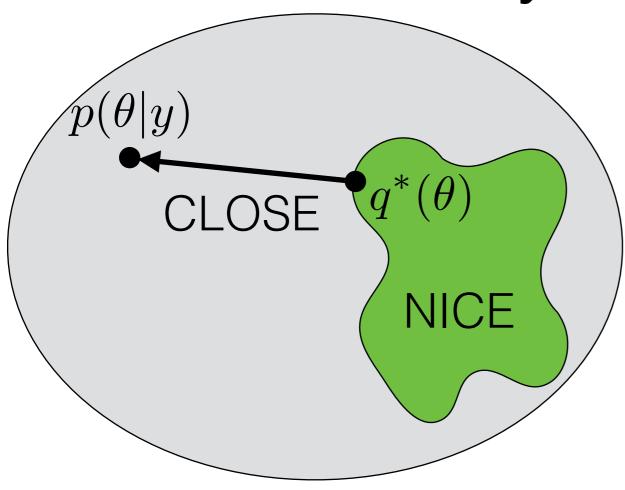
 Mean-field variational Bayes (MFVB)

$$Q_{MFVB} := \left\{ q : q(\theta) = \prod_{j=1}^{J} q_j(\theta_j) \right\}$$

- Often also exponential family
- Not a modeling assumption



$$q^* = \operatorname{argmin}_{q \in Q} \operatorname{KL}(q(\cdot)||p(\cdot|y))$$



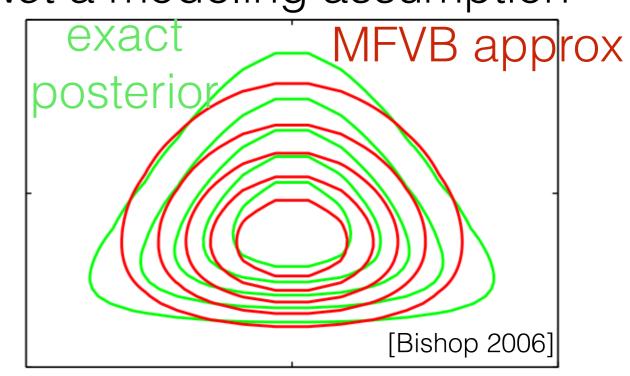
Choose "NICE" distributions

 Mean-field variational Bayes (MFVB)

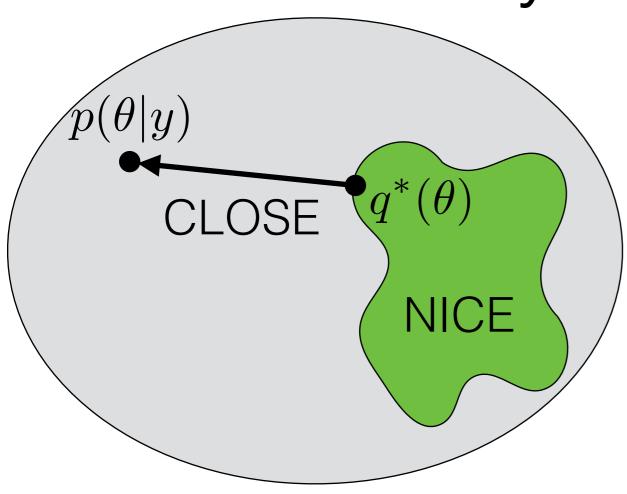
$$Q_{MFVB} := \left\{ q : q(\theta) = \prod_{j=1}^{J} q_j(\theta_j) \right\}$$

- Often also exponential family
- Not a modeling assumption

Now we have an optimization problem; how to solve it?



$$q^* = \operatorname{argmin}_{q \in Q} \operatorname{KL}(q(\cdot)||p(\cdot|y))$$



Choose "NICE" distributions

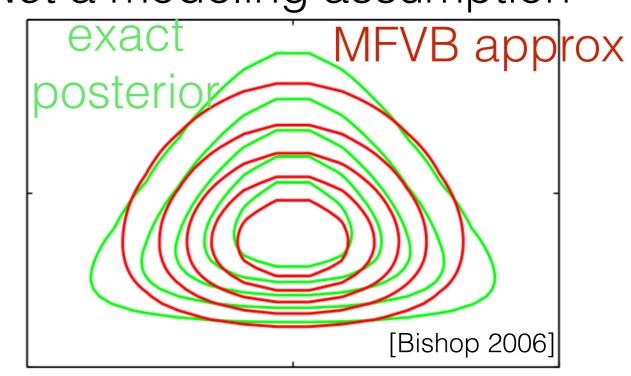
 Mean-field variational Bayes (MFVB)

$$Q_{MFVB} := \left\{ q : q(\theta) = \prod_{j=1}^{J} q_j(\theta_j) \right\}$$

- Often also exponential family
- Not a modeling assumption

Now we have an optimization problem; how to solve it?

 One option: Coordinate descent in  $q_1, \ldots, q_J$ 



Use  $q^*$  to approximate  $p(\cdot|y)$ 

Use  $q^*$  to approximate  $p(\cdot|y)$ 

Use  $q^*$  to approximate  $p(\cdot|y)$ 

Optimization 
$$q^* = \operatorname{argmin}_{q \in Q} f(q(\cdot), p(\cdot|y))$$

Variational Bayes  $q^* = \operatorname{argmin}_{q \in Q} KL(q(\cdot)||p(\cdot|y))$ 

Use  $q^*$  to approximate  $p(\cdot|y)$ 

Optimization 
$$q^* = \operatorname{argmin}_{q \in Q} f(q(\cdot), p(\cdot|y))$$

Variational Bayes

$$q^* = \operatorname{argmin}_{q \in Q} KL(q(\cdot)||p(\cdot|y))$$

Mean-field variational Bayes

$$q^* = \operatorname{argmin}_{q \in Q_{\text{MFVB}}} KL(q(\cdot)||p(\cdot|y))$$

Use  $q^*$  to approximate  $p(\cdot|y)$ 

Optimization 
$$q^* = \operatorname{argmin}_{q \in Q} f(q(\cdot), p(\cdot|y))$$

Variational Bayes

$$q^* = \operatorname{argmin}_{q \in Q} KL(q(\cdot)||p(\cdot|y))$$

Mean-field variational Bayes

$$q^* = \operatorname{argmin}_{q \in Q_{\text{MFVB}}} KL(q(\cdot)||p(\cdot|y))$$

Use  $q^*$  to approximate  $p(\cdot|y)$ 

Variational Bayes

$$q^* = \operatorname{argmin}_{q \in Q} KL(q(\cdot)||p(\cdot|y))$$

Mean-field variational Bayes

$$q^* = \operatorname{argmin}_{q \in Q_{\text{MFVB}}} KL(q(\cdot)||p(\cdot|y))$$

Coordinate descent

Use  $q^*$  to approximate  $p(\cdot|y)$ 

Optimization 
$$q^* = \operatorname{argmin}_{q \in Q} f(q(\cdot), p(\cdot|y))$$

Variational Bayes

$$q^* = \operatorname{argmin}_{q \in Q} KL(q(\cdot)||p(\cdot|y))$$

Mean-field variational Bayes

$$q^* = \operatorname{argmin}_{q \in Q_{\text{MFVB}}} KL(q(\cdot)||p(\cdot|y))$$

- Coordinate descent
- Stochastic variational inference (\$VI) [Hoffman et al/2013]

Use  $q^*$  to approximate  $p(\cdot|y)$ 

Optimization 
$$q^* = \operatorname{argmin}_{q \in Q} f(q(\cdot), p(\cdot|y))$$

Variational Bayes

$$q^* = \operatorname{argmin}_{q \in Q} KL(q(\cdot)||p(\cdot|y))$$

Mean-field variational Bayes

$$q^* = \operatorname{argmin}_{q \in Q_{\text{MFVB}}} KL(q(\cdot)||p(\cdot|y))$$

- Coordinate descent
- Stochastic variational inference (\$VI) [Hoffman et al 2013]
- Automatic differentiation variational inference (ADVI) [Kucukelbir et al 2015, 2017]

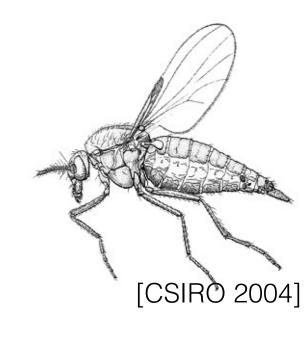
#### Roadmap

- Bayes & Approximate Bayes review
- What is:
  - Variational Bayes (VB)
  - Mean-field variational Bayes (MFVB)
- Why use MFVB?
- When can we trust MFVB?
- Where do we go from here?

#### Roadmap

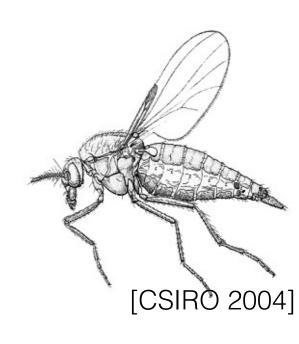
- Bayes & Approximate Bayes review
- What is:
  - Variational Bayes (VB)
  - Mean-field variational Bayes (MFVB)
- Why use MFVB?
- When can we trust MFVB?
- Where do we go from here?

• Catalogued midge wing lengths (mm)  $y = (y_1, \dots, y_N)$ 



- Catalogued midge wing lengths (mm)  $y = (y_1, \dots, y_N)$
- Model:

$$p(y|\theta): \quad y_n \stackrel{iid}{\sim} \mathcal{N}(\mu, \sigma^2), \quad n = 1, \dots, N$$



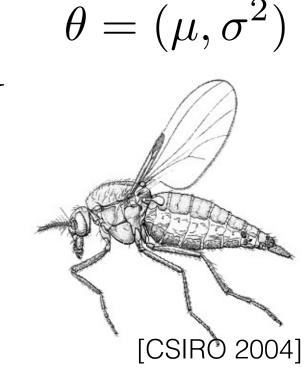
- Catalogued midge wing lengths (mm)  $y = (y_1, \dots, y_N)$
- Parameters of interest: population mean and variance  $\theta = (\mu, \sigma^2)$
- Model:

$$p(y|\theta): \quad y_n \stackrel{iid}{\sim} \mathcal{N}(\mu, \sigma^2), \quad n = 1, \dots, N$$



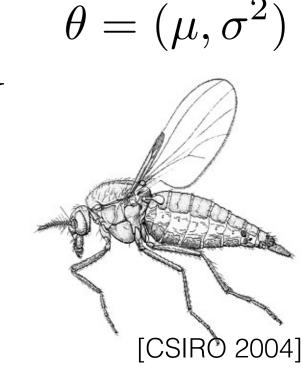
- Catalogued midge wing lengths (mm)  $y = (y_1, \dots, y_N)$
- Parameters of interest: population mean and variance
- Model:

$$p(y|\theta): \quad y_n \stackrel{iid}{\sim} \mathcal{N}(\mu, \sigma^2), \quad n = 1, \dots, N$$
$$p(\theta): \quad (\sigma^2)^{-1} \sim \text{Gamma}(a_0, b_0)$$
$$\mu|\sigma^2 \sim \mathcal{N}(\mu_0, \lambda_0 \sigma^2)$$



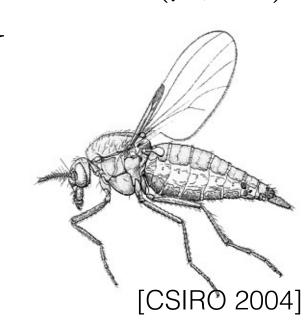
- Catalogued midge wing lengths (mm)  $y = (y_1, \dots, y_N)$
- Parameters of interest: population mean and variance

Model (conjugate prior): 
$$p(y|\theta): \quad y_n \overset{iid}{\sim} \mathcal{N}(\mu, \sigma^2), \quad n = 1, \dots, N$$
 
$$p(\theta): \quad (\sigma^2)^{-1} \sim \operatorname{Gamma}(a_0, b_0)$$
 
$$\mu|\sigma^2 \sim \mathcal{N}(\mu_0, \lambda_0 \sigma^2)$$



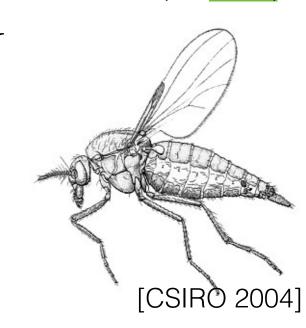
- Catalogued midge wing lengths (mm)  $y = (y_1, \dots, y_N)$
- Parameters of interest: population mean and variance  $\theta = (\mu, \sigma^2)$

Model (conjugate prior): [Exercise: find the posterior] 
$$p(y|\theta): \quad y_n \overset{iid}{\sim} \mathcal{N}(\mu,\sigma^2), \quad n=1,\dots,N$$
 
$$p(\theta): \quad (\sigma^2)^{-1} \sim \mathrm{Gamma}(a_0,b_0)$$
 
$$\mu|\sigma^2 \sim \mathcal{N}(\mu_0,\lambda_0\sigma^2)$$



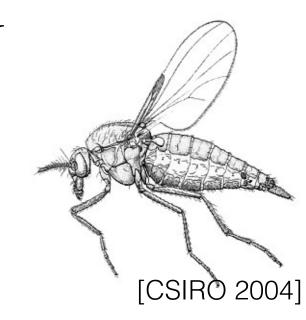
- Catalogued midge wing lengths (mm)  $y = (y_1, \dots, y_N)$
- Parameters of interest: population mean and variance  $\theta = (\mu, \sigma^2)$
- Model (conjugate prior): [Exercise: find the posterior]

$$p(y|\theta): y_n \stackrel{iid}{\sim} \mathcal{N}(\mu, \sigma^2), \quad n = 1, \dots, N$$
  
 $p(\theta): (\sigma^2)^{-1} \sim \text{Gamma}(a_0, b_0)$   
 $\mu|\sigma^2 \sim \mathcal{N}(\mu_0, \lambda_0 \sigma^2)$ 



- Catalogued midge wing lengths (mm)  $y = (y_1, \dots, y_N)$
- Parameters of interest: population mean and precision  $\theta = (u, \tau)$
- Model (conjugate prior): [Exercise: find the posterior]

$$p(y|\theta)$$
:  $y_n \stackrel{iid}{\sim} \mathcal{N}(\mu, \sigma^2)$ ,  $n = 1, ..., N$   
 $p(\theta)$ :  $(\sigma^2)^{-1} \sim \text{Gamma}(a_0, b_0)$   
 $\mu|\sigma^2 \sim \mathcal{N}(\mu_0, \lambda_0 \sigma^2)$ 



- Catalogued midge wing lengths (mm)  $y = (y_1, \dots, y_N)$
- Parameters of interest: population mean and precision
- Model (conjugate prior): [Exercise: find the posterior]

$$p(y|\theta): y_n \stackrel{iid}{\sim} \mathcal{N}(\mu, \sigma^2), \quad n = 1, \dots, N$$

$$p(\theta): (\sigma^2)^{-1} \sim \text{Gamma}(a_0, b_0)$$

$$\mu | \sigma^2 \sim \mathcal{N}(\mu_0, \lambda_0 \sigma^2)$$

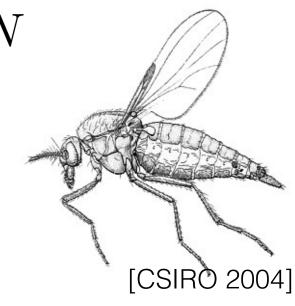


- Catalogued midge wing lengths (mm)  $y = (y_1, \dots, y_N)$
- Parameters of interest: population mean and precision
- Model (conjugate prior): [Exercise: find the posterior]

$$p(y|\theta): y_n \stackrel{iid}{\sim} \mathcal{N}(\mu, \tau^{-1}), \quad n = 1, \dots, N$$

$$p(\theta): \quad \boldsymbol{\tau} \sim \operatorname{Gamma}(a_0, b_0)$$

$$\mu | \tau \sim \mathcal{N}(\mu_0, (\rho_0 \tau)^{-1})$$



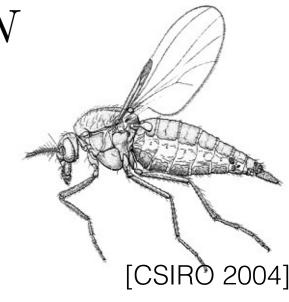
 $\theta = (\mu, \tau)$ 

- Catalogued midge wing lengths (mm)  $y = (y_1, \dots, y_N)$
- Parameters of interest: population mean and precision
- Model (conjugate prior): [Exercise: find the posterior]

$$p(y|\theta): y_n \stackrel{iid}{\sim} \mathcal{N}(\mu, \tau^{-1}), \quad n = 1, \dots, N$$

$$p(\theta): \quad \tau \sim \text{Gamma}(a_0, b_0)$$

$$\mu | \tau \sim \mathcal{N}(\mu_0, (\rho_0 \tau)^{-1})$$



 $\theta = (\mu, \tau)$ 

- Catalogued midge wing lengths (mm)  $y = (y_1, \dots, y_N)$
- Parameters of interest: population mean and precision  $\theta = (\mu, \tau)$
- Model (conjugate prior): [Exercise: find the posterior]

$$p(y|\theta): y_n \stackrel{iid}{\sim} \mathcal{N}(\mu, \tau^{-1}), \quad n = 1, \dots, N$$

$$p(\theta): \quad \tau \sim \text{Gamma}(a_0, b_0)$$

$$\mu | \tau \sim \mathcal{N}(\mu_0, (\rho_0 \tau)^{-1})$$

• Exercise: check  $p(\mu, \tau | y) \neq f_1(\mu, y) f_2(\tau, y)$ 



- Catalogued midge wing lengths (mm)  $y = (y_1, \dots, y_N)$
- Parameters of interest: population mean and precision
- Model (conjugate prior): [Exercise: find the posterior]

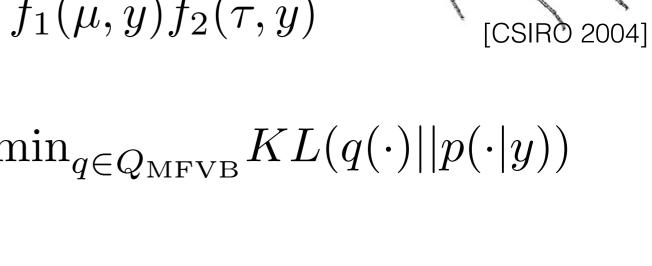
$$p(y|\theta): y_n \stackrel{iid}{\sim} \mathcal{N}(\mu, \tau^{-1}), \quad n = 1, \dots, N$$

$$p(\theta): \quad \tau \sim \text{Gamma}(a_0, b_0)$$

$$\mu | \tau \sim \mathcal{N}(\mu_0, (\rho_0 \tau)^{-1})$$

- Exercise: check  $p(\mu, \tau | y) \neq f_1(\mu, y) f_2(\tau, y)$
- MFVB approximation:

$$q^*(\mu, \tau) = q_{\mu}^*(\mu)q_{\tau}^*(\tau) = \operatorname{argmin}_{q \in Q_{\text{MFVB}}} KL(q(\cdot)||p(\cdot|y))$$



 $\theta = (\mu, \tau)$ 

- Catalogued midge wing lengths (mm)  $y = (y_1, \dots, y_N)$
- Parameters of interest: population mean and precision
- Model (conjugate prior): [Exercise: find the posterior]  $\theta = (\mu, \tau)$

$$p(y|\theta): y_n \stackrel{iid}{\sim} \mathcal{N}(\mu, \tau^{-1}), \quad n = 1, \dots, N$$

$$p(\theta): \quad \tau \sim \text{Gamma}(a_0, b_0)$$

$$\mu | \tau \sim \mathcal{N}(\mu_0, (\rho_0 \tau)^{-1})$$

- Exercise: check  $p(\mu, \tau | y) \neq f_1(\mu, y) f_2(\tau, y)$
- MFVB approximation:

$$q^*(\mu, \tau) = q_{\mu}^*(\mu)q_{\tau}^*(\tau) = \operatorname{argmin}_{q \in Q_{\text{MFVB}}} KL(q(\cdot)||p(\cdot|y))$$

• Coordinate descent [Exercise: derive this] [Bishop 2006, Sec 10.1.3]

[CSIRO 2004]

- Catalogued midge wing lengths (mm)  $y = (y_1, \dots, y_N)$
- Parameters of interest: population mean and precision
- Model (conjugate prior): [Exercise: find the posterior]  $\theta = (\mu, \tau)$

$$p(y|\theta): y_n \stackrel{iid}{\sim} \mathcal{N}(\mu, \tau^{-1}), \quad n = 1, \dots, N$$

$$p(\theta): \quad \tau \sim \text{Gamma}(a_0, b_0)$$

$$\mu | \tau \sim \mathcal{N}(\mu_0, (\rho_0 \tau)^{-1})$$

- Exercise: check  $p(\mu, \tau | y) \neq f_1(\mu, y) f_2(\tau, y)$
- MFVB approximation:

$$q^*(\mu, \tau) = q_{\mu}^*(\mu)q_{\tau}^*(\tau) = \operatorname{argmin}_{q \in Q_{\text{MFVB}}} KL(q(\cdot)||p(\cdot|y))$$

• Coordinate descent [Exercise: derive this] [Bishop 2006, Sec 10.1.3]

$$q_{\mu}^*(\mu) = \mathcal{N}(\mu|\mu_N, \rho_N^{-1}) \qquad q_{\tau}^*(\tau) = \operatorname{Gamma}(\tau|a_N, b_N)$$

[CSIRO 2004]

- Catalogued midge wing lengths (mm)  $y = (y_1, \dots, y_N)$
- Parameters of interest: population mean and precision
- Model (conjugate prior): [Exercise: find the posterior]

$$p(y|\theta): y_n \stackrel{iid}{\sim} \mathcal{N}(\mu, \tau^{-1}), \quad n = 1, \dots, N$$

$$p(\theta): \quad \tau \sim \text{Gamma}(a_0, b_0)$$

$$\mu | \tau \sim \mathcal{N}(\mu_0, (\rho_0 \tau)^{-1})$$

- Exercise: check  $p(\mu, \tau | y) \neq f_1(\mu, y) f_2(\tau, y)$
- MFVB approximation:

$$q^*(\mu, \tau) = q_{\mu}^*(\mu)q_{\tau}^*(\tau) = \operatorname{argmin}_{q \in Q_{\text{MFVB}}} KL(q(\cdot)||p(\cdot|y))$$

• Coordinate descent [Exercise: derive this] [Bishop 2006, Sec 10.1.3]

$$q_{\mu}^{*}(\mu) = \mathcal{N}(\mu|\mu_{N}, \rho_{N}^{-1})$$
  $q_{\tau}^{*}(\tau) = \text{Gamma}(\tau|a_{N}, b_{N})$ 

"variational parameters"

[CSIRO 2004]

 $\theta = (\mu, \tau)$ 

- Catalogued midge wing lengths (mm)  $y = (y_1, \dots, y_N)$
- Parameters of interest: population mean and precision
- Model (conjugate prior): [Exercise: find the posterior]

$$p(y|\theta): y_n \stackrel{iid}{\sim} \mathcal{N}(\mu, \tau^{-1}), \quad n = 1, \dots, N$$

$$p(\theta): \quad \tau \sim \text{Gamma}(a_0, b_0)$$

$$\mu | \tau \sim \mathcal{N}(\mu_0, (\rho_0 \tau)^{-1})$$

- Exercise: check  $p(\mu, \tau | y) \neq f_1(\mu, y) f_2(\tau, y)$
- MFVB approximation:

$$q^*(\mu, \tau) = q_{\mu}^*(\mu)q_{\tau}^*(\tau) = \operatorname{argmin}_{q \in Q_{\text{MFVB}}} KL(q(\cdot)||p(\cdot|y))$$

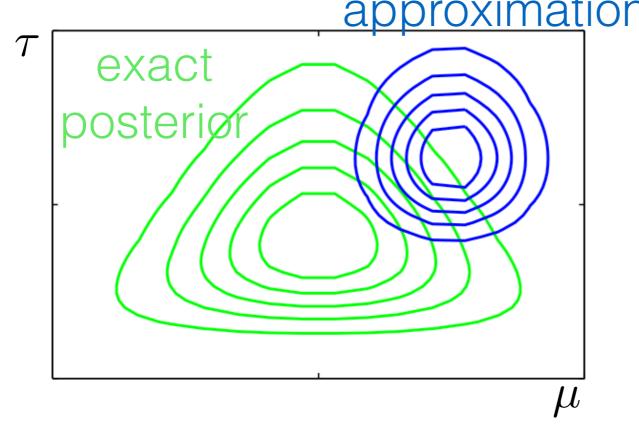
• Coordinate descent [Exercise: derive this] [Bishop 2006, Sec 10.1.3]

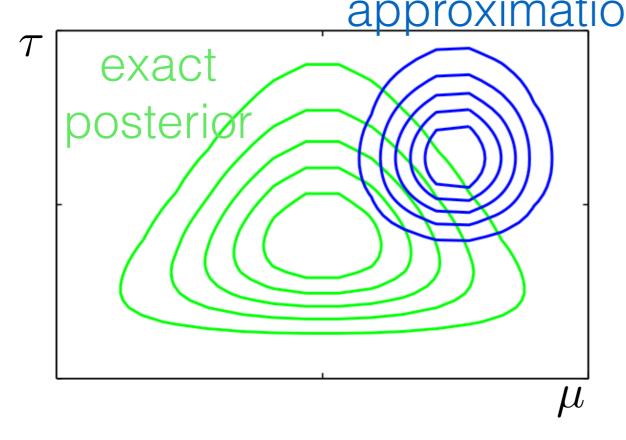
$$q_{\mu}^{*}(\mu) = \mathcal{N}(\mu|\mu_{N}, \rho_{N}^{-1})$$
  $q_{\tau}^{*}(\tau) = \text{Gamma}(\tau|a_{N}, b_{N})$ 

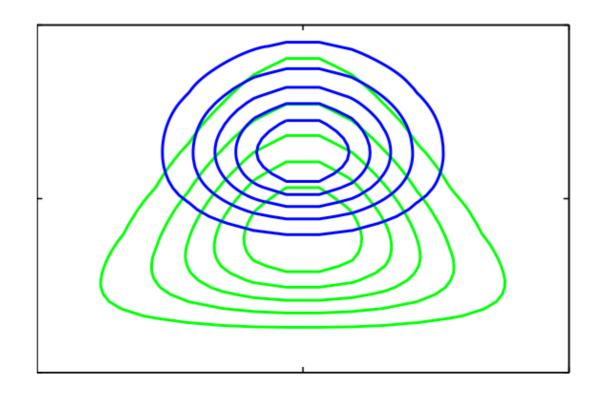
• Iterate:  $(\mu_N, \rho_N) = f(a_N, b_N)$  "variational  $(a_N, b_N) = g(\mu_N, \rho_N)$  parameters"

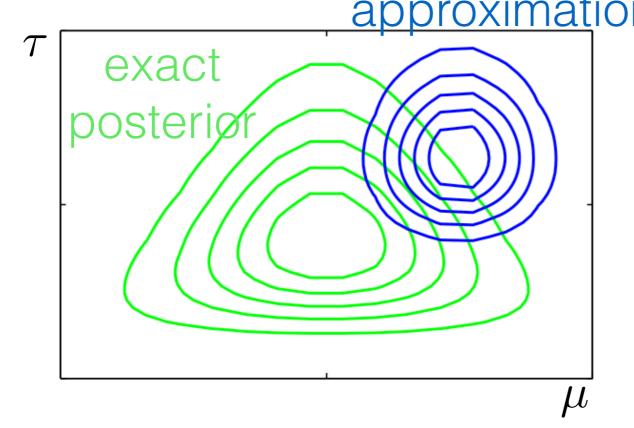
 $\theta = (\mu, \tau)$ 

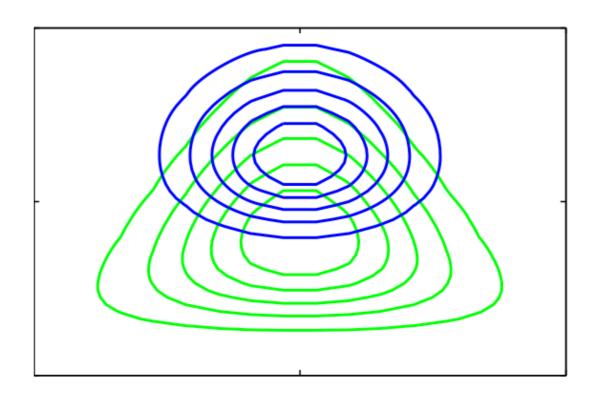
[CSIRO 2004]

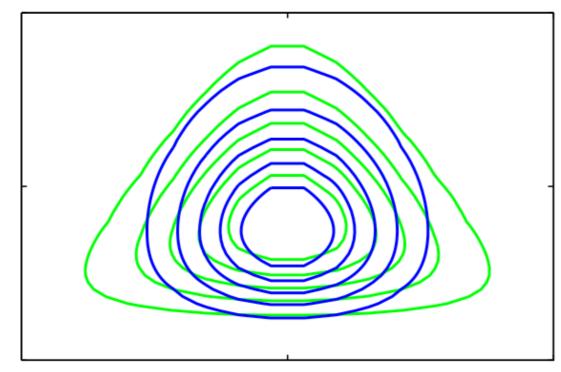


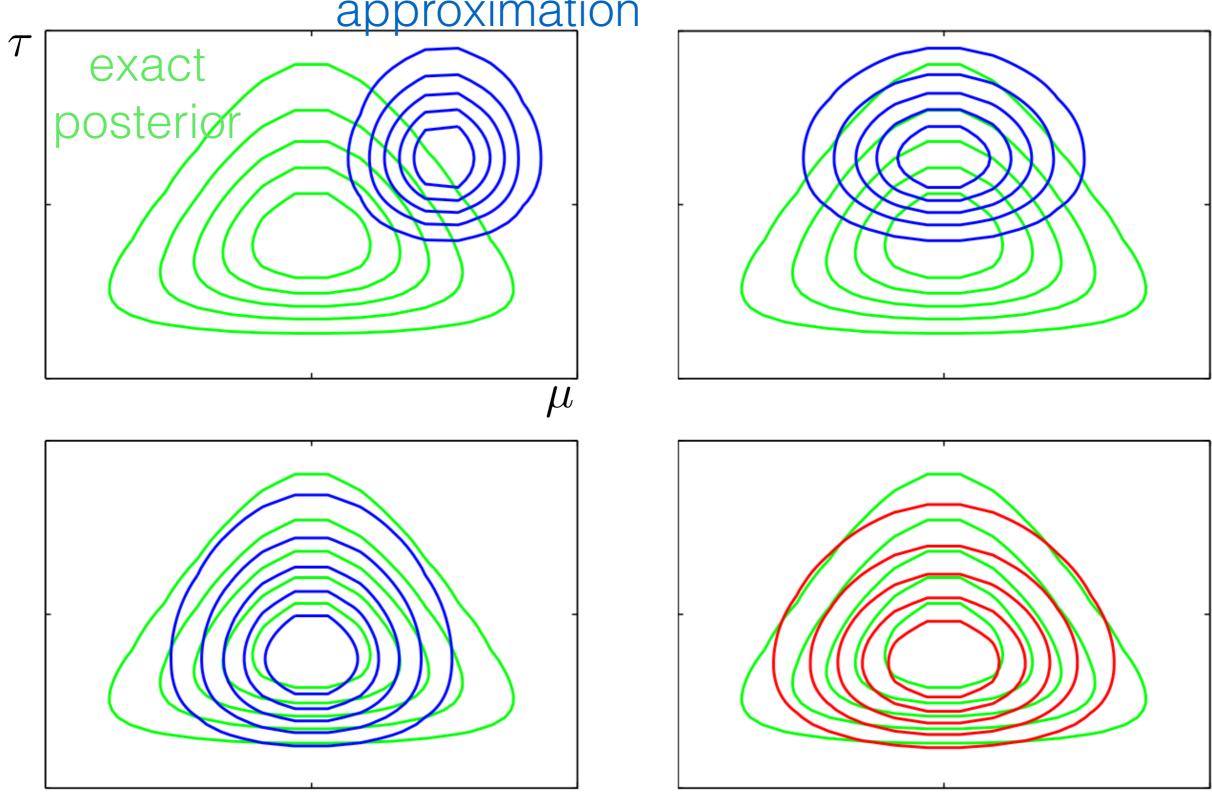


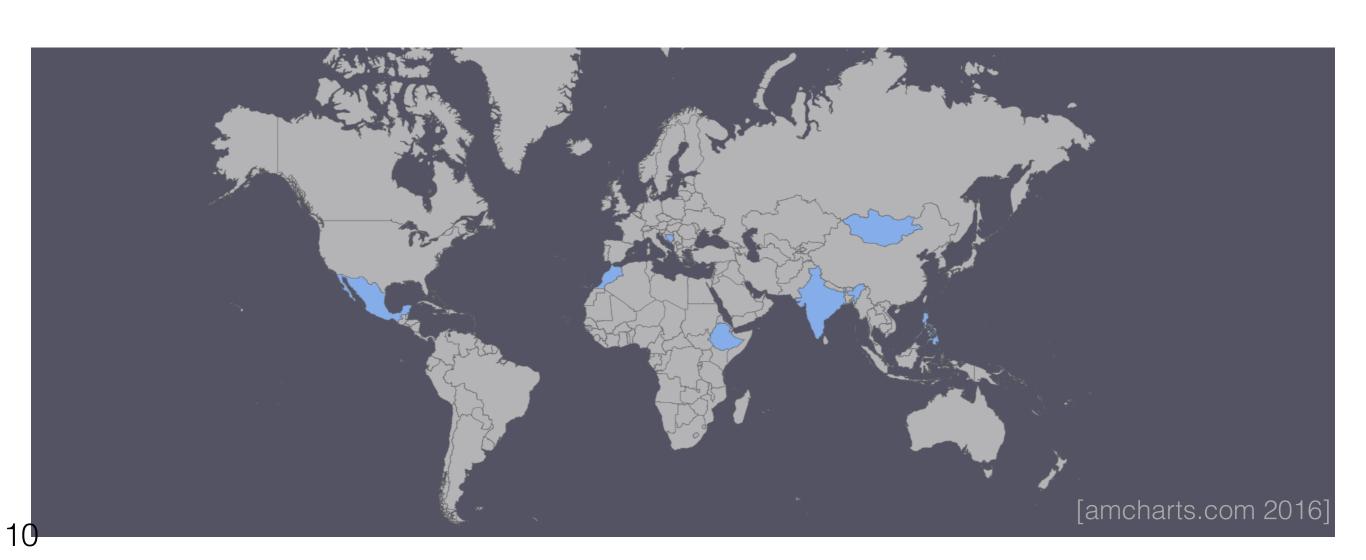




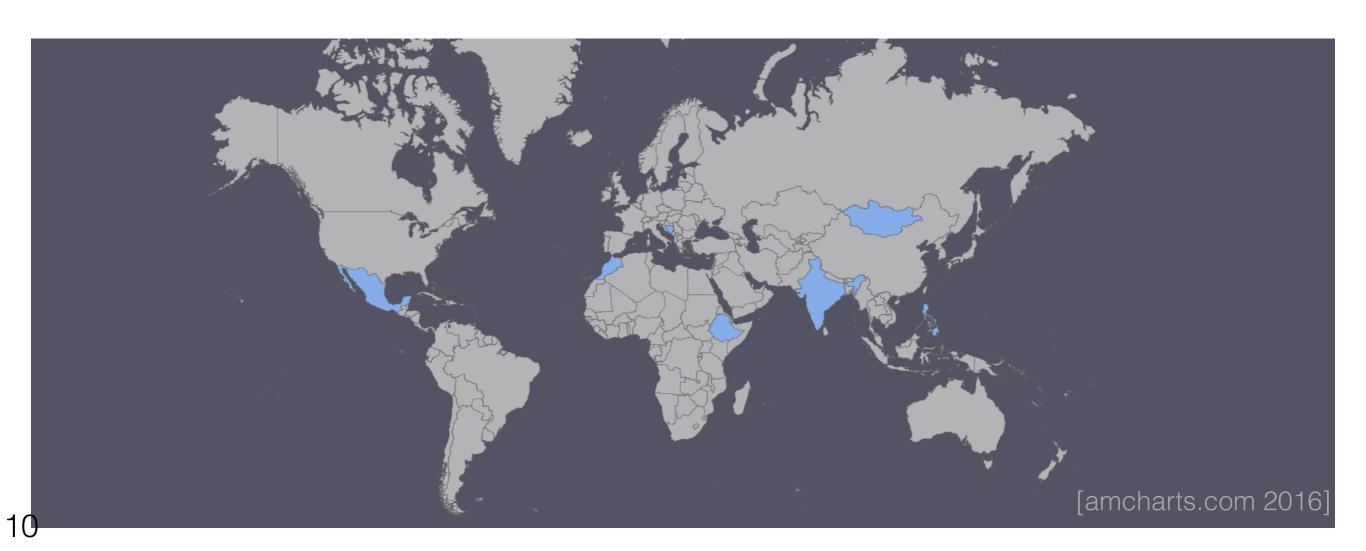




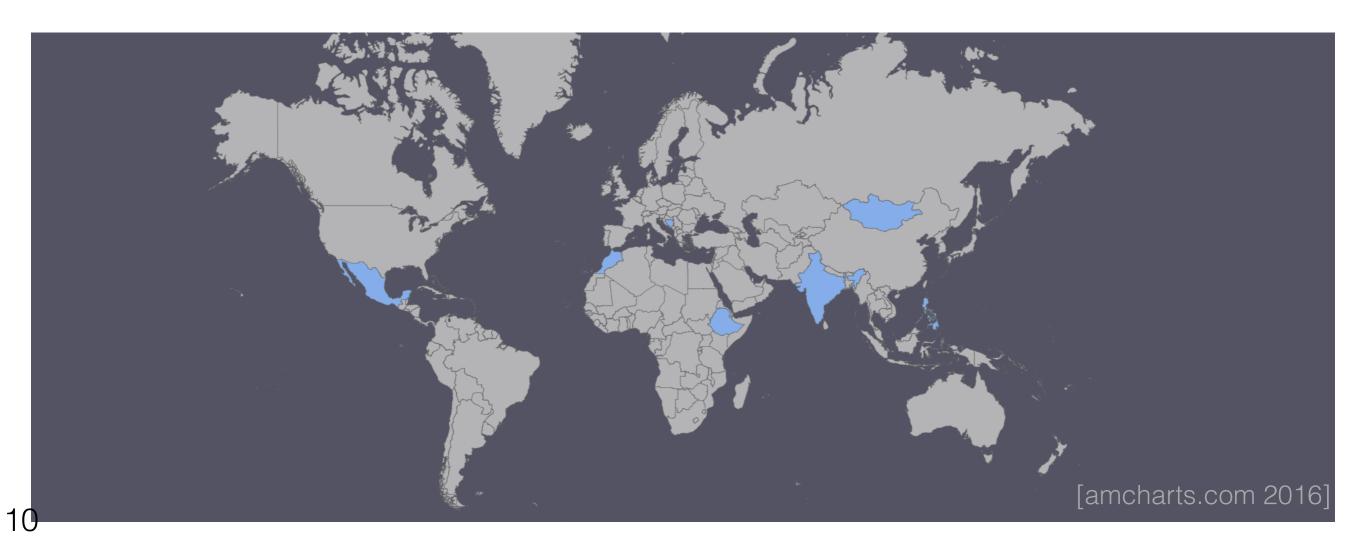




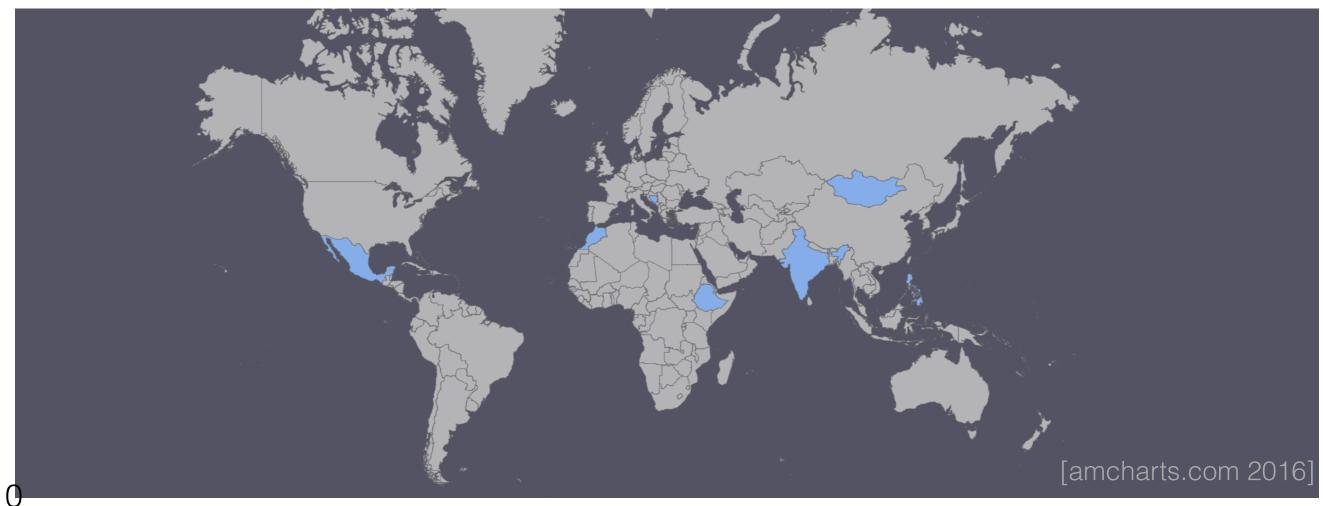
• Simplified from Meager (2018a)



- Simplified from Meager (2018a)
- K = 7 microcredit trials (Mexico, Mongolia, Bosnia, India, Morocco, Philippines, Ethiopia)



- Simplified from Meager (2018a)
- K = 7 microcredit trials (Mexico, Mongolia, Bosnia, India, Morocco, Philippines, Ethiopia)
- $N_k$  businesses in kth site (~900 to ~17K)



- Simplified from Meager (2018a)
- K = 7 microcredit trials (Mexico, Mongolia, Bosnia, India, Morocco, Philippines, Ethiopia)
- $N_k$  businesses in kth site (~900 to ~17K)

- Simplified from Meager (2018a)
- K = 7 microcredit trials (Mexico, Mongolia, Bosnia, India, Morocco, Philippines, Ethiopia)
- $N_k$  businesses in kth site (~900 to ~17K)
- Profit of nth business at kth site:

- Simplified from Meager (2018a)
- K = 7 microcredit trials (Mexico, Mongolia, Bosnia, India, Morocco, Philippines, Ethiopia)
- $N_k$  businesses in kth site (~900 to ~17K)
- Profit of nth business at kth site:



- Simplified from Meager (2018a)
- K = 7 microcredit trials (Mexico, Mongolia, Bosnia, India, Morocco, Philippines, Ethiopia)
- $N_k$  businesses in kth site (~900 to ~17K)
- Profit of nth business at kth site:

```
profit y_{kn} \stackrel{indep}{\sim} \mathcal{N}( , )
```

- Simplified from Meager (2018a)
- K = 7 microcredit trials (Mexico, Mongolia, Bosnia, India, Morocco, Philippines, Ethiopia)
- $N_k$  businesses in kth site (~900 to ~17K)
- Profit of nth business at kth site:

```
profit y_{kn} \stackrel{indep}{\sim} \mathcal{N}(\mu_k) , )
```

- Simplified from Meager (2018a)
- K = 7 microcredit trials (Mexico, Mongolia, Bosnia, India, Morocco, Philippines, Ethiopia)
- $N_k$  businesses in kth site (~900 to ~17K)
- Profit of nth business at kth site:

```
profit y_{kn} \stackrel{indep}{\sim} \mathcal{N}(\mu_k + T_{kn}\tau_k, )
```

- Simplified from Meager (2018a)
- K = 7 microcredit trials (Mexico, Mongolia, Bosnia, India, Morocco, Philippines, Ethiopia)

1 if microcredit

- $N_k$  businesses in kth site (~900 to ~17K)
- Profit of *n*th business at *k*th site:

profit  $y_{kn} \overset{indep}{\sim} \mathcal{N}(\mu_k + T_{kn}\tau_k, \quad )$ 

- Simplified from Meager (2018a)
- K = 7 microcredit trials (Mexico, Mongolia, Bosnia, India, Morocco, Philippines, Ethiopia)

1 if microcredit

- $N_k$  businesses in kth site (~900 to ~17K)
- Profit of *n*th business at *k*th site:

profit  $y_{kn} \stackrel{indep}{\sim} \mathcal{N}(\mu_k + T_{kn} \tau_k, \quad )$ 

- Simplified from Meager (2018a)
- K = 7 microcredit trials (Mexico, Mongolia, Bosnia, India, Morocco, Philippines, Ethiopia)

1 if microcredit

- $N_k$  businesses in kth site (~900 to ~17K)
- Profit of *n*th business at *k*th site:

profit  $y_{kn} \overset{indep}{\sim} \mathcal{N}(\mu_k + T_{kn}\tau_k, \sigma_k^2)$ 

- Simplified from Meager (2018a)
- K = 7 microcredit trials (Mexico, Mongolia, Bosnia, India, Morocco, Philippines, Ethiopia)

\_1 if microcredit

- $N_k$  businesses in kth site (~900 to ~17K)
- Profit of *n*th business at *k*th site:

profit  $y_{kn} \overset{indep}{\sim} \mathcal{N}(\mu_k + T_{kn}\tau_k, \sigma_k^2)$ 

- Simplified from Meager (2018a)
- K = 7 microcredit trials (Mexico, Mongolia, Bosnia, India, Morocco, Philippines, Ethiopia)

1 if microcredit

- $N_k$  businesses in kth site (~900 to ~17K)
- Profit of *n*th business at *k*th site:

profit  $y_{kn} \overset{indep}{\sim} \mathcal{N}(\mu_k + T_{kn}\tau_k, \sigma_k^2)$ 

Priors and hyperpriors:

- Simplified from Meager (2018a)
- K = 7 microcredit trials (Mexico, Mongolia, Bosnia, India, Morocco, Philippines, Ethiopia)

→ 1 if microcredit

- $N_k$  businesses in kth site (~900 to ~17K)
- Profit of *n*th business at *k*th site:

profit 
$$y_{kn} \stackrel{indep}{\sim} \mathcal{N}(\mu_k + T_{kn}\tau_k, \sigma_k^2)$$

Priors and hyperpriors:

$$\left(\begin{array}{c} \mu_k \\ \tau_k \end{array}\right) \stackrel{iid}{\sim} \mathcal{N}\left(\left(\begin{array}{c} \mu \\ \tau \end{array}\right), C\right)$$

- Simplified from Meager (2018a)
- K = 7 microcredit trials (Mexico, Mongolia, Bosnia, India, Morocco, Philippines, Ethiopia)

→1 if microcredit

- $N_k$  businesses in kth site (~900 to ~17K)
- Profit of nth business at kth site:

profit 
$$y_{kn} \stackrel{indep}{\sim} \mathcal{N}(\mu_k + T_{kn}\tau_k, \sigma_k^2)$$

Priors and hyperpriors:

$$\begin{pmatrix} \mu_k \\ \tau_k \end{pmatrix} \stackrel{iid}{\sim} \mathcal{N} \left( \begin{pmatrix} \mu \\ \tau \end{pmatrix}, C \right)$$

$$\sigma_k^{-2} \stackrel{iid}{\sim} \Gamma(a,b)$$

- Simplified from Meager (2018a)
- K = 7 microcredit trials (Mexico, Mongolia, Bosnia, India, Morocco, Philippines, Ethiopia)
- $N_k$  businesses in kth site (~900 to ~17K)
- Profit of nth business at kth site:

profit 
$$y_{kn} \stackrel{indep}{\sim} \mathcal{N}(\mu_k + T_{kn}\tau_k, \sigma_k^2)$$

Priors and hyperpriors:

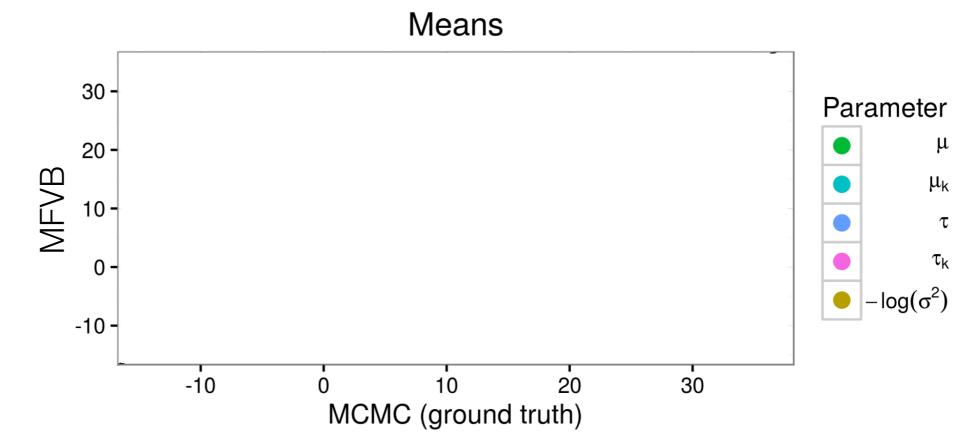
$$\begin{pmatrix} \mu_k \\ \tau_k \end{pmatrix} \stackrel{iid}{\sim} \mathcal{N} \left( \begin{pmatrix} \mu \\ \tau \end{pmatrix}, C \right) \qquad \begin{pmatrix} \mu \\ \tau \end{pmatrix} \stackrel{iid}{\sim} \mathcal{N} \left( \begin{pmatrix} \mu_0 \\ \tau_0 \end{pmatrix}, \Lambda^{-1} \right)$$

$$\sigma_k^{-2} \stackrel{iid}{\sim} \Gamma(a, b)$$
  $C \sim \text{Sep&LKJ}(\eta, c, d)$ 

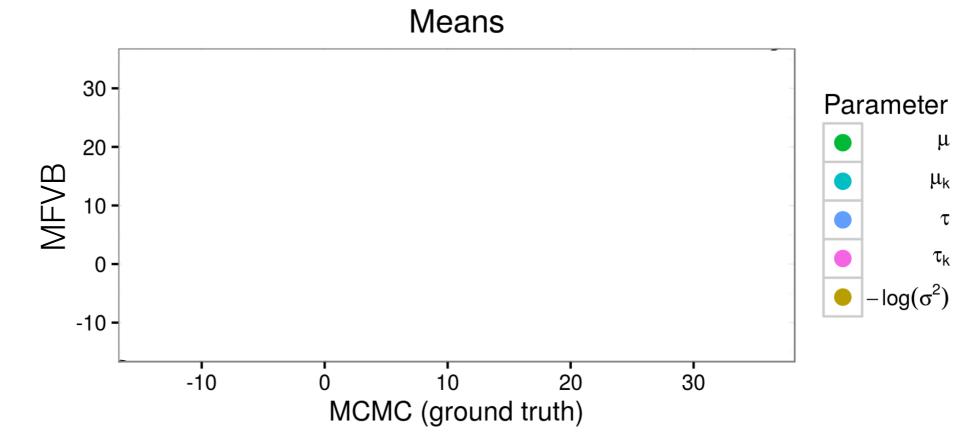
√1 if microcredit

MFVB: Do we need to check the output?

MFVB: How will we know if it's working?



One set of 2500 MCMC draws:45 minutes

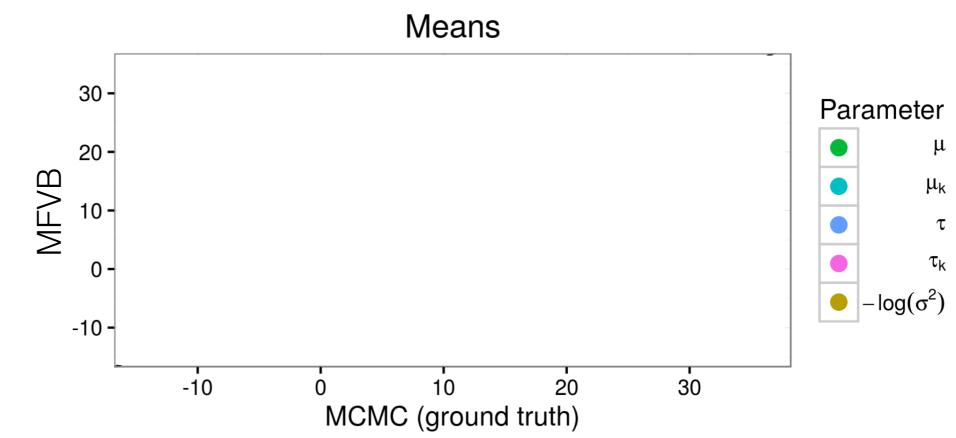


 One set of 2500 MCMC draws:

#### 45 minutes

MFVB optimization:

#### <1 min

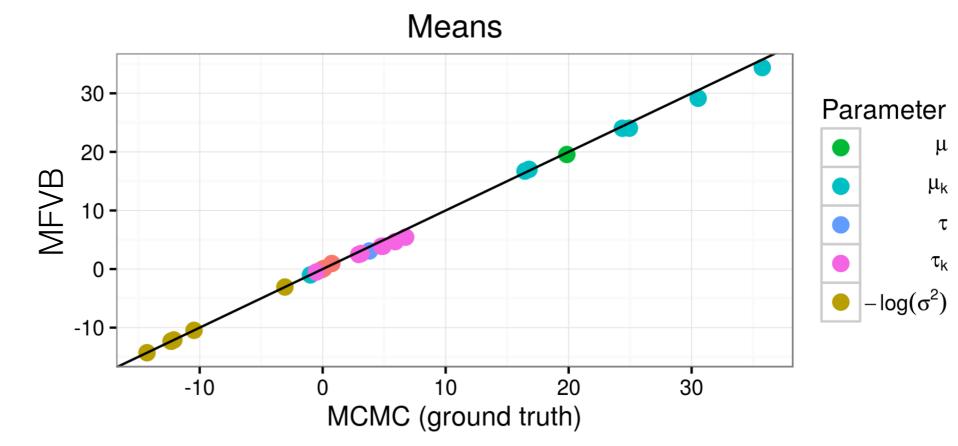


 One set of 2500 MCMC draws:

#### 45 minutes

MFVB optimization:

<1 min

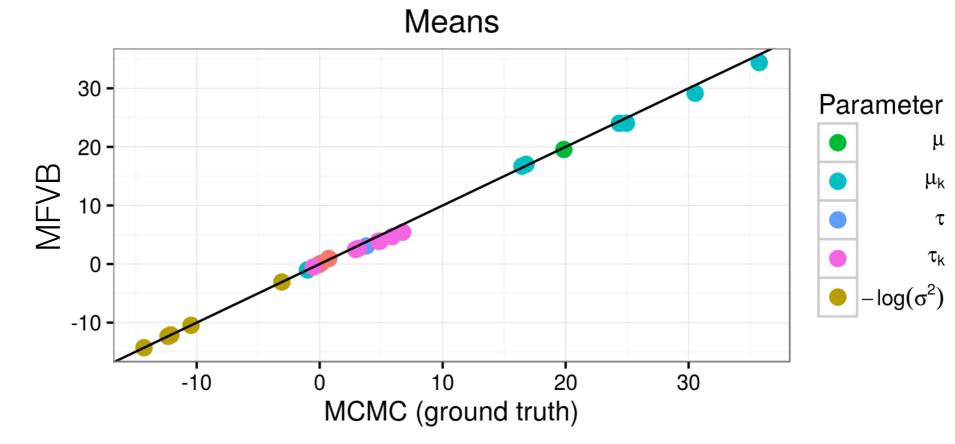


 One set of 2500 MCMC draws:

#### 45 minutes

MFVB optimization:

<1 min



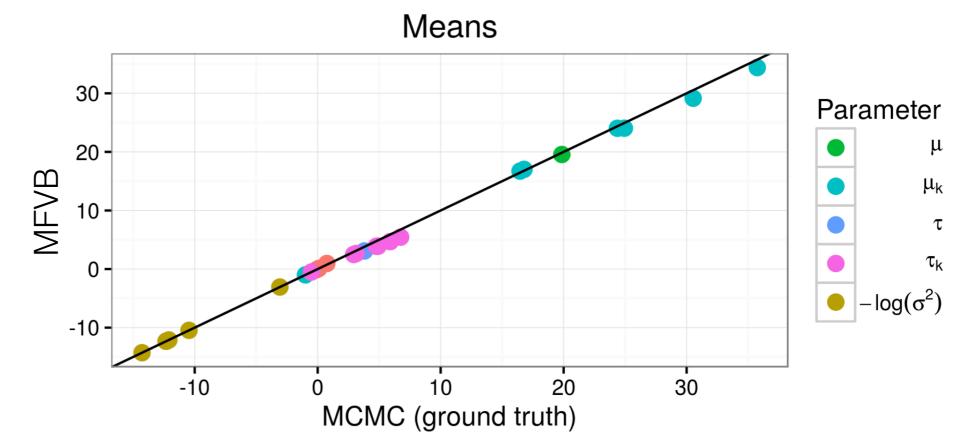
- Click-through conversion prediction
- Q: Will a customer (e.g.) buy a product after clicking?

 One set of 2500 MCMC draws:

#### 45 minutes

MFVB optimization:

<1 min



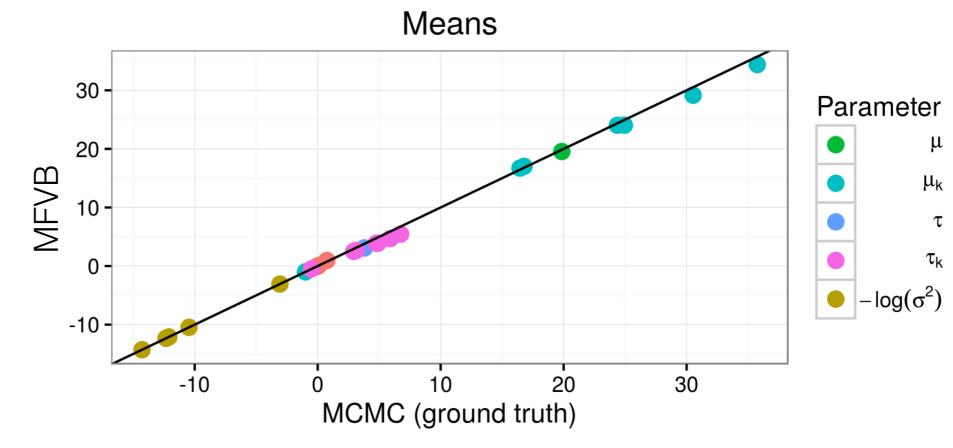
- Click-through conversion prediction
- Q: Will a customer (e.g.) buy a product after clicking?
- Q: How predictive of conversion are different features?

 One set of 2500 MCMC draws:

#### 45 minutes

MFVB optimization:

<1 min



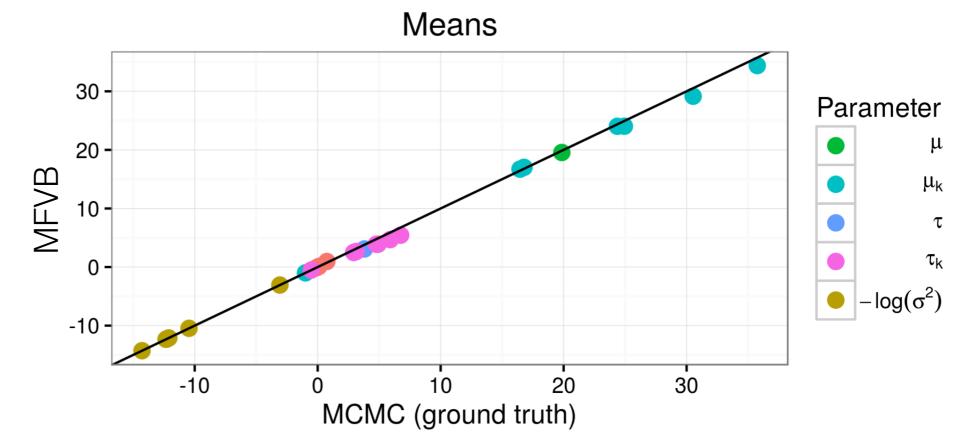
- Click-through conversion prediction
- Q: Will a customer (e.g.) buy a product after clicking?
- Q: How predictive of conversion are different features?
- Logistic GLMM

 One set of 2500 MCMC draws:

#### 45 minutes

MFVB optimization:

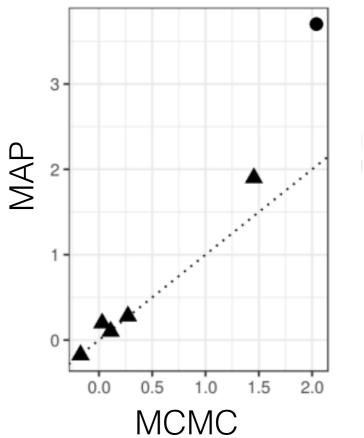
<1 min



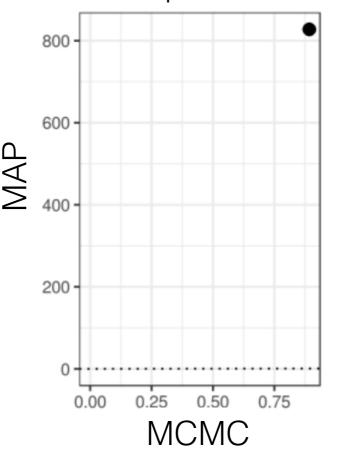
- Click-through conversion prediction
- Q: Will a customer (e.g.) buy a product after clicking?
- Q: How predictive of conversion are different features?
- Logistic GLMM; N = 61,895 subset to compare to MCMC

• MAP: **12 s** 

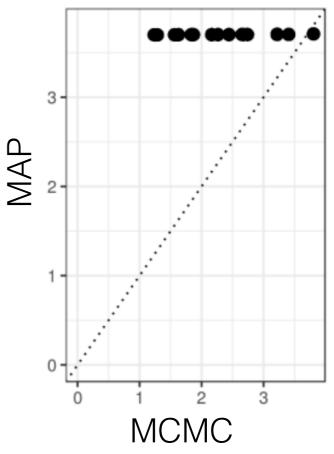
Global parameters (-τ)



Global parameter  $\tau$ 

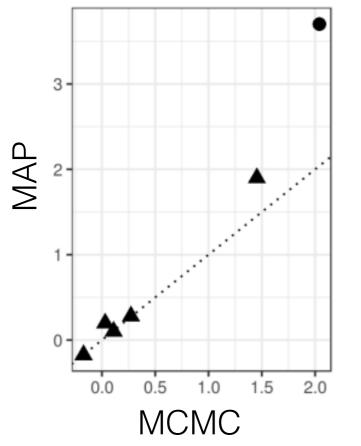


Local parameters

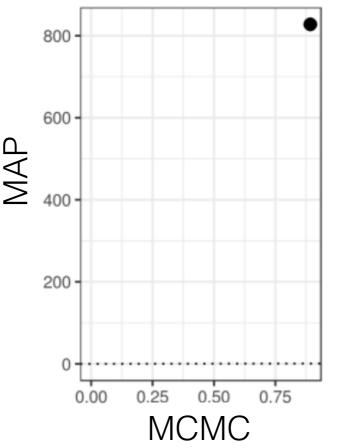


• MAP: **12 s** 

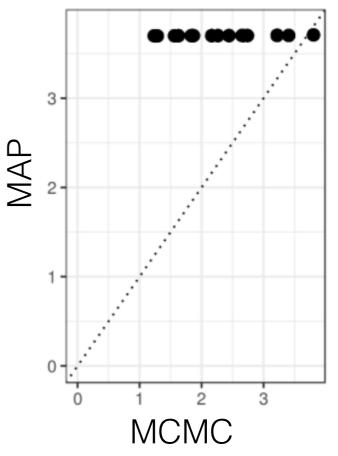
Global parameters (-τ)



Global parameter  $\tau$ 



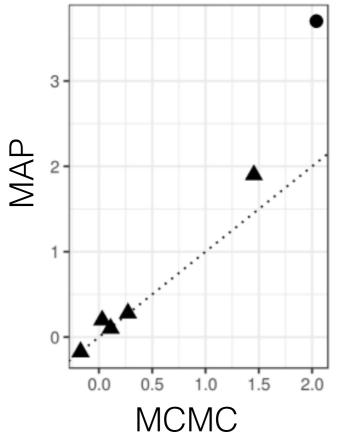
Local parameters



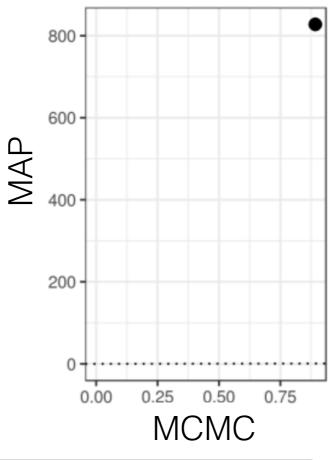
• MAP: **12 s** 

• MFVB: **57** s

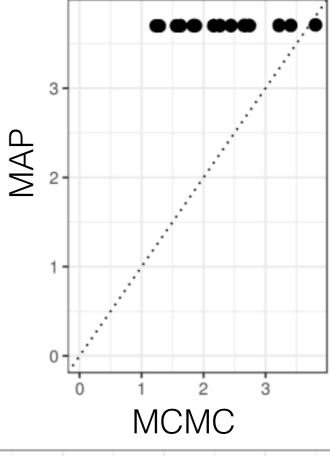
Global parameters (-τ)





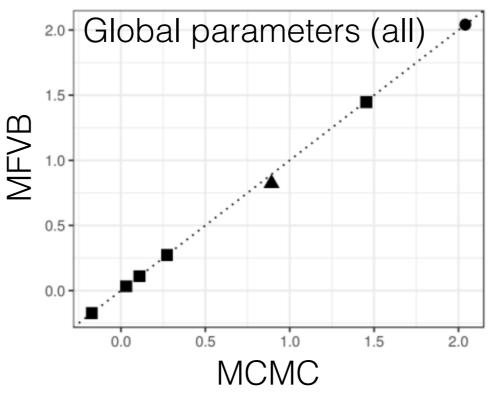


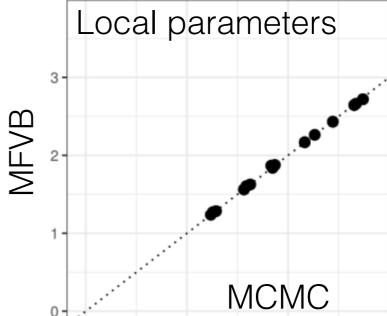
Local parameters





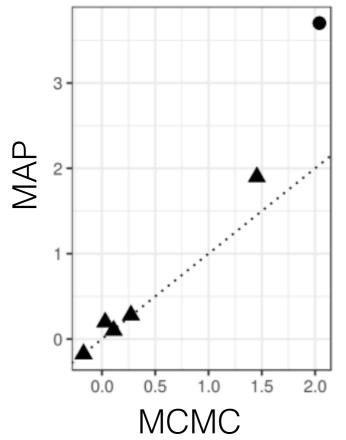
• MFVB: **57** s

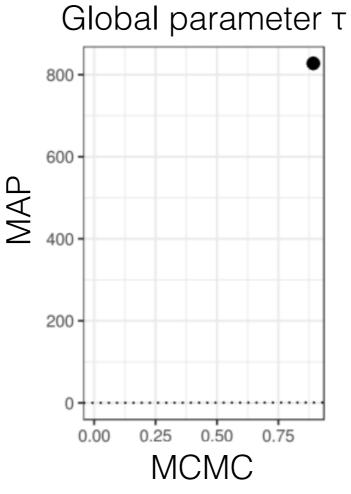


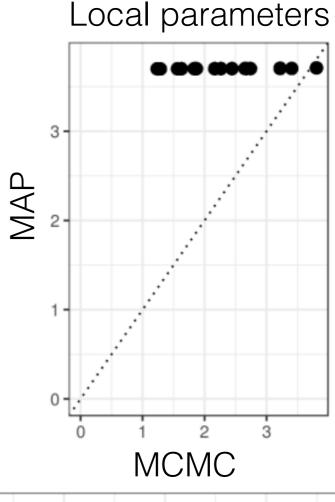


[Giordano, Broderick, Jordan 2018]

Global parameters (-τ)



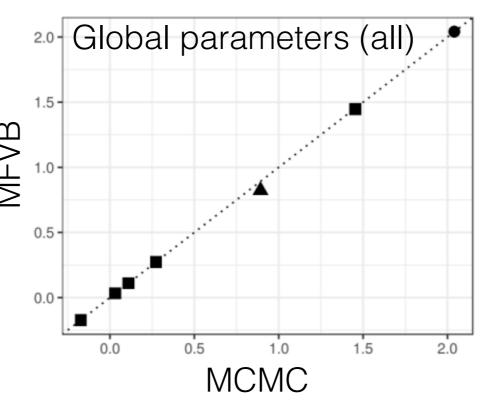


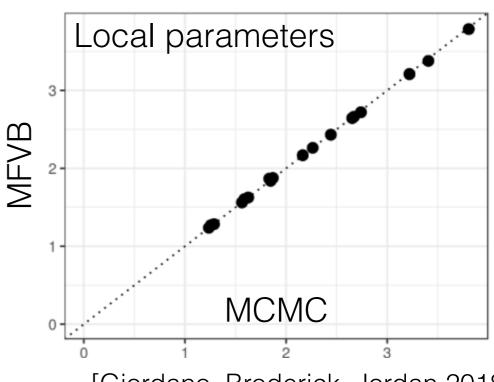


• MAP: **12** s

• MFVB: **57** s

MCMC (5K samples):
21,066 s
(5.85 h)





## Why use MFVB?

Topic discovery

"Arts"	${ m `Budgets''}$	"Children"	"Education"
NEW	MILLION	CHILDREN	SCHOOL
FILM	TAX	WOMEN	STUDENTS
SHOW	PROGRAM	PEOPLE	SCHOOLS
MUSIC	BUDGET	CHILD	EDUCATION
MOVIE	BILLION	YEARS	TEACHERS
PLAY	FEDERAL	FAMILIES	HIGH
MUSICAL	YEAR	WORK	PUBLIC
BEST	SPENDING	PARENTS	TEACHER
ACTOR	NEW	SAYS	BENNETT
FIRST	STATE	FAMILY	MANIGAT
YORK	PLAN	WELFARE	NAMPHY
OPERA	MONEY	MEN	STATE
THEATER	PROGRAMS	PERCENT	PRESIDENT
ACTRESS	GOVERNMENT	CARE	ELEMENTARY
LOVE	CONGRESS	LIFE	HAITI

The William Randolph Hearst Foundation will give \$1.25 million to Lincoln Center, Metropolitan Opera Co., New York Philharmonic and Juilliard School. "Our board felt that we had a real opportunity to make a mark on the future of the performing arts with these grants an act every bit as important as our traditional areas of support in health, medical research, education and the social services," Hearst Foundation President Randolph A. Hearst said Monday in announcing the grants. Lincoln Center's share will be \$200,000 for its new building, which will house young artists and provide new public facilities. The Metropolitan Opera Co. and New York Philharmonic will receive \$400,000 each. The Juilliard School, where music and the performing arts are taught, will get \$250,000. The Hearst Foundation, a leading supporter of the Lincoln Center Consolidated Corporate Fund, will make its usual annual \$100,000 donation, too.

## Why use MFVB?

- Topic discovery
  - Latent Dirichlet allocation (LDA)

"Arts"	${ m `Budgets''}$	"Children"	"Education"
NEW	MILLION	CHILDREN	SCHOOL
$\operatorname{FILM}$	TAX	WOMEN	STUDENTS
SHOW	PROGRAM	PEOPLE	SCHOOLS
MUSIC	BUDGET	CHILD	EDUCATION
MOVIE	BILLION	YEARS	TEACHERS
PLAY	FEDERAL	FAMILIES	HIGH
MUSICAL	YEAR	WORK	PUBLIC
BEST	SPENDING	PARENTS	TEACHER
ACTOR	NEW	SAYS	BENNETT
FIRST	STATE	FAMILY	MANIGAT
YORK	PLAN	WELFARE	NAMPHY
OPERA	MONEY	MEN	STATE
THEATER	PROGRAMS	PERCENT	PRESIDENT
ACTRESS	GOVERNMENT	CARE	ELEMENTARY
LOVE	CONGRESS	LIFE	HAITI

The William Randolph Hearst Foundation will give \$1.25 million to Lincoln Center, Metropolitan Opera Co., New York Philharmonic and Juilliard School. "Our board felt that we had a real opportunity to make a mark on the future of the performing arts with these grants an act every bit as important as our traditional areas of support in health, medical research, education and the social services," Hearst Foundation President Randolph A. Hearst said Monday in announcing the grants. Lincoln Center's share will be \$200,000 for its new building, which will house young artists and provide new public facilities. The Metropolitan Opera Co. and New York Philharmonic will receive \$400,000 each. The Juilliard School, where music and the performing arts are taught, will get \$250,000. The Hearst Foundation, a leading supporter of the Lincoln Center Consolidated Corporate Fund, will make its usual annual \$100,000 donation, too.

## Why use MFVB?

- Topic discovery
  - Latent Dirichlet allocation (LDA): 27,700+ citations

"Arts"	"Budgets"	"Children"	"Education"
NEW	MILLION	CHILDREN	SCHOOL
FILM	TAX	WOMEN	STUDENTS
SHOW	PROGRAM	PEOPLE	SCHOOLS
MUSIC	BUDGET	CHILD	<b>EDUCATION</b>
MOVIE	BILLION	YEARS	TEACHERS
PLAY	FEDERAL	FAMILIES	HIGH
MUSICAL	YEAR	WORK	PUBLIC
BEST	SPENDING	PARENTS	TEACHER
ACTOR	NEW	SAYS	BENNETT
FIRST	STATE	FAMILY	MANIGAT
YORK	PLAN	WELFARE	NAMPHY
OPERA	MONEY	MEN	STATE
THEATER	PROGRAMS	PERCENT	PRESIDENT
ACTRESS	GOVERNMENT	CARE	ELEMENTARY
LOVE	CONGRESS	LIFE	HAITI

The William Randolph Hearst Foundation will give \$1.25 million to Lincoln Center, Metropolitan Opera Co., New York Philharmonic and Juilliard School. "Our board felt that we had a real opportunity to make a mark on the future of the performing arts with these grants an act every bit as important as our traditional areas of support in health, medical research, education and the social services," Hearst Foundation President Randolph A. Hearst said Monday in announcing the grants. Lincoln Center's share will be \$200,000 for its new building, which will house young artists and provide new public facilities. The Metropolitan Opera Co. and New York Philharmonic will receive \$400,000 each. The Juilliard School, where music and the performing arts are taught, will get \$250,000. The Hearst Foundation, a leading supporter of the Lincoln Center Consolidated Corporate Fund, will make its usual annual \$100,000 donation, too.

### Roadmap

- Bayes & Approximate Bayes review
- What is:
  - Variational Bayes (VB)
  - Mean-field variational Bayes (MFVB)
- Why use MFVB?
- When can we trust MFVB?
- Where do we go from here?

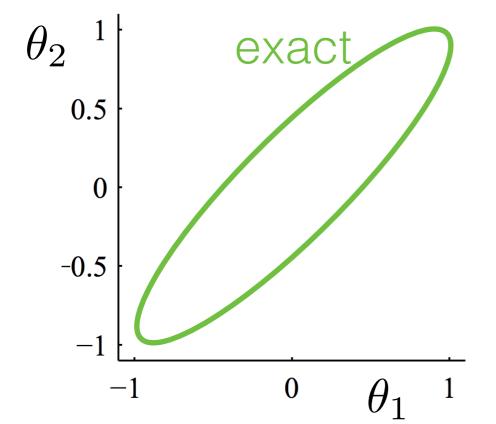
### Roadmap

- Bayes & Approximate Bayes review
- What is:
  - Variational Bayes (VB)
  - Mean-field variational Bayes (MFVB)
- Why use MFVB?
- When can we trust MFVB?
- Where do we go from here?

$$KL(q||p(\cdot|y)) = \int_{\theta} q(\theta) \log \frac{q(\theta)}{p(\theta|y)} d\theta$$

$$q(\theta) = \prod_{j=1}^{J} q_j(\theta_j)$$

$$KL(q||p(\cdot|y)) = \int_{\theta} q(\theta) \log \frac{q(\theta)}{p(\theta|y)} d\theta$$

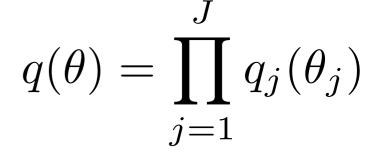


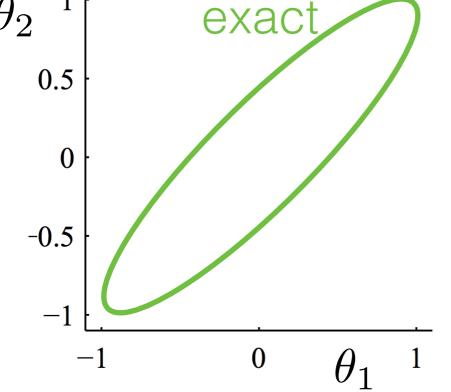
$$q(\theta) = \prod_{j=1}^{J} q_j(\theta_j)$$

[Turner & Sahani 2011; MacKay 2003; Bishop 2006; Wang, Titterington 2004]

$$KL(q||p(\cdot|y)) = \int_{\theta} q(\theta) \log \frac{q(\theta)}{p(\theta|y)} d\theta$$

$$KL(q||p(\cdot|y)) = \int_{\theta} q(\theta) \log \frac{q(\theta)}{p(\theta|y)} d\theta$$





[Turner & Sahani 2011; MacKay 2003; Bishop 2006; Wang, Titterington 2004]

Conjugate linear regression

$$KL(q||p(\cdot|y)) = \int_{\theta} q(\theta) \log \frac{q(\theta)}{p(\theta|y)} d\theta$$

$$\theta_2$$
 exact

0.5

 $\theta_2$ 
 $\theta_2$ 
 $\theta_2$ 
 $\theta_2$ 
 $\theta_2$ 
 $\theta_3$ 
 $\theta_4$ 
 $\theta_4$ 
 $\theta_4$ 
 $\theta_4$ 
 $\theta_4$ 
 $\theta_4$ 
 $\theta_4$ 

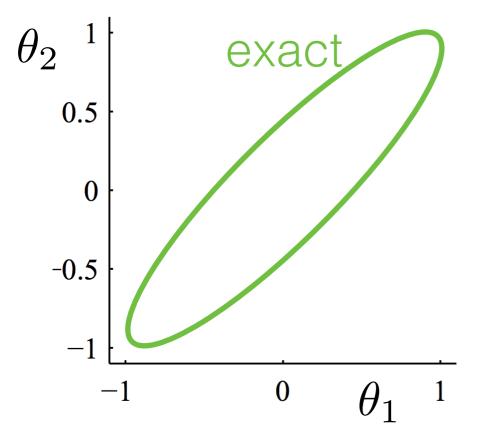
$$q(\theta) = \prod_{j=1}^{J} q_j(\theta_j)$$

[Turner & Sahani 2011; MacKay 2003; Bishop 2006; Wang, Titterington 2004]

- Conjugate linear regression
- Bayesian central limit theorem

$$KL(q||p(\cdot|y)) = \int_{\theta} q(\theta) \log \frac{q(\theta)}{p(\theta|y)} d\theta$$

$$q(\theta) = \prod_{j=1}^{J} q_j(\theta_j)$$

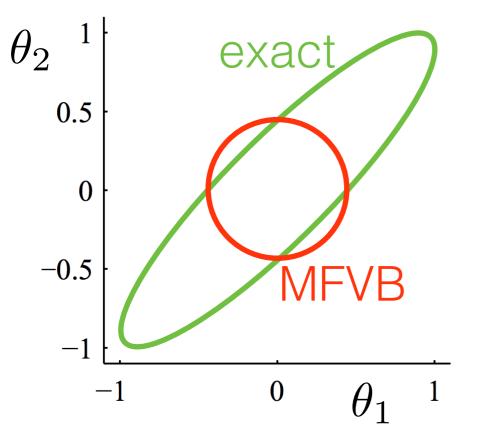


[Turner & Sahani 2011; MacKay 2003; Bishop 2006; Wang, Titterington 2004]

- Conjugate linear regression
- Bayesian central limit theorem

$$KL(q||p(\cdot|y)) = \int_{\theta} q(\theta) \log \frac{q(\theta)}{p(\theta|y)} d\theta$$

$$q(\theta) = \prod_{j=1}^{J} q_j(\theta_j)$$

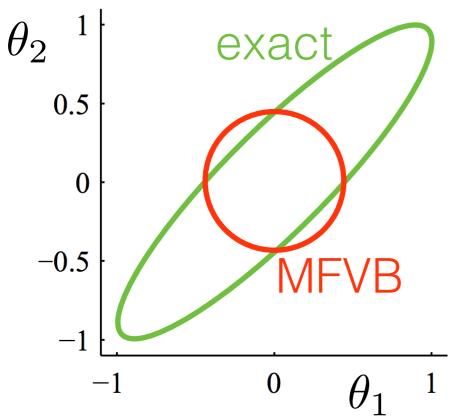


[Turner & Sahani 2011; MacKay 2003; Bishop 2006; Wang, Titterington 2004]

- Conjugate linear regression
- Bayesian central limit theorem

$$KL(q||p(\cdot|y)) = \int_{\theta} q(\theta) \log \frac{q(\theta)}{p(\theta|y)} d\theta$$

$$q(\theta) = \prod_{j=1}^{J} q_j(\theta_j)$$



[Turner & Sahani 2011; MacKay 2003; Bishop 2006; Wang, Titterington 2004]

Underestimates variance (sometimes severely)

- Conjugate linear regression
- Bayesian central limit theorem

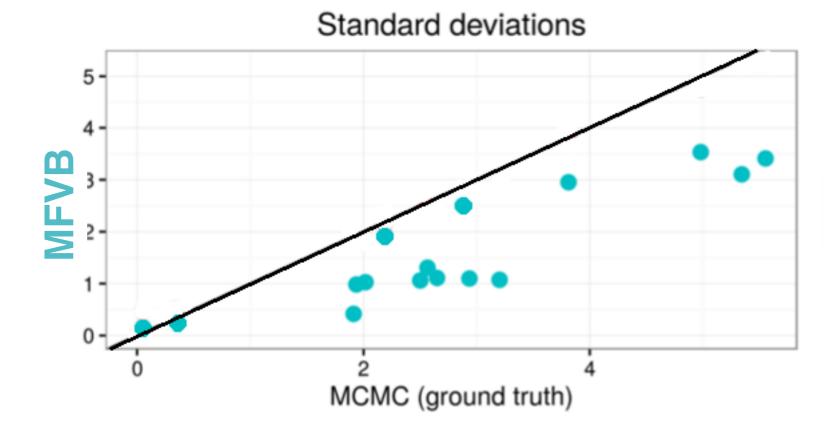
Underestimates variance (sometimes severely)

- Conjugate linear regression
- Bayesian central limit theorem

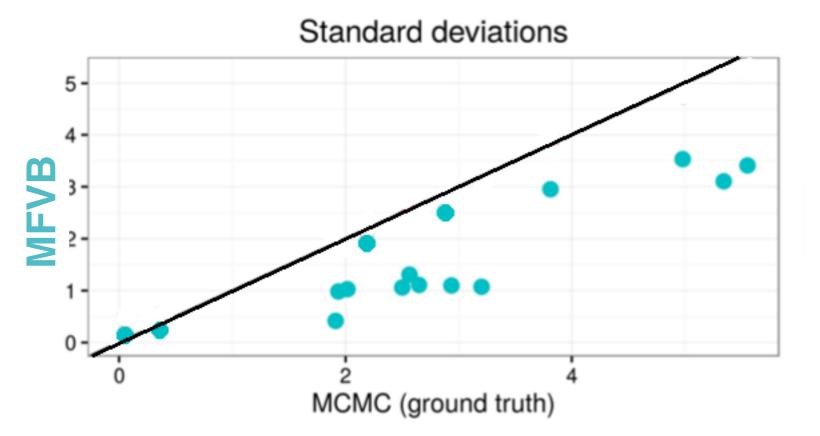
- Underestimates variance (sometimes severely)
- No covariance estimates
- Conjugate linear regression
- Bayesian central limit theorem

Microcredit

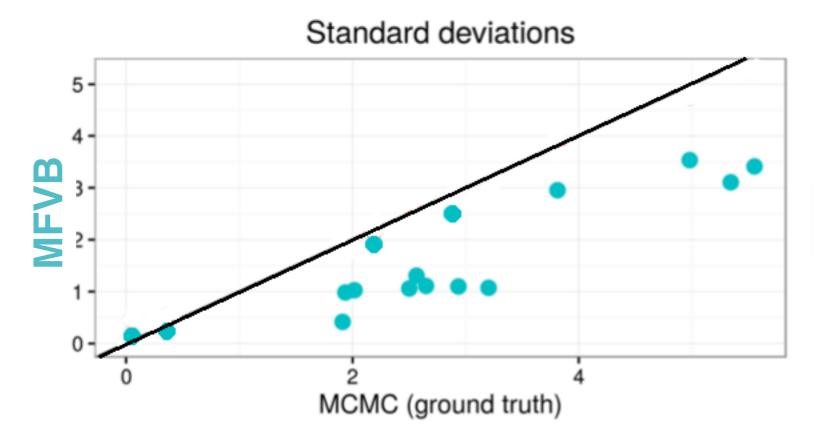
Microcredit



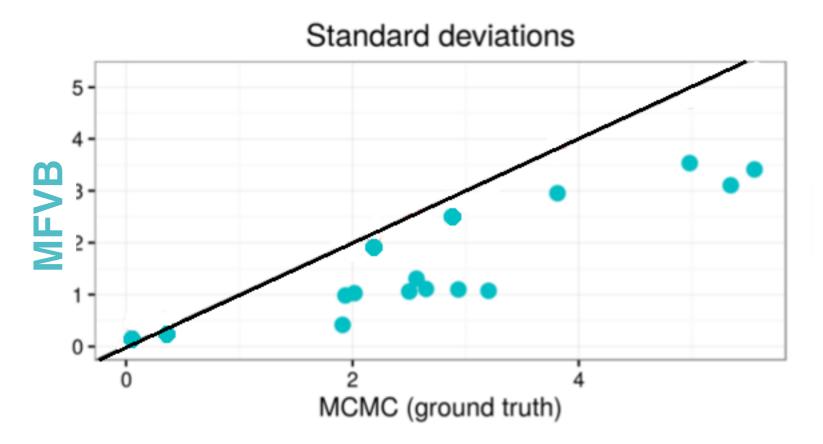
- Microcredit effect
- t mean:3.08 USD PPP



- Microcredit effect
- t mean:3.08 USD PPP
- τ std dev:1.83 USD PPP

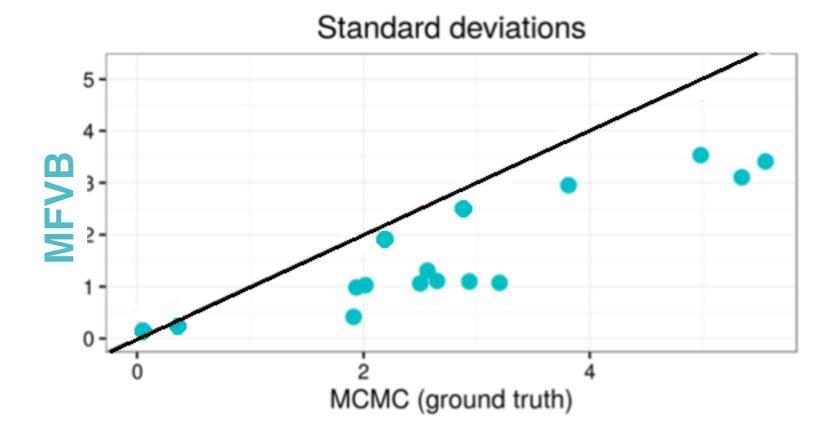


- Microcredit effect
- t mean:3.08 USD PPP
- *t* std dev:
   1.83 USD PPP
- Mean is 1.68 std dev from 0

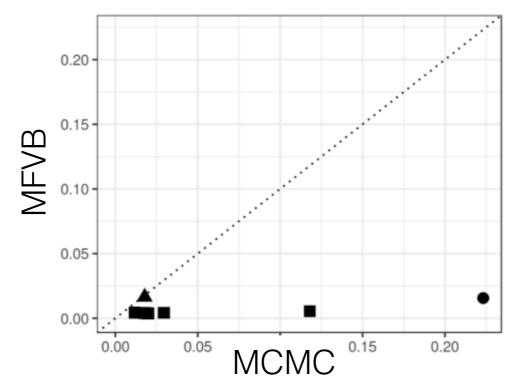


- Microcredit effect
- t mean:3.08 USD PPP
- *t* std dev:
   1.83 USD PPP
- Mean is 1.68 std dev from 0

Criteo
online ads
experiment



#### Standard deviations

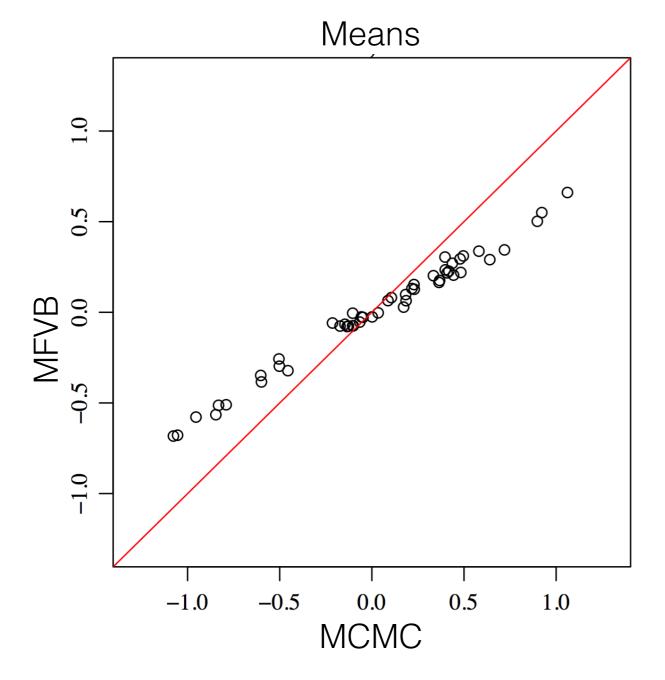


#### What about means?

- Model for relational data with covariates
- When 1000+ nodes, MCMC > 1 day [Fosdick 2013, Ch 4]

#### What about means?

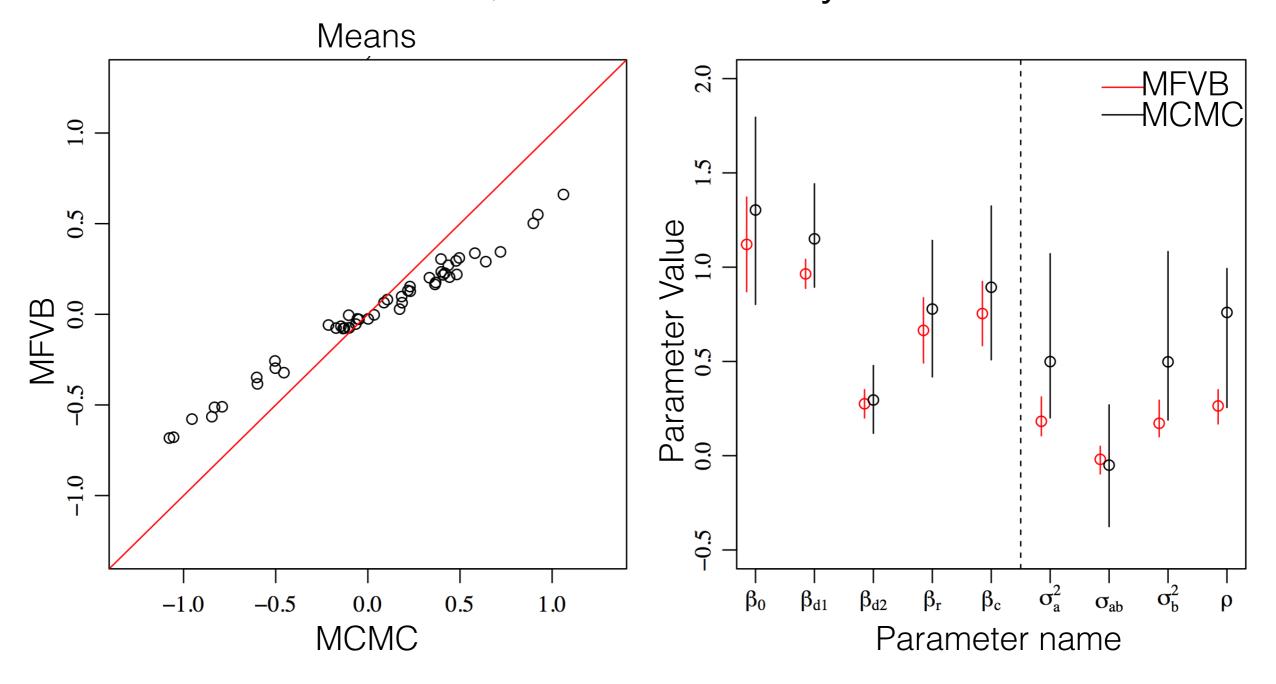
- Model for relational data with covariates
- When 1000+ nodes, MCMC > 1 day [Fosdick 2013, Ch 4]



[Fosdick 2013, Ch 4, Fig 4.3]

#### What about means?

- Model for relational data with covariates
- When 1000+ nodes, MCMC > 1 day [Fosdick 2013, Ch 4]



[Fosdick 2013, Ch 4, Fig 4.3]

#### Posterior means: revisited

• Want to predict college GPA  $y_n$ 

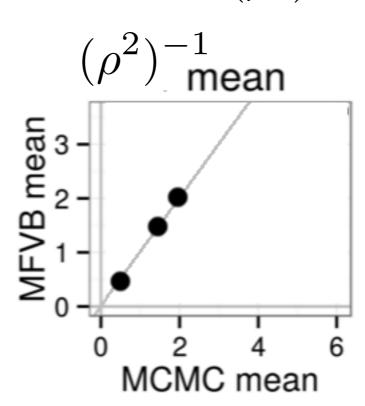
- Want to predict college GPA  $y_n$
- Collect: standardized test scores (e.g., SAT, ACT)  $x_n$

- Want to predict college GPA  $y_n$
- Collect: standardized test scores (e.g., SAT, ACT)  $x_n$
- Collect: regional test scores  $r_n$

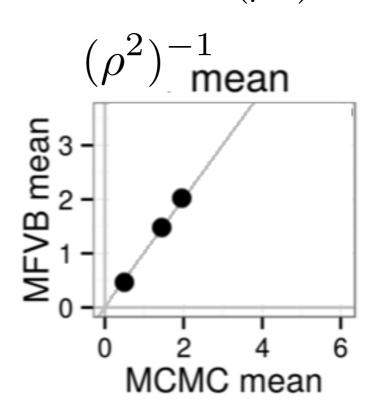
- Want to predict college GPA  $y_n$
- Collect: standardized test scores (e.g., SAT, ACT)  $x_n$
- Collect: regional test scores  $r_n$
- Model:  $y_n | \beta, z, \sigma^2 \stackrel{indep}{\sim} \mathcal{N}(\beta^T x_n + z_{k(n)} r_n, \sigma^2)$

- Want to predict college GPA  $y_n$
- Collect: standardized test scores (e.g., SAT, ACT)  $x_n$
- Collect: regional test scores  $r_n$
- Model:  $y_n | \beta, z, \sigma^2 \stackrel{indep}{\sim} \mathcal{N}(\beta^T x_n + z_{k(n)} r_n, \sigma^2)$   $z_k | \rho^2 \stackrel{iid}{\sim} \mathcal{N}(0, \rho^2) \qquad (\sigma^2)^{-1} \sim \operatorname{Gamma}(a_{\sigma^2}, b_{\sigma^2})$  $\beta \sim \mathcal{N}(0, \Sigma) \qquad (\rho^2)^{-1} \sim \operatorname{Gamma}(a_{\rho^2}, b_{\rho^2})$

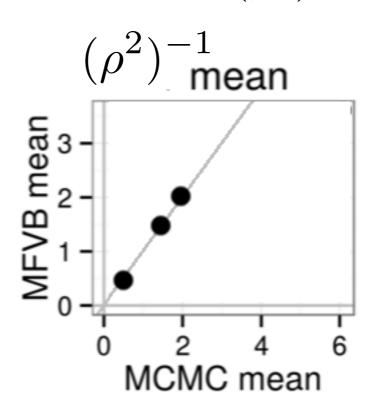
- Want to predict college GPA  $y_n$
- Collect: standardized test scores (e.g., SAT, ACT)  $x_n$
- Collect: regional test scores  $r_n$
- Model:  $y_n | \beta, z, \sigma^2 \stackrel{indep}{\sim} \mathcal{N}(\beta^T x_n + z_{k(n)} r_n, \sigma^2)$   $z_k | \rho^2 \stackrel{iid}{\sim} \mathcal{N}(0, \rho^2) \qquad (\sigma^2)^{-1} \sim \operatorname{Gamma}(a_{\sigma^2}, b_{\sigma^2})$   $\beta \sim \mathcal{N}(0, \Sigma) \qquad (\rho^2)^{-1} \sim \operatorname{Gamma}(a_{\rho^2}, b_{\rho^2})$ 
  - Data simulated from model (3 data sets, 300 data points):



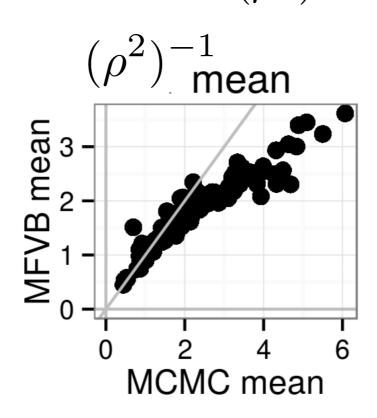
- Want to predict college GPA  $y_n$
- Collect: standardized test scores (e.g., SAT, ACT)  $x_n$
- Collect: regional test scores  $r_n$
- Model:  $y_n | \beta, z, \sigma^2 \stackrel{indep}{\sim} \mathcal{N}(\beta^T x_n + z_{k(n)} r_n, \sigma^2)$   $z_k | \rho^2 \stackrel{iid}{\sim} \mathcal{N}(0, \rho^2) \qquad (\sigma^2)^{-1} \sim \operatorname{Gamma}(a_{\sigma^2}, b_{\sigma^2})$   $\beta \sim \mathcal{N}(0, \Sigma) \qquad (\rho^2)^{-1} \sim \operatorname{Gamma}(a_{\rho^2}, b_{\rho^2})$ 
  - Data simulated from model (3) data sets, 300 data points):



- Want to predict college GPA  $y_n$
- Collect: standardized test scores (e.g., SAT, ACT)  $x_n$
- Collect: regional test scores  $r_n$
- Model:  $y_n | \beta, z, \sigma^2 \stackrel{indep}{\sim} \mathcal{N}(\beta^T x_n + z_{k(n)} r_n, \sigma^2)$   $z_k | \rho^2 \stackrel{iid}{\sim} \mathcal{N}(0, \rho^2) \qquad (\sigma^2)^{-1} \sim \operatorname{Gamma}(a_{\sigma^2}, b_{\sigma^2})$   $\beta \sim \mathcal{N}(0, \Sigma) \qquad (\rho^2)^{-1} \sim \operatorname{Gamma}(a_{\rho^2}, b_{\rho^2})$ 
  - Data simulated from model (100 data sets, 300 data points):



- Want to predict college GPA  $y_n$
- Collect: standardized test scores (e.g., SAT, ACT)  $x_n$
- Collect: regional test scores  $r_n$
- Model:  $y_n | \beta, z, \sigma^2 \stackrel{indep}{\sim} \mathcal{N}(\beta^T x_n + z_{k(n)} r_n, \sigma^2)$   $z_k | \rho^2 \stackrel{iid}{\sim} \mathcal{N}(0, \rho^2) \qquad (\sigma^2)^{-1} \sim \operatorname{Gamma}(a_{\sigma^2}, b_{\sigma^2})$   $\beta \sim \mathcal{N}(0, \Sigma) \qquad (\rho^2)^{-1} \sim \operatorname{Gamma}(a_{\rho^2}, b_{\rho^2})$ 
  - Data simulated from model (100 data sets, 300 data points):



Use  $q^*$  to approximate  $p(\cdot|y)$ 

Optimization 
$$q^* = \operatorname{argmin}_{q \in Q} f(q(\cdot), p(\cdot|y))$$

Variational Bayes  $q^* = \mathrm{argmin}_{q \in Q} KL(q(\cdot)||p(\cdot|y))$ 

How deep is the issue?

Mean-field variational Bayes

$$q^* = \operatorname{argmin}_{q \in Q_{\text{MFVB}}} KL(q(\cdot)||p(\cdot|y))$$

- Coordinate descent
- Stochastic variational inference (\$VI) [Hoffman et al 2013]
- Automatic differentiation variational inference (ADVI) [Kucukelbir et al 2015, 2017]

Use  $q^*$  to approximate  $p(\cdot|y)$ 

Optimization 
$$q^* = \operatorname{argmin}_{q \in Q} f(q(\cdot), p(\cdot|y))$$

Variational Bayes  $q^* = \operatorname{argmin}_{q \in Q} KL(q(\cdot)||p(\cdot|y))$ 

How deep is the issue?

Mean-field variational Bayes

$$q^* = \operatorname{argmin}_{q \in Q_{\text{MFVB}}} KL(q(\cdot)||p(\cdot|y))$$

Algorithm

Use  $q^*$  to approximate  $p(\cdot|y)$ 

Variational Bayes  $q^* = \mathrm{argmin}_{q \in Q} KL(q(\cdot)||p(\cdot|y))$ 

How deep is the issue?

Mean-field variational Bayes

$$q^* = \operatorname{argmin}_{q \in Q_{\text{MFVB}}} KL(q(\cdot)||p(\cdot|y))$$

Algorithm

Implementation

Use  $q^*$  to approximate  $p(\cdot|y)$ 

Variational Bayes  $q^* = \operatorname{argmin}_{q \in Q} KL(q(\cdot)||p(\cdot|y))$ 

How deep is the issue?

Mean-field variational Bayes

$$q^* = \operatorname{argmin}_{q \in Q_{\text{MFVB}}} KL(q(\cdot)||p(\cdot|y))$$

Algorithm

**Implementation** 

Use  $q^*$  to approximate  $p(\cdot|y)$ 

Variational Bayes

$$q^* = \operatorname{argmin}_{q \in Q} KL(q(\cdot)||p(\cdot|y))$$

How deep is the issue?

Mean-field variational Bayes

$$q^* = \operatorname{argmin}_{q \in Q_{\text{MFVB}}} KL(q(\cdot)||p(\cdot|y))$$

Algorithm

**Implementation** 

Gaussian example was exact

Use  $q^*$  to approximate  $p(\cdot|y)$ 

Variational Bayes

$$q^* = \operatorname{argmin}_{q \in Q} KL(q(\cdot)||p(\cdot|y))$$

How deep is the issue?

Mean-field variational Bayes

$$q^* = \operatorname{argmin}_{q \in Q_{\text{MFVB}}} KL(q(\cdot)||p(\cdot|y))$$

#### **Algorithm**

Implementation

Gaussian example was exact

Use  $q^*$  to approximate  $p(\cdot|y)$ 

Variational Bayes  $q^* = \operatorname{argmin}_{q \in Q} KL(q(\cdot)||p(\cdot|y))$ 

# How deep is the issue?

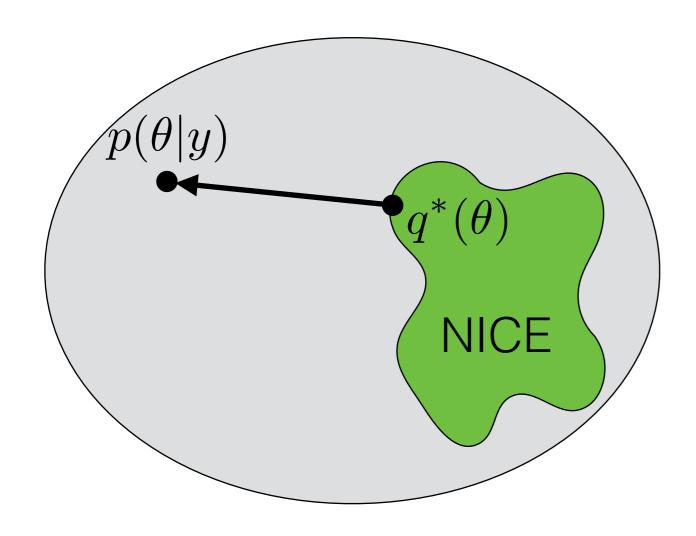
#### Mean-field variational Bayes

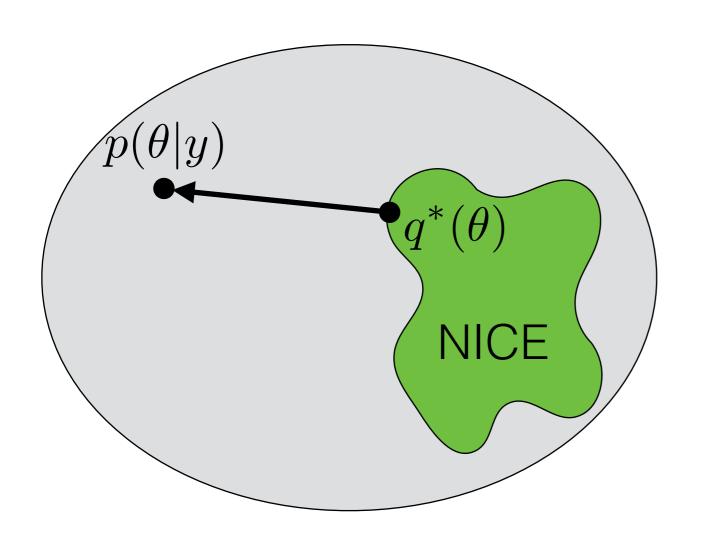
$$q^* = \operatorname{argmin}_{q \in Q_{\text{MFVB}}} KL(q(\cdot)||p(\cdot|y))$$

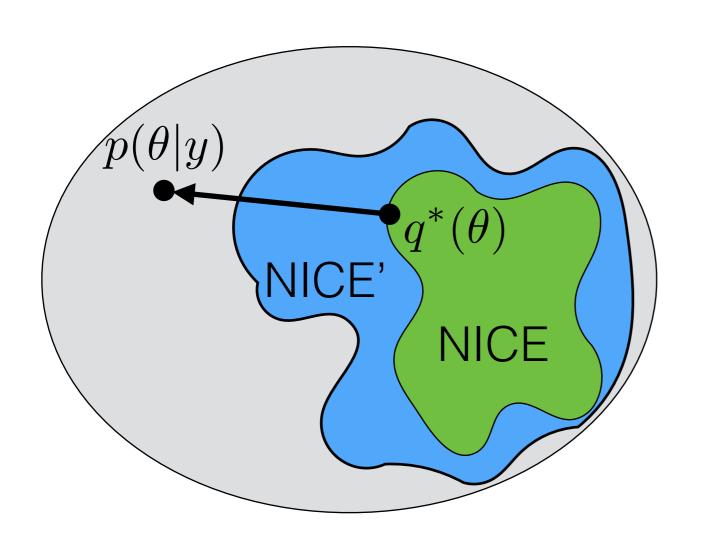
Algorithm

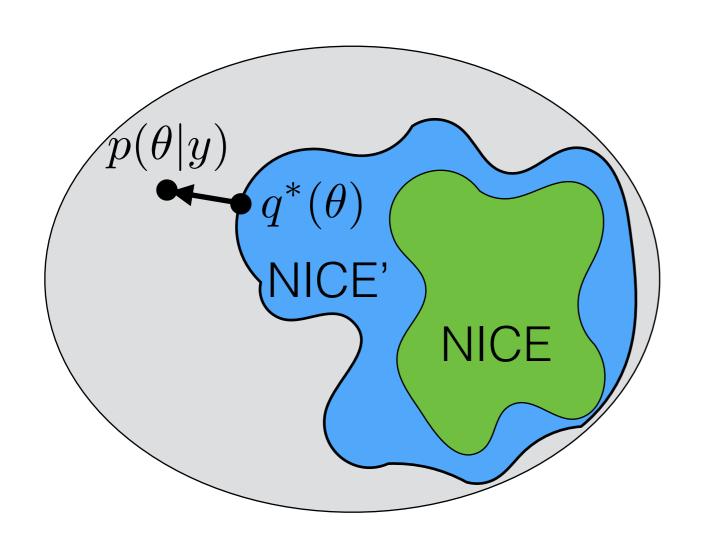
Implementation

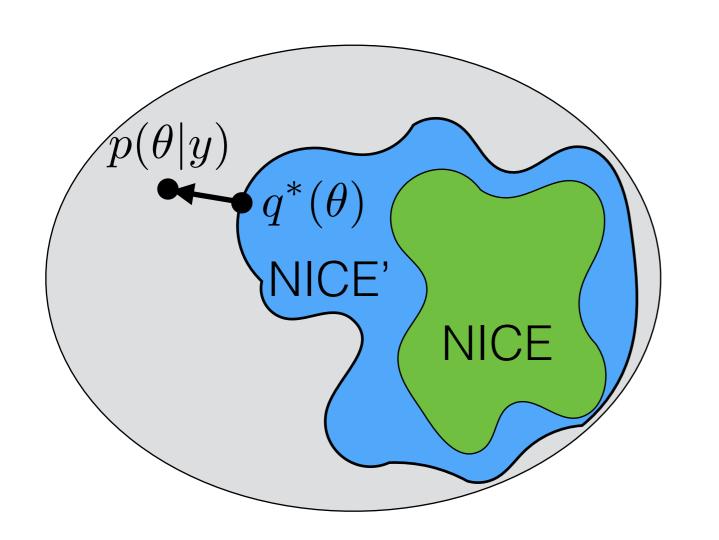
Gaussian example was exact



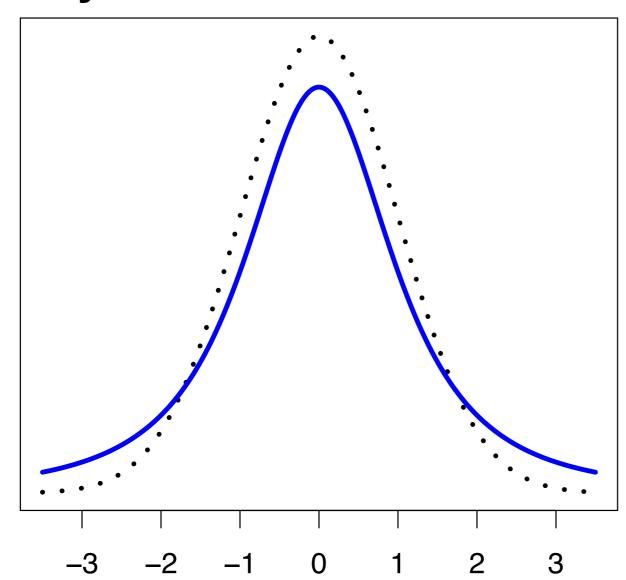


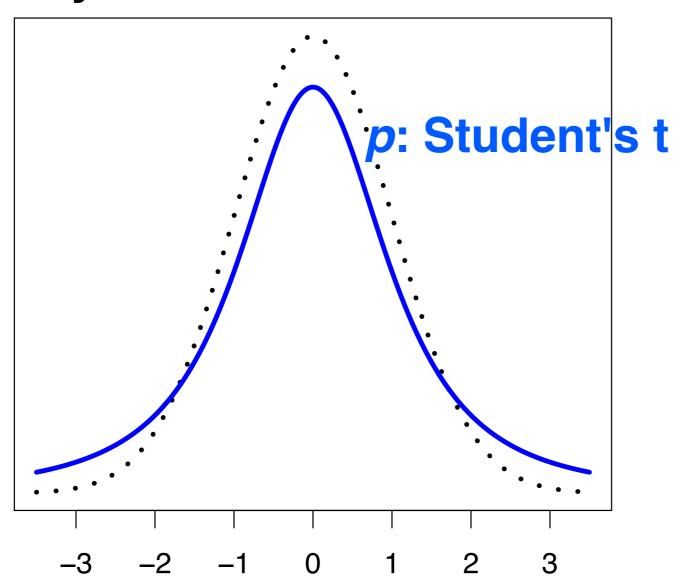


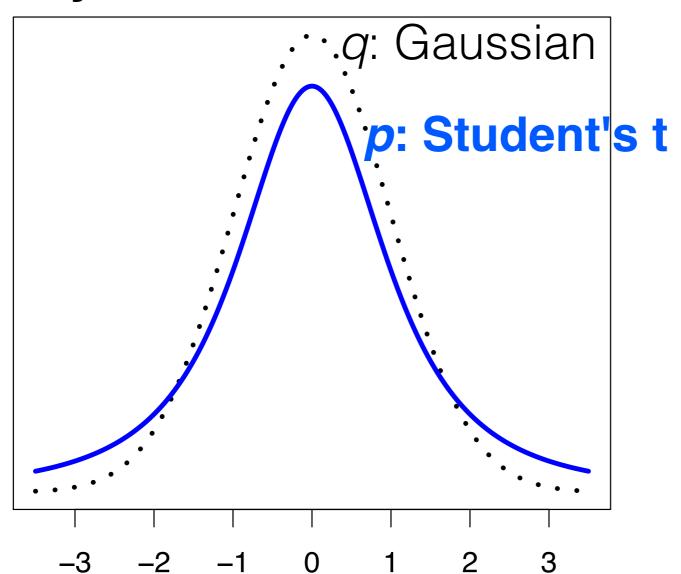


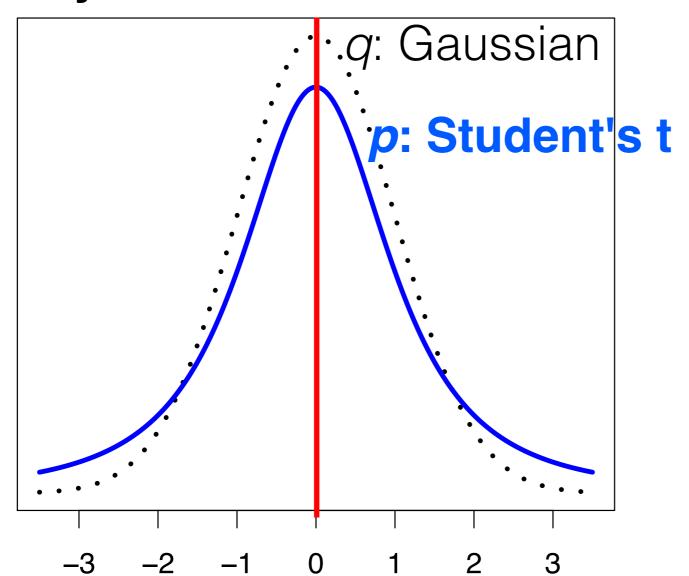


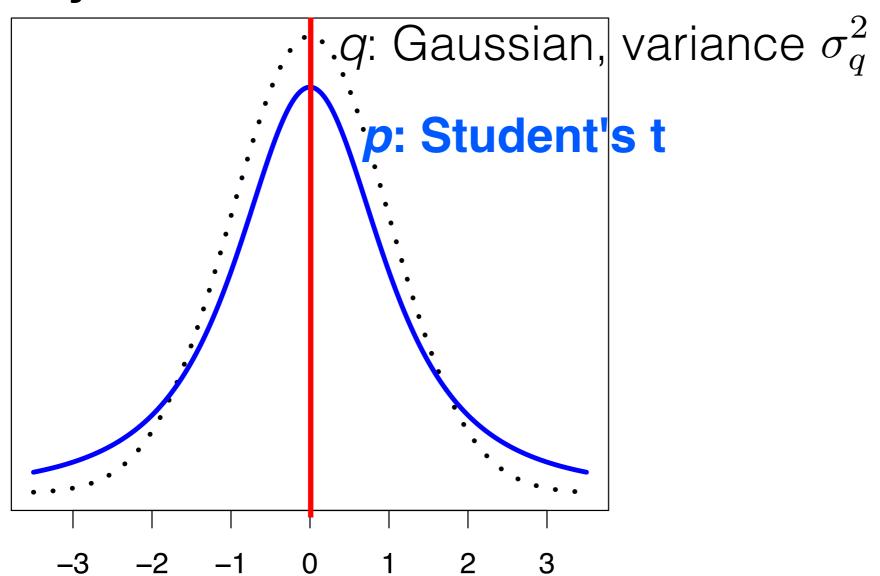
- Takeaway: A smaller KL does not imply better mean and variance estimates
- Exercise: show this

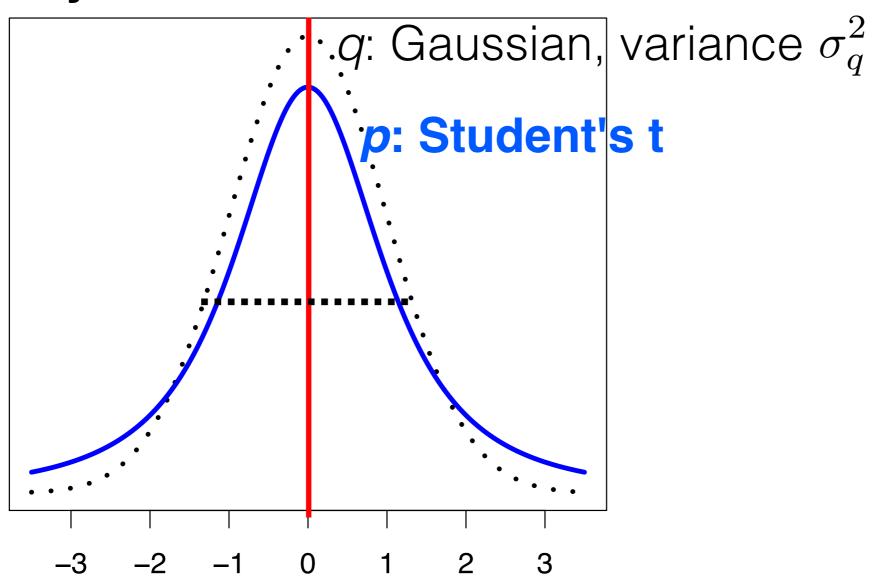


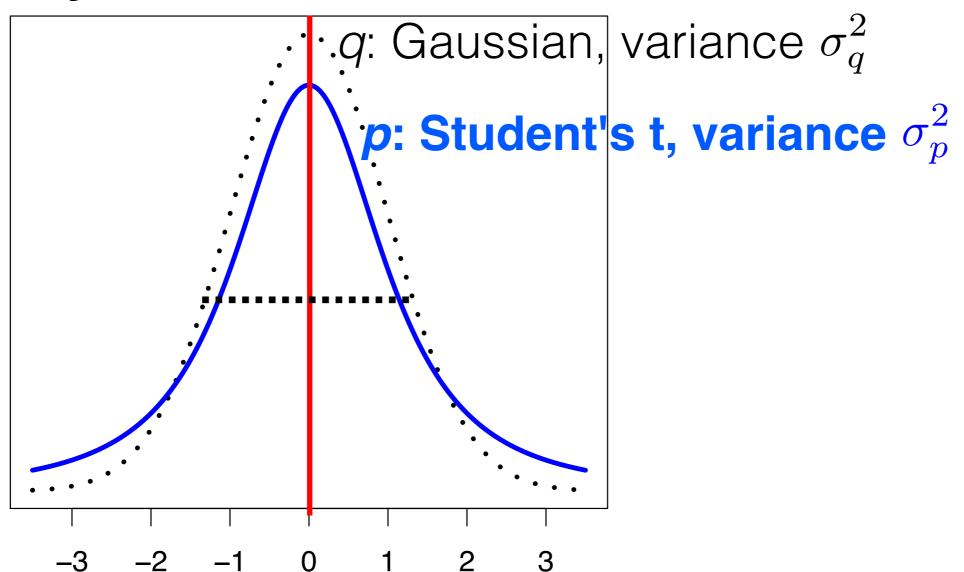


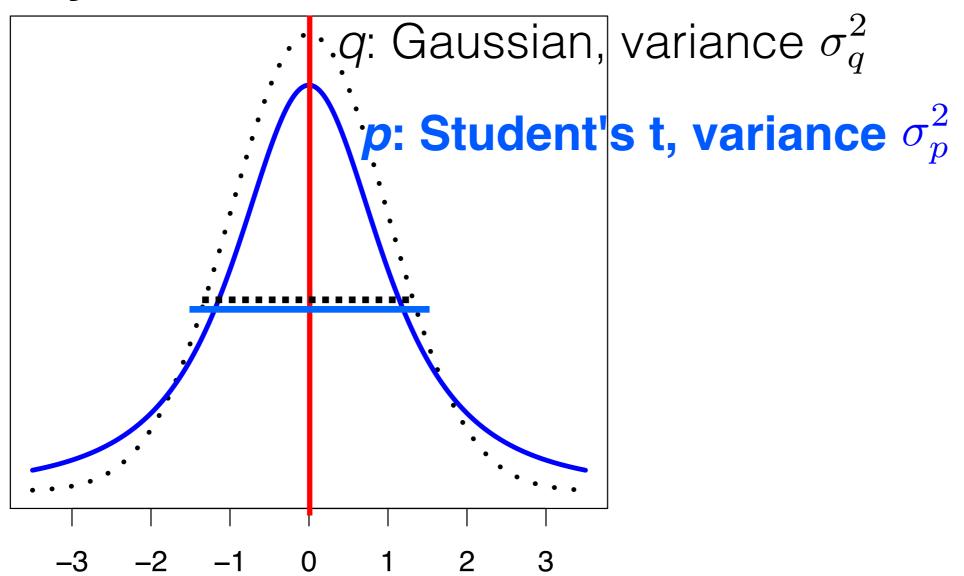


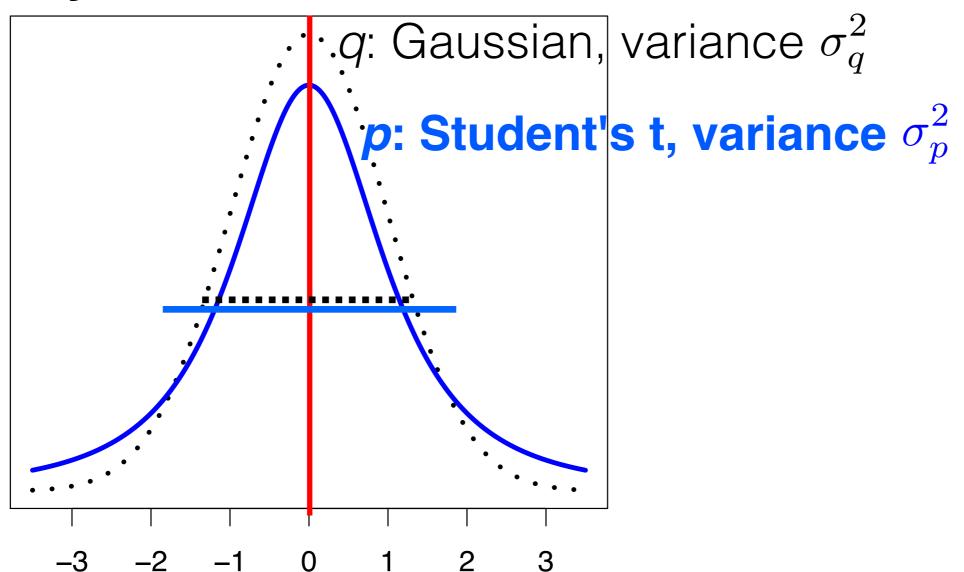


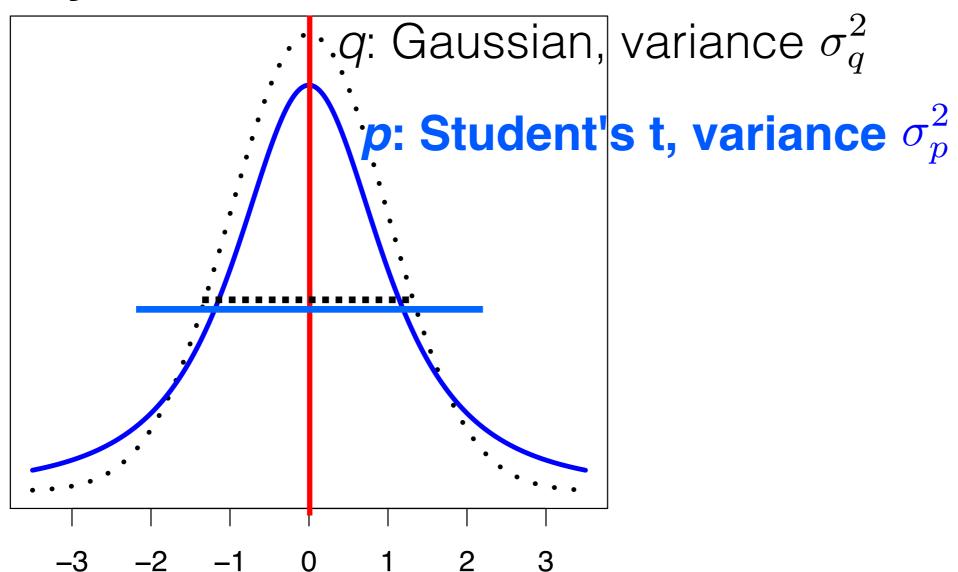


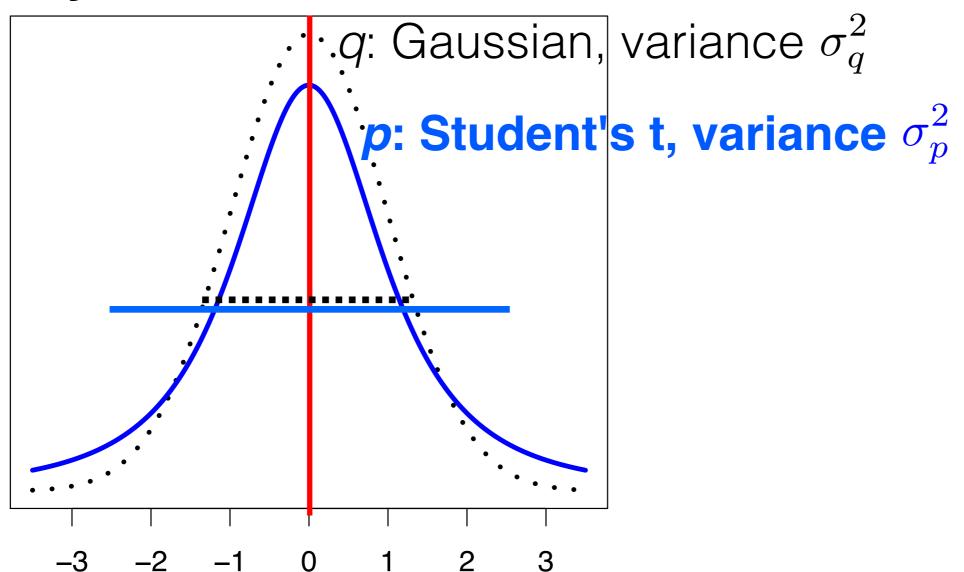


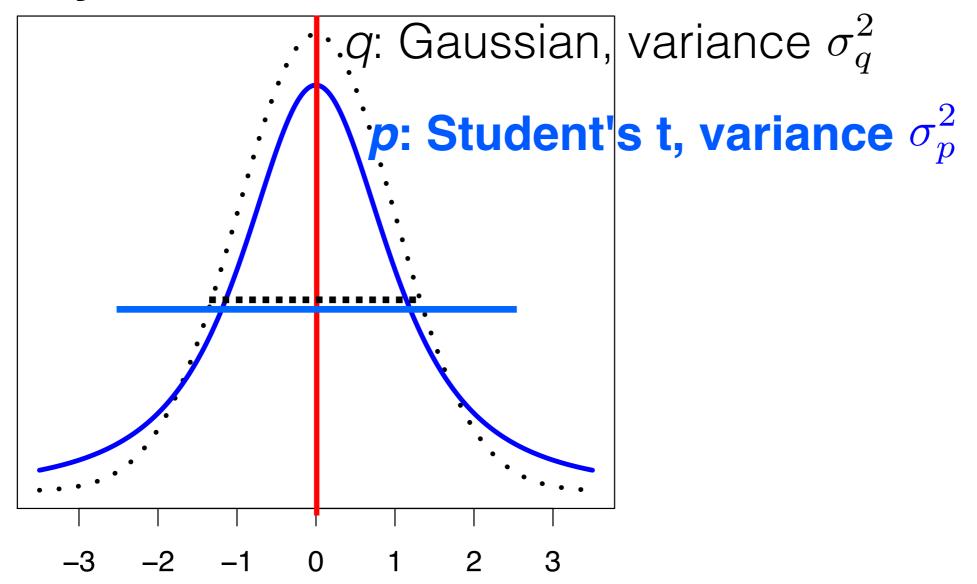




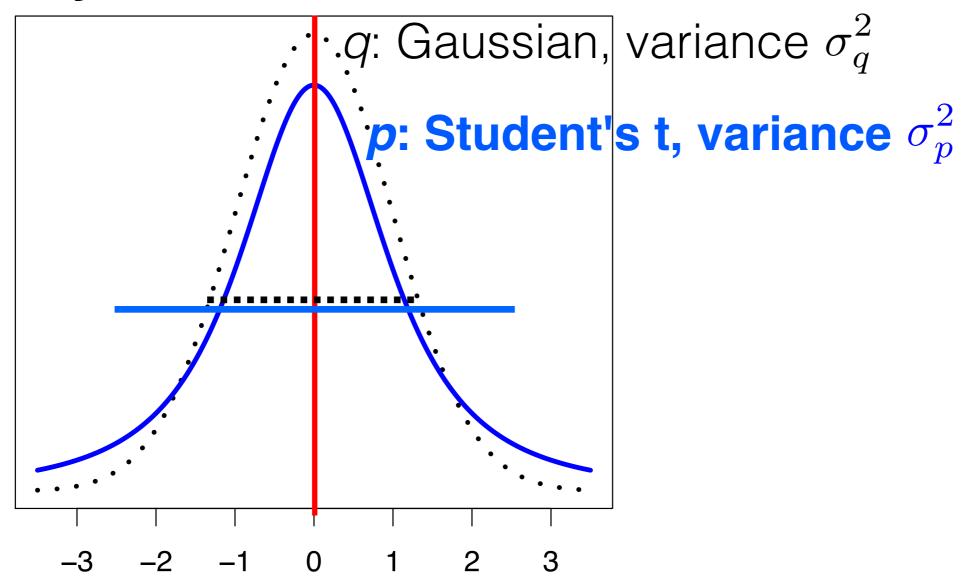




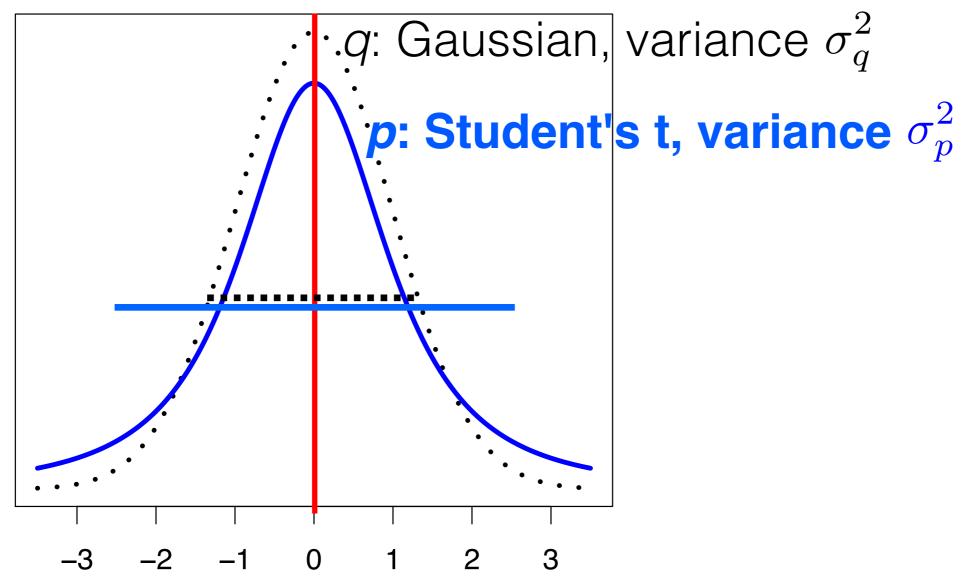




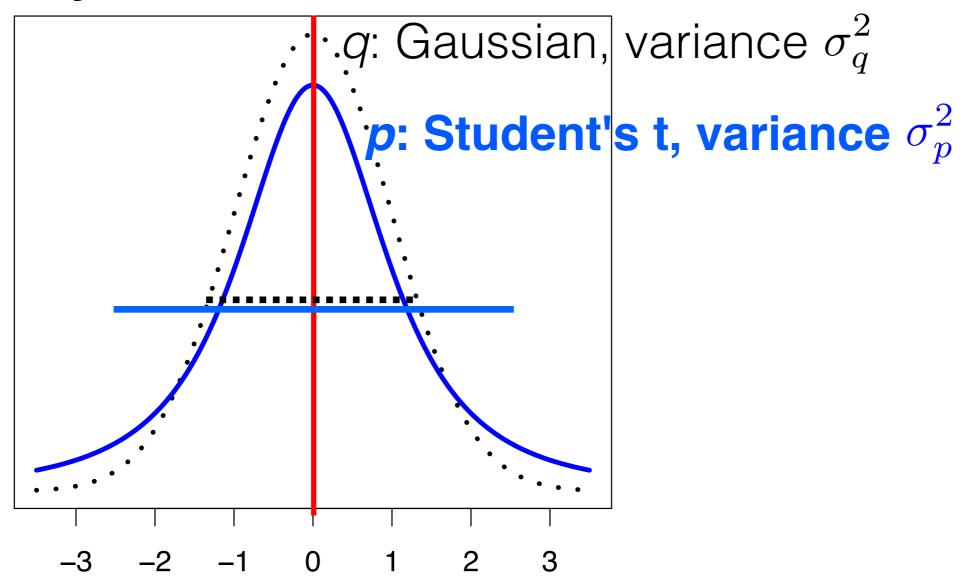
$$\sigma_p^2 \ge c\sigma_q^2$$



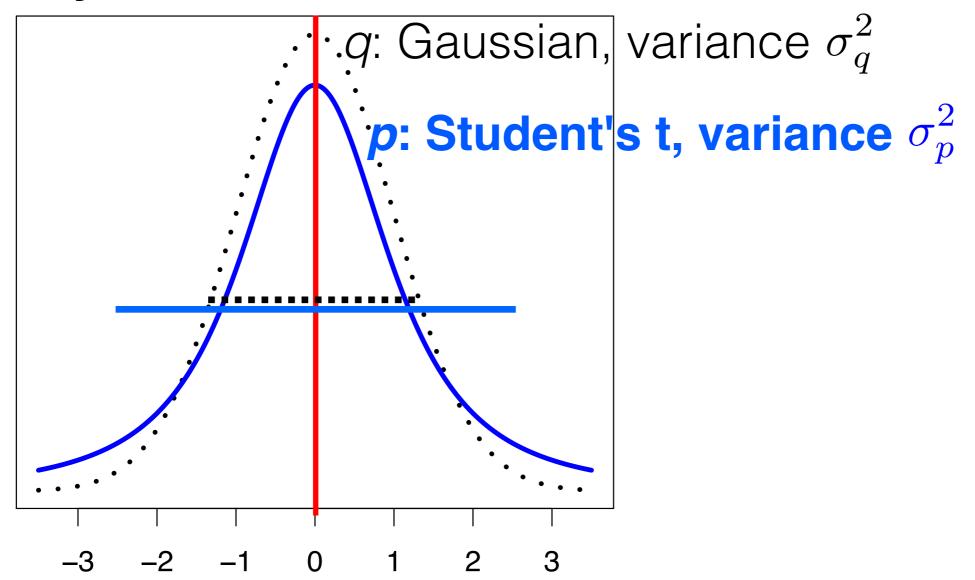
$$KL(q||p) < 0.802$$
 but also  $\sigma_p^2 \ge c\sigma_q^2$ 



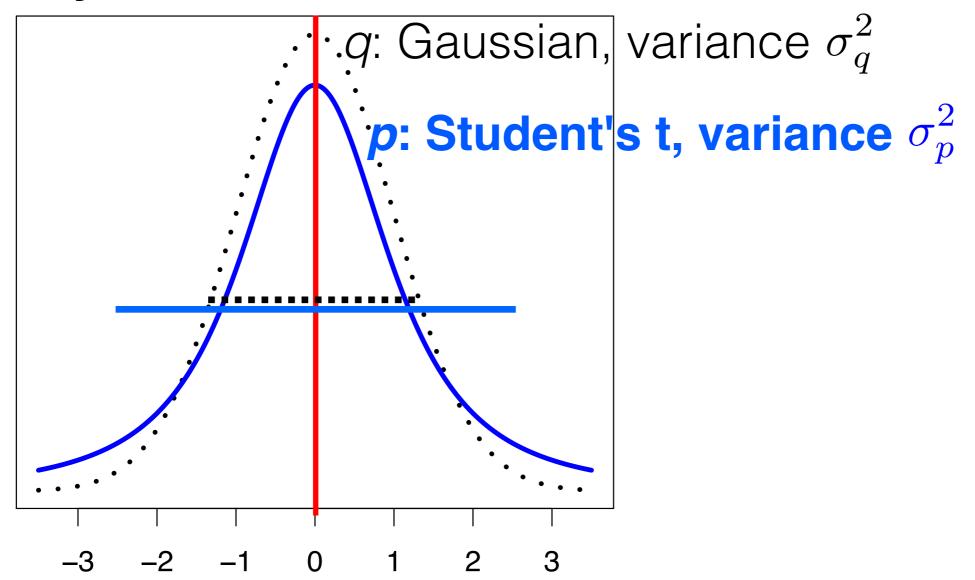
**Proposition (HKCB)**. For any c>1, there exist zeromean, unimodal distributions q and p such that KL(q||p)<0.802 but also  $\sigma_p^2\geq c\sigma_q^2$ 



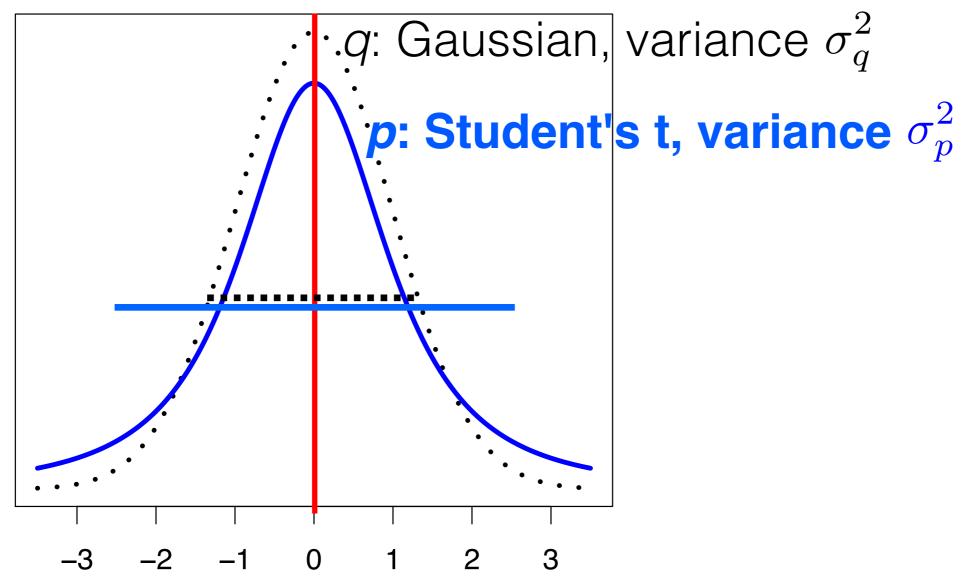
**Proposition (HKCB)**. For any c>1, there exist zeromean, unimodal distributions q and p such that KL(q||p)<0.802 but also  $\sigma_p^2\geq c\sigma_q^2$ 



**Proposition (HKCB).** For any c > 1, there exist zeromean, unimodal distributions q and p such that KL(q||p) < 0.802 but also  $\sigma_p^2 \ge c\sigma_q^2$ 

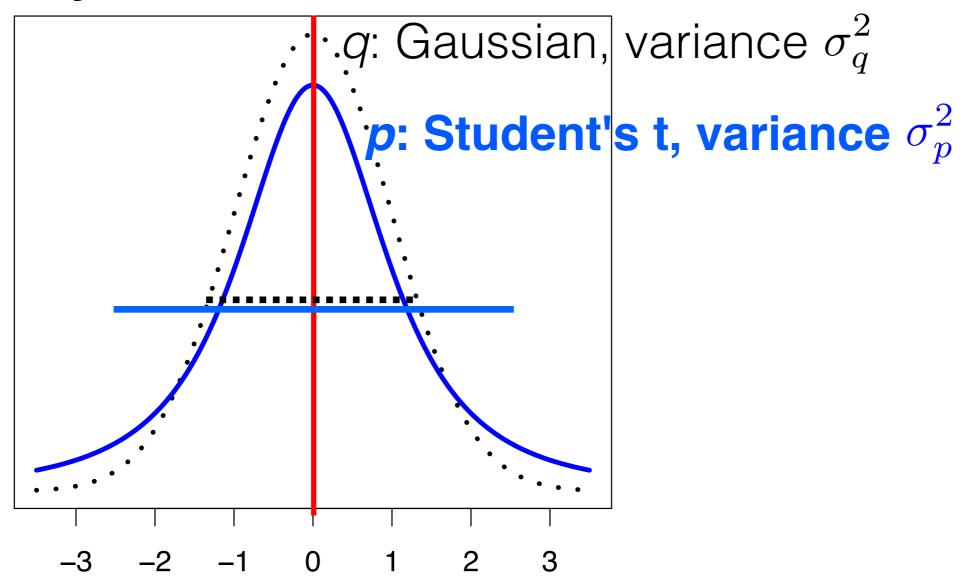


**Conjecture (HKCB)**. For any c > 1, there exist zeromean, unimodal distributions q and p such that KL(q||p) < 0.802 but also  $\sigma_p^2 \ge c\sigma_q^2$ 



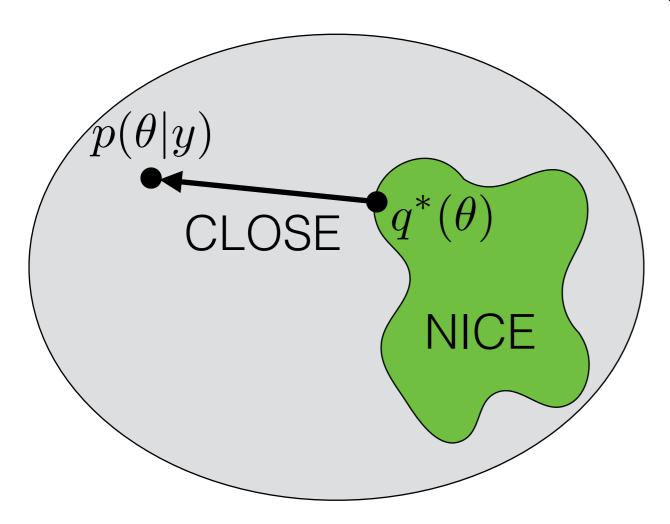
Conjecture (HKCB). For any c > 1, there exist zeromean, unimodal distributions q and p such that

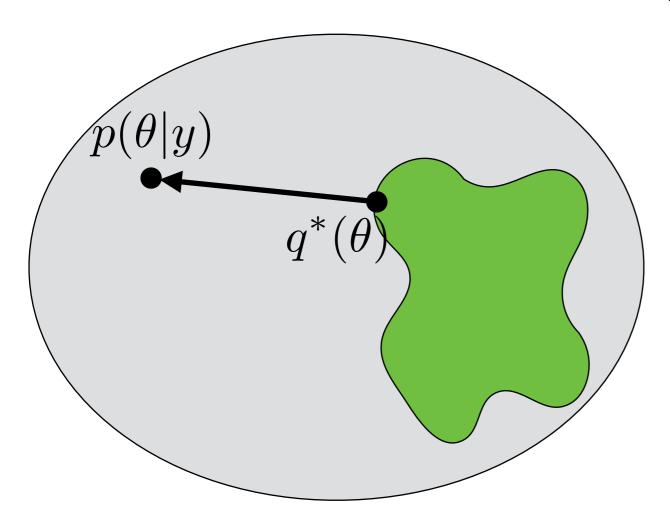
$$KL(q||\mathbf{p}) < 0.802$$
 but also  $\sigma_{\mathbf{p}}^2 \geq c\sigma_q^2$ 

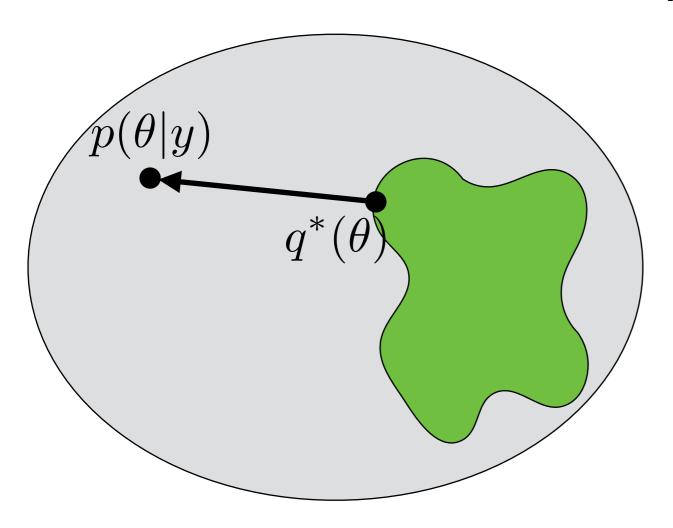


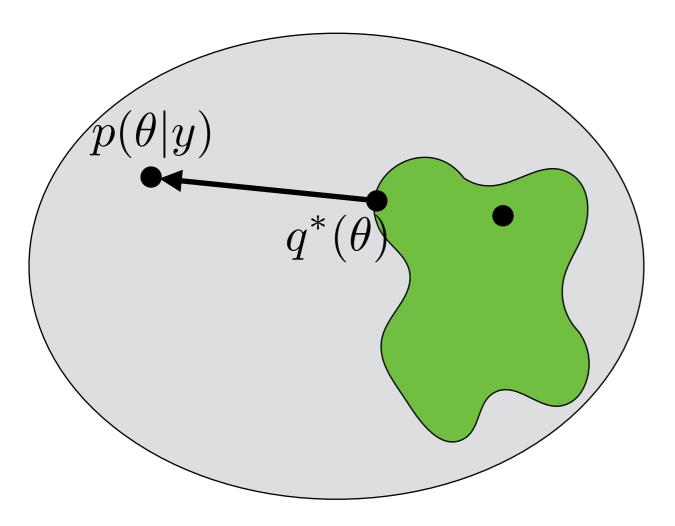
Conjecture (HKCB). For any c > 1, there exist zeromean, unimodal distributions q and p such that

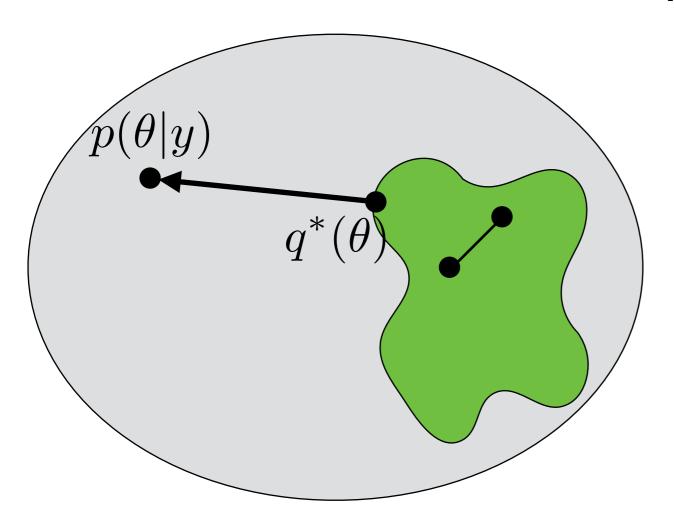
$$KL(q||\mathbf{p}) < \mathbf{0.12}$$
 but also  $\sigma_{\mathbf{p}}^2 \geq c\sigma_q^2$ 

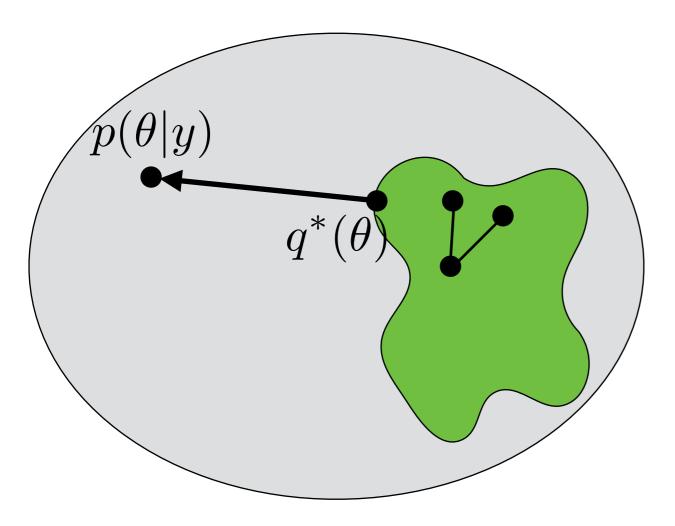


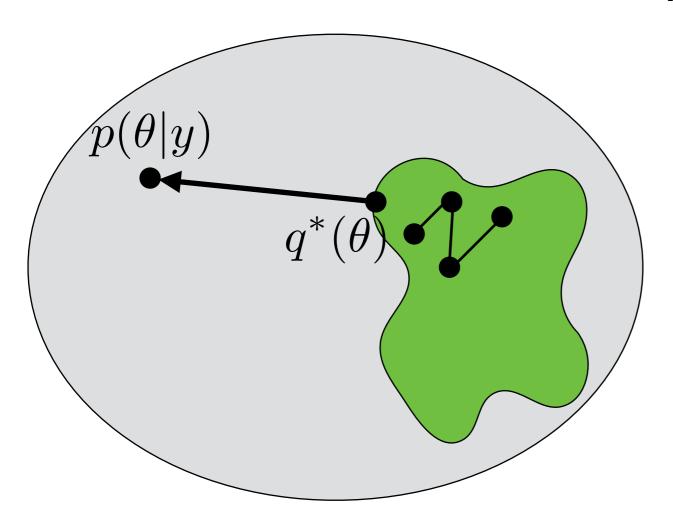


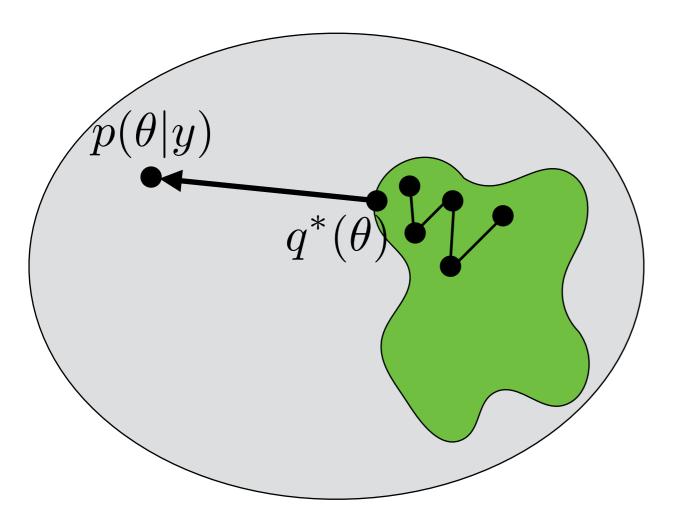


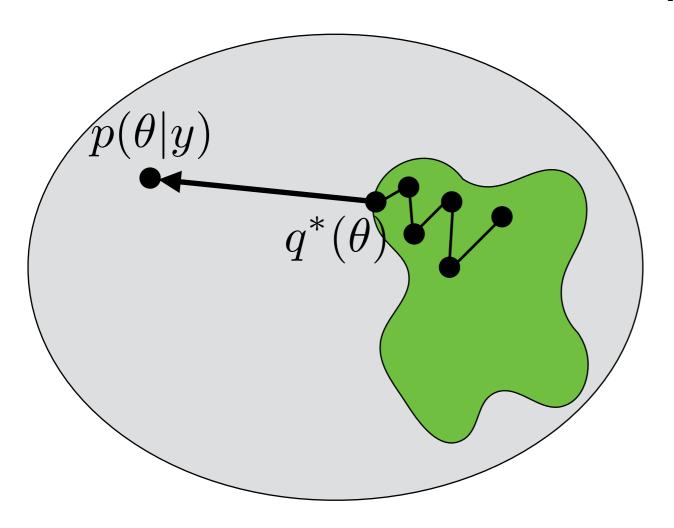


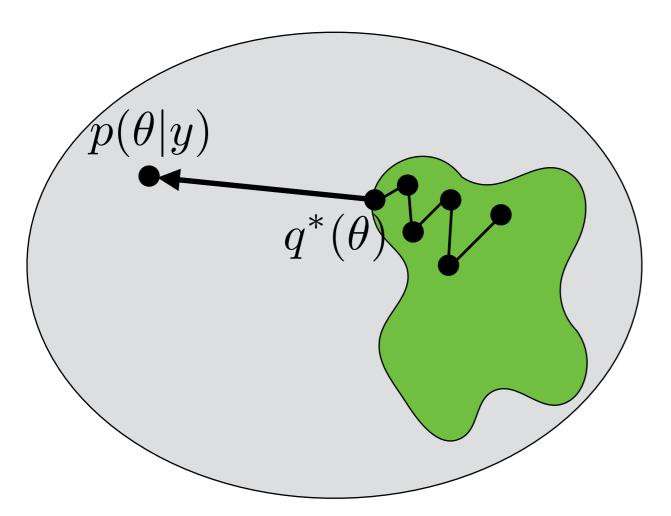




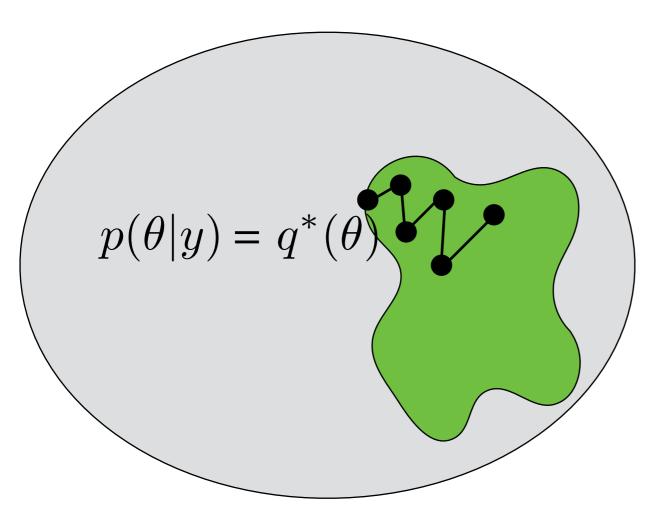




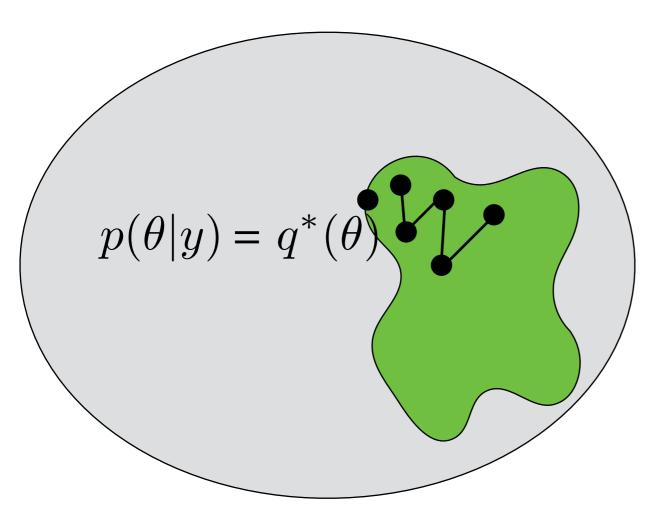




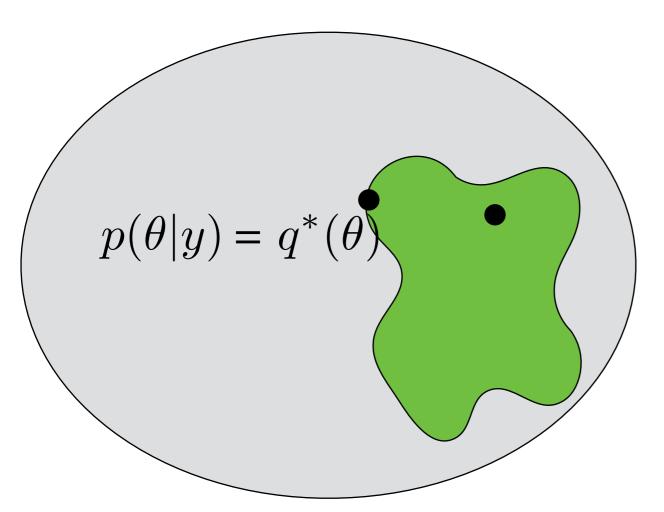
- Often optimum has non-zero KL (MFVB, Gaussian VB)
- Even if optimum has zero KL, algorithm might not reach zero



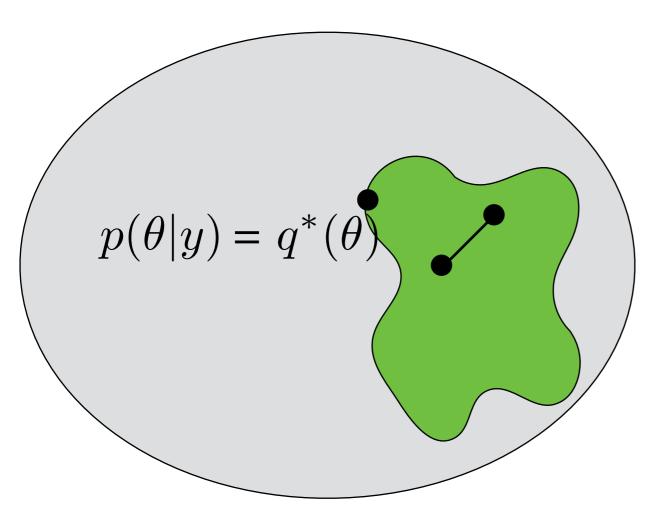
- Often optimum has non-zero KL (MFVB, Gaussian VB)
- Even if optimum has zero KL, algorithm might not reach zero



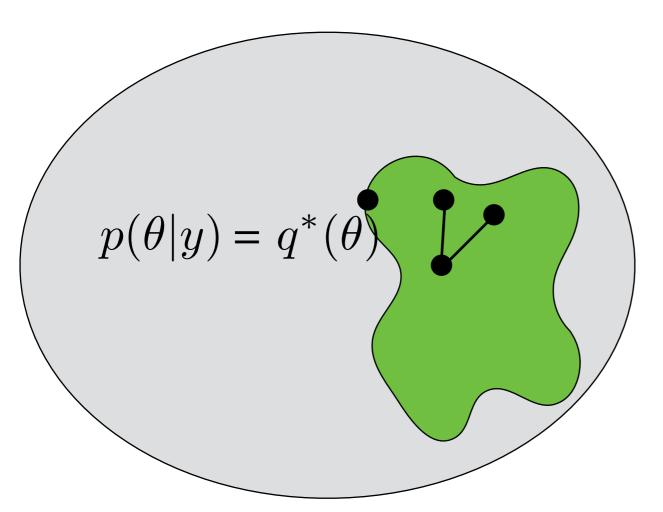
- Often optimum has non-zero KL (MFVB, Gaussian VB)
- Even if optimum has zero KL, algorithm might not reach zero



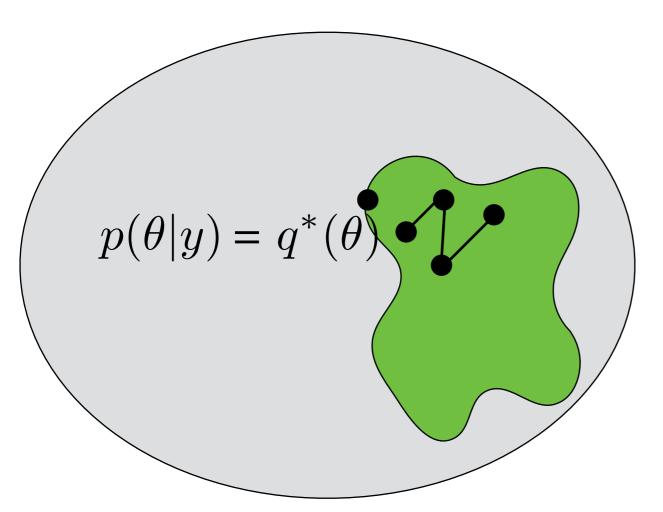
- Often optimum has non-zero KL (MFVB, Gaussian VB)
- Even if optimum has zero KL, algorithm might not reach zero



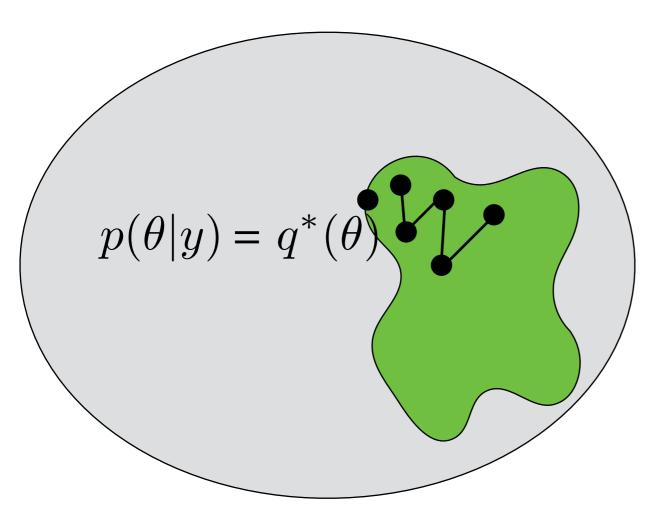
- Often optimum has non-zero KL (MFVB, Gaussian VB)
- Even if optimum has zero KL, algorithm might not reach zero



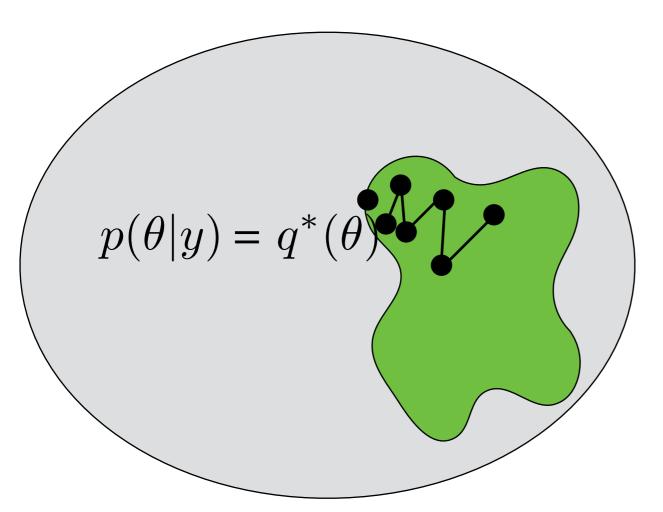
- Often optimum has non-zero KL (MFVB, Gaussian VB)
- Even if optimum has zero KL, algorithm might not reach zero



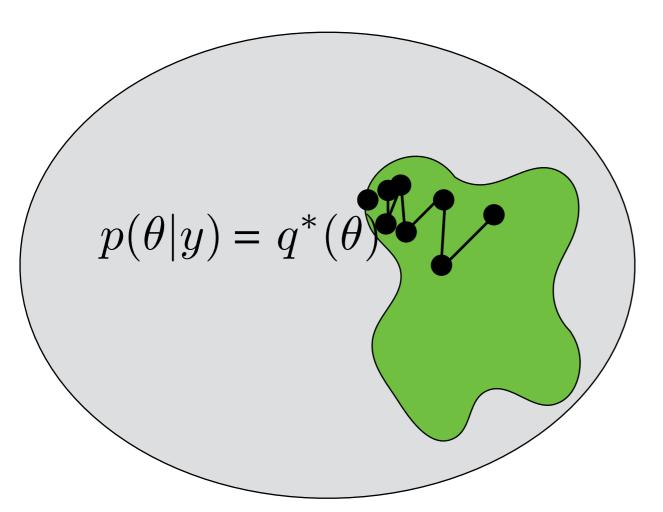
- Often optimum has non-zero KL (MFVB, Gaussian VB)
- Even if optimum has zero KL, algorithm might not reach zero



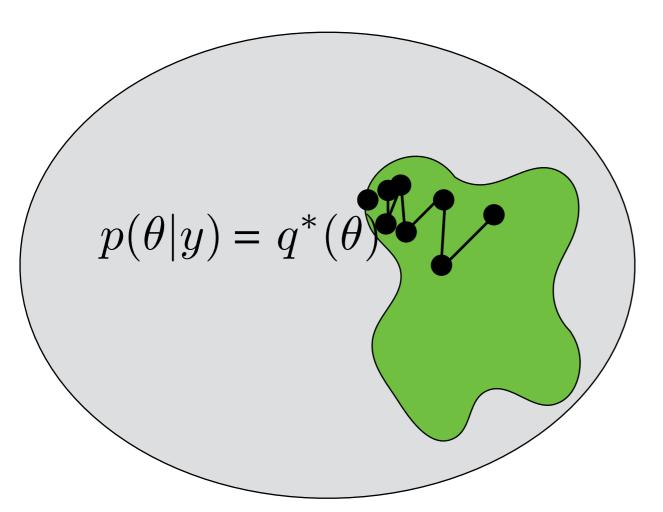
- Often optimum has non-zero KL (MFVB, Gaussian VB)
- Even if optimum has zero KL, algorithm might not reach zero



- Often optimum has non-zero KL (MFVB, Gaussian VB)
- Even if optimum has zero KL, algorithm might not reach zero

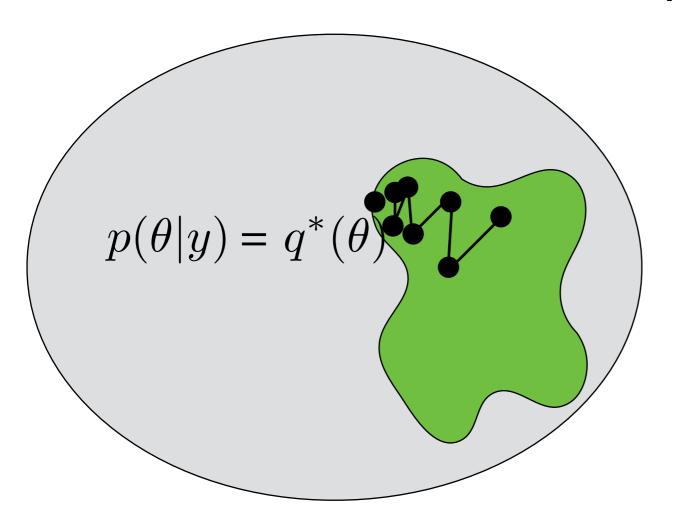


- Often optimum has non-zero KL (MFVB, Gaussian VB)
- Even if optimum has zero KL, algorithm might not reach zero



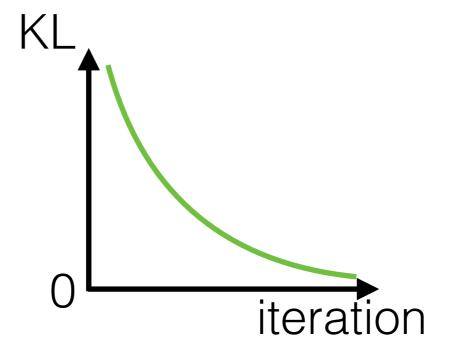
- Often optimum has non-zero KL (MFVB, Gaussian VB)
- Even if optimum has zero KL, algorithm might not reach zero

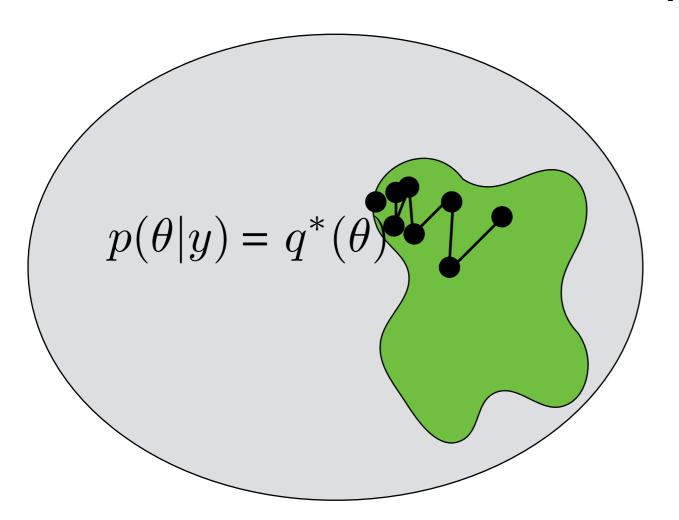
 Recall: the KL value isn't free



- Often optimum has non-zero KL (MFVB, Gaussian VB)
- Even if optimum has zero KL, algorithm might not reach zero

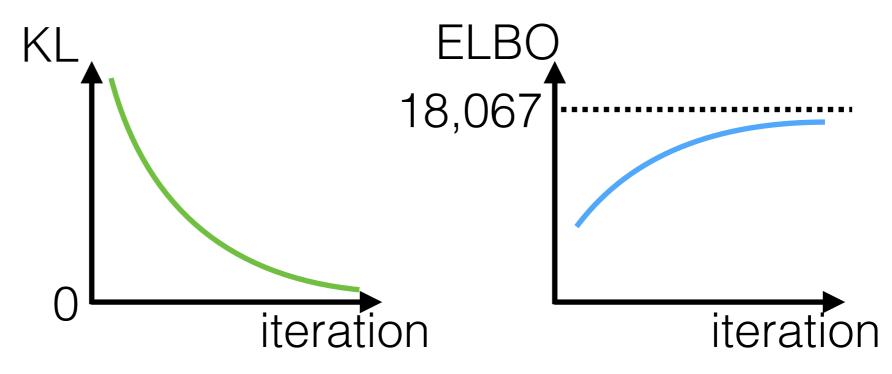
 Recall: the KL value isn't free





- Often optimum has non-zero KL (MFVB, Gaussian VB)
- Even if optimum has zero KL, algorithm might not reach zero

 Recall: the KL value isn't free



### Approximate Bayesian inference

Use  $q^*$  to approximate  $p(\cdot|y)$ 

Variational Bayes  $q^* = \mathrm{argmin}_{q \in Q} KL(q(\cdot)||p(\cdot|y))$ 

# How deep is the issue?

#### Mean-field variational Bayes

$$q^* = \operatorname{argmin}_{q \in Q_{\text{MFVB}}} KL(q(\cdot)||p(\cdot|y))$$

Algorithm

Implementation

Gaussian example was exact

#### Approximate Bayesian inference

Use  $q^*$  to approximate  $p(\cdot|y)$ 

#### **Variational Bayes**

$$q^* = \operatorname{argmin}_{q \in Q} KL(q(\cdot)||p(\cdot|y))$$

# How deep is the issue?

#### Mean-field variational Bayes

$$q^* = \operatorname{argmin}_{q \in Q_{\text{MFVB}}} KL(q(\cdot)||p(\cdot|y))$$

Algorithm

Implementation

Gaussian example was exact

- Bayes & Approximate Bayes review
- What is:
  - Variational Bayes (VB)
  - Mean-field variational Bayes (MFVB)
- Why use MFVB?
- When can we trust MFVB?
- Where do we go from here?

- Bayes & Approximate Bayes review
- What is:
  - Variational Bayes (VB)
  - Mean-field variational Bayes (MFVB)
- Why use MFVB?
- When can we trust MFVB?
- Where do we go from here?

- Bayes & Approximate Bayes review
- What is:
  - Variational Bayes (VB)
  - Mean-field variational Bayes (MFVB)
- Why use MFVB?
- When can we trust VB?
- Where do we go from here?

- Bayes & Approximate Bayes review
- What is:
  - Variational Bayes (VB)
  - Mean-field variational Bayes (MFVB)
- Why use MFVB?
- When can we trust VB?
- Where do we go from here?

- Bayes & Approximate Bayes review
- What is:
  - Variational Bayes (VB)
  - Mean-field variational Bayes (MFVB)
- Why use MFVB?
- When can we trust VB?
- Where do we go from here?

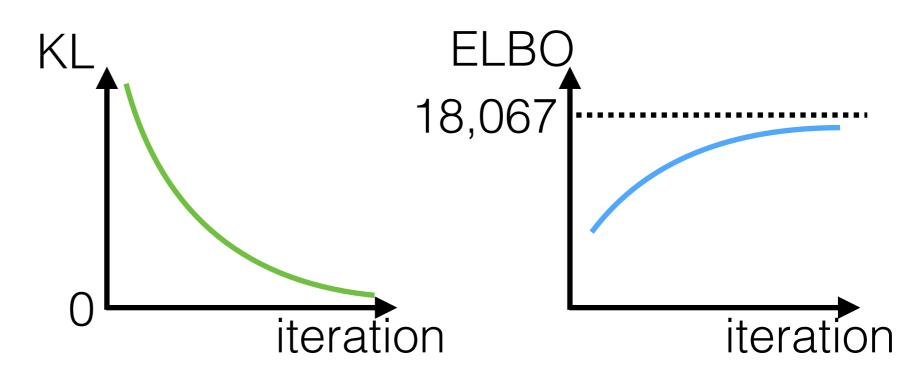
# Latent Dirichlet Allocation (LDA)

"Arts"	"Budgets"	"Children"	"Education"
NEW	MILLION	CHILDREN	SCHOOL
FILM	TAX	WOMEN	STUDENTS
SHOW	PROGRAM	PEOPLE	SCHOOLS
MUSIC	BUDGET	CHILD	EDUCATION
MOVIE	BILLION	YEARS	TEACHERS
PLAY	FEDERAL	FAMILIES	HIGH
MUSICAL	YEAR	WORK	PUBLIC
BEST	SPENDING	PARENTS	TEACHER
ACTOR	NEW	SAYS	BENNETT
FIRST	STATE	FAMILY	MANIGAT
YORK	PLAN	WELFARE	NAMPHY
OPERA	MONEY	MEN	STATE
THEATER	PROGRAMS	PERCENT	PRESIDENT
ACTRESS	GOVERNMENT	CARE	ELEMENTARY
LOVE	CONGRESS	LIFE	HAITI

The William Randolph Hearst Foundation will give \$1.25 million to Lincoln Center, Metropolitan Opera Co., New York Philharmonic and Juilliard School. "Our board felt that we had a real opportunity to make a mark on the future of the performing arts with these grants an act every bit as important as our traditional areas of support in health, medical research, education and the social services," Hearst Foundation President Randolph A. Hearst said Monday in announcing the grants. Lincoln Center's share will be \$200,000 for its new building, which will house young artists and provide new public facilities. The Metropolitan Opera Co. and New York Philharmonic will receive \$400,000 each. The Juilliard School, where music and the performing arts are taught, will get \$250,000. The Hearst Foundation, a leading supporter of the Lincoln Center Consolidated Corporate Fund, will make its usual annual \$100,000 donation, too.

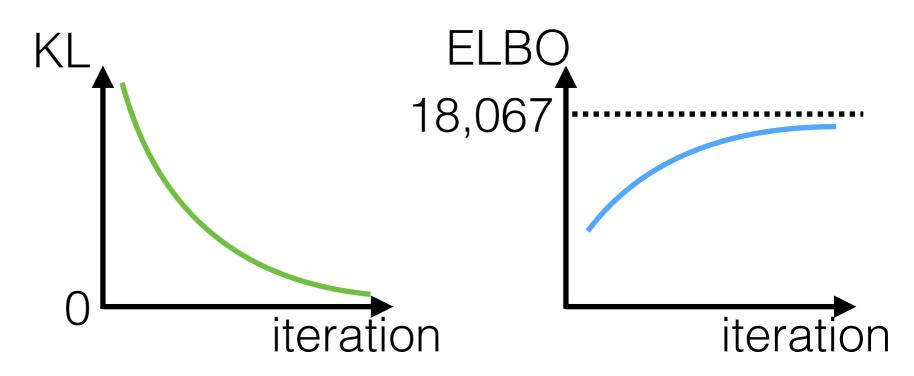
Reliable diagnostics

- Reliable diagnostics
  - cf. KL, ELBO



- Reliable diagnostics
  - cf. KL, ELBO

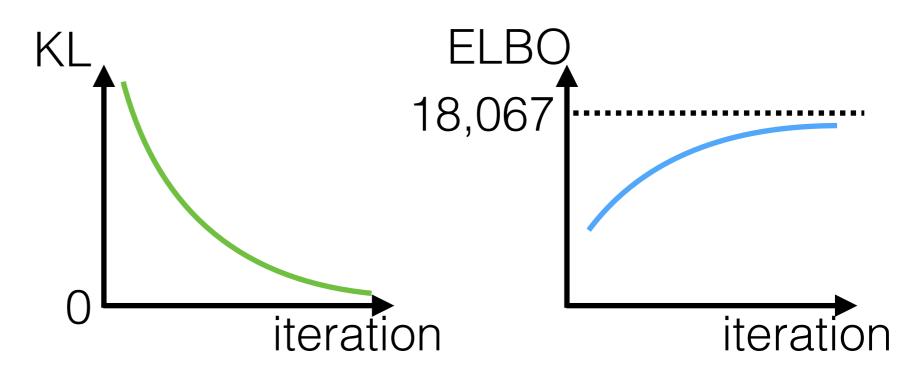
[Gorham, Mackey 2015, 2017; Chwialkowski, Strathmann, Gretton 2016; Jitkrittum et al 2017; Talts et al 2018; Yao et al 2018, etc.]



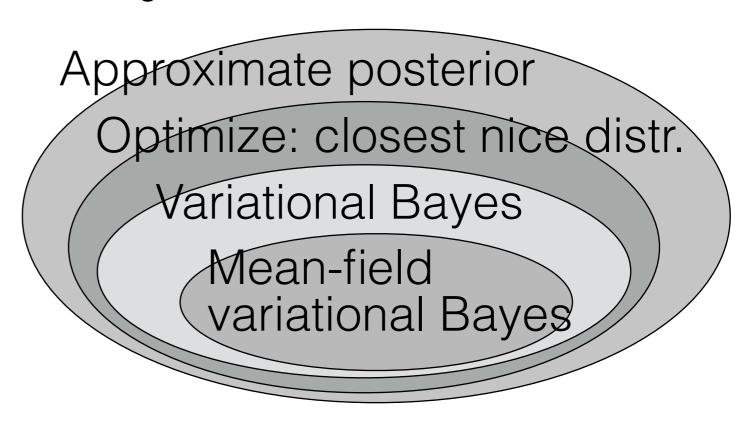
"Yes, but did it work? Evaluating variational inference" ICML 2018

- Reliable diagnostics
  - cf. KL, ELBO

[Gorham, Mackey 2015, 2017; Chwialkowski, Strathmann, Gretton 2016; Jitkrittum et al 2017; Talts et al 2018; Yao et al 2018, etc.]

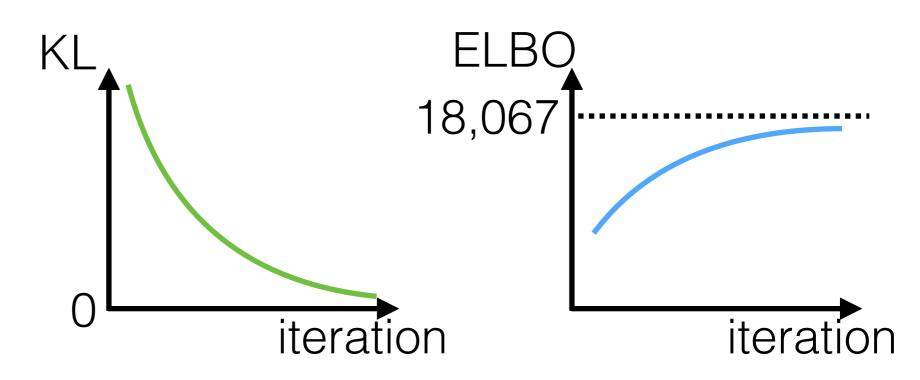


"Yes, but did it work? Evaluating variational inference" ICML 2018



- Reliable diagnostics
  - cf. KL, ELBO

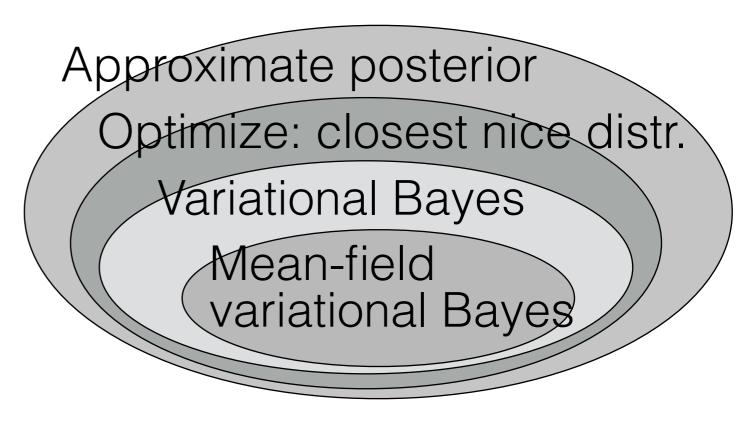
[Gorham, Mackey 2015, 2017; Chwialkowski, Strathmann, Gretton 2016; Jitkrittum et al 2017; Talts et al 2018; Yao et al 2018, etc.]



"Yes, but did it work? Evaluating variational inference" ICML 2018

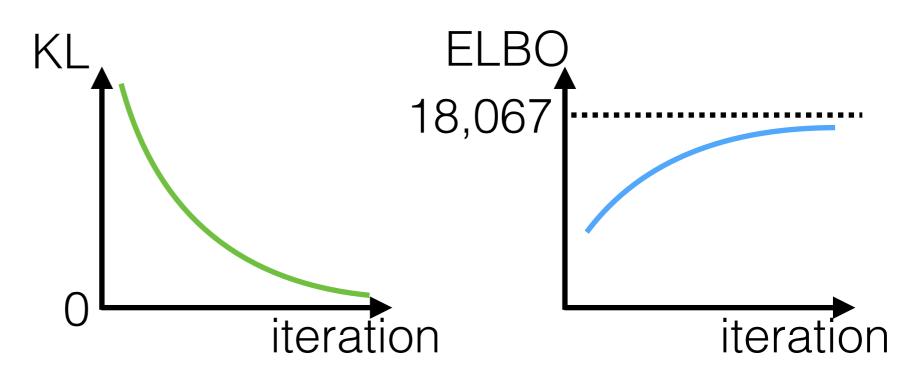
Alternative divergences:
 Time & accuracy

[Huggins, Kasprzak, Campbell, Broderick, 2018]



- Reliable diagnostics
  - cf. KL, ELBO

[Gorham, Mackey 2015, 2017; Chwialkowski, Strathmann, Gretton 2016; Jitkrittum et al 2017; Talts et al 2018; Yao et al 2018, etc.]



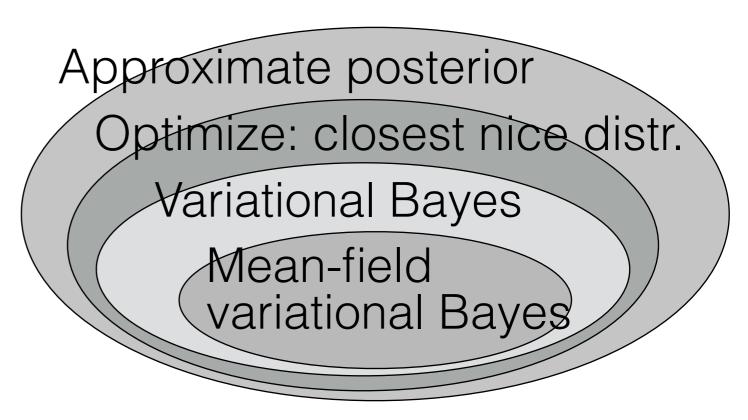
"Yes, but did it work? Evaluating variational inference" ICML 2018

Alternative divergences:
 Time & accuracy

[Huggins, Kasprzak, Campbell, Broderick, 2018]

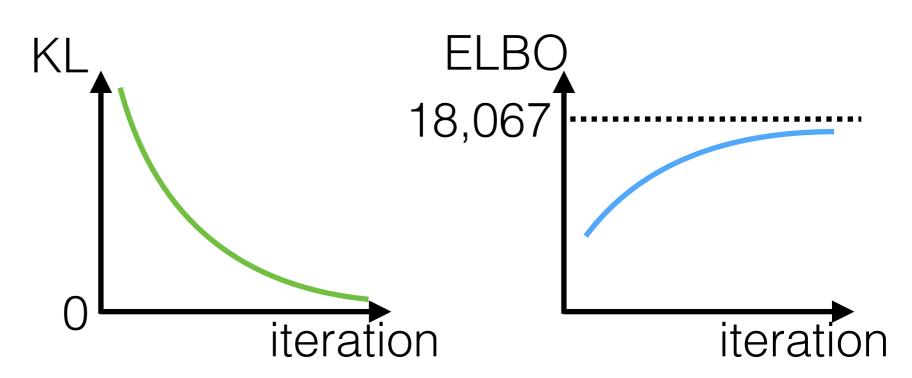
Corrections

[Giordano, Broderick, Jordan 2018]



- Reliable diagnostics
  - cf. KL, ELBO

[Gorham, Mackey 2015, 2017; Chwialkowski, Strathmann, Gretton 2016; Jitkrittum et al 2017; Talts et al 2018; Yao et al 2018, etc.]



"Yes, but did it work? Evaluating variational inference" ICML 2018

Alternative divergences:
 Time & accuracy

[Huggins, Kasprzak, Campbell, Broderick, 2018]

- Corrections [Giordano, Broderick, Jordan 2018]
- Theoretical guarantees on finite-data quality [Huggins, Car Broderick 201

[Huggins, Campbell, Broderick 2016; Campbell, Broderick 2018, 2019]

Approximate posterior

Optimize: closest nice distr.

Variational Bayes

Mean-field

variational Bayes

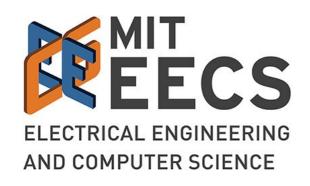
#### What to read next

#### Textbooks and Reviews

- Bishop. Pattern Recognition and Machine Learning, Ch 10. 2006.
- Blei, Kucukelbir, McAuliffe. Variational inference: A review for statisticians, *JASA* 2016.
- MacKay. Information Theory, Inference, and Learning Algorithms, Ch 33. 2003.
- Murphy. Machine Learning: A Probabilistic Perspective, Ch 21. 2012.
- Ormerod, Wand. Explaining Variational Approximations. Amer Stat 2010.
- Turner, Sahani. Two problems with variational expectation maximisation for time-series models. In Bayesian Time Series Models, 2011.
- Wainwright, Jordan. Graphical models, exponential families, and variational inference.
   Foundations and Trends in Machine Learning, 2008.

#### Our Experiments

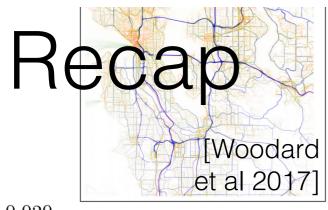
- RJ Giordano, T Broderick, and MI Jordan. Linear response methods for accurate covariance estimates from mean field variational Bayes. NeurIPS 2015.
- RJ Giordano, T Broderick, R Meager, J Huggins, and MI Jordan. Fast robustness
  quantification with variational Bayes. ICML Data4Good Workshop 2016.
- RJ Giordano, T Broderick, and MI Jordan. Covariances, robustness, and variational Bayes. Journal of Machine Learning Research, 2018.
- J Huggins, M Kasprzak, T Campbell, T Broderick. Practical bounds on the error of Bayesian posterior approximations: A nonasymptotic approach, 2018. ArXiv: 1809.09505.



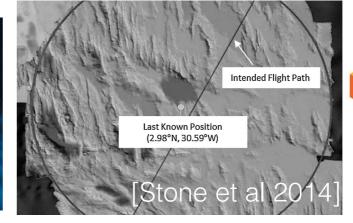




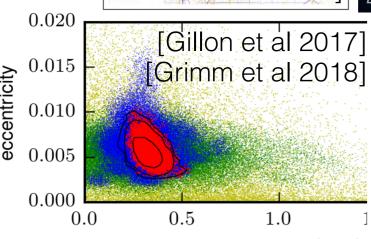
# Automated, Scalable Bayesian Inference via Data Summarization







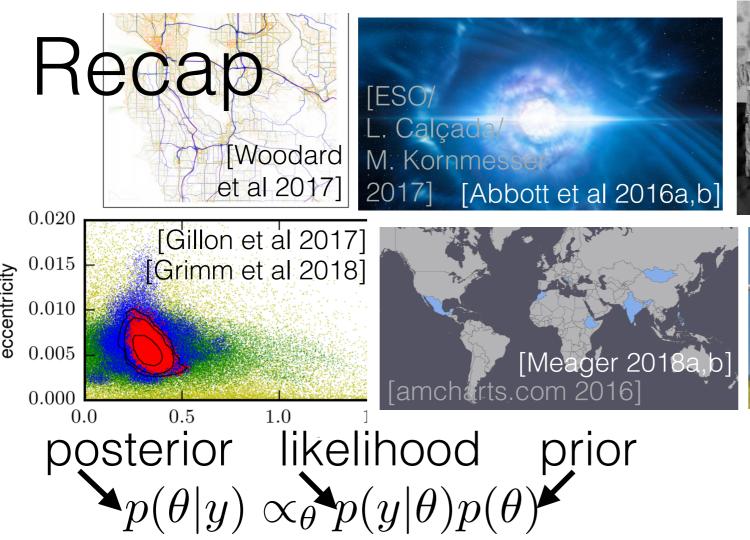








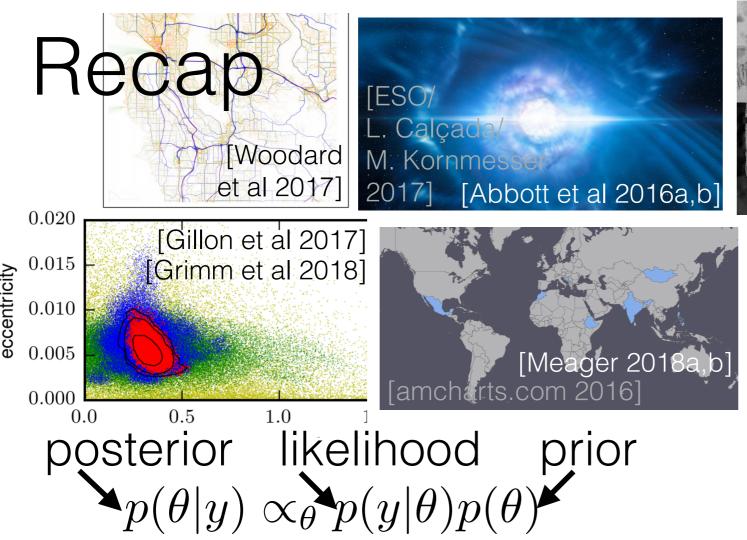








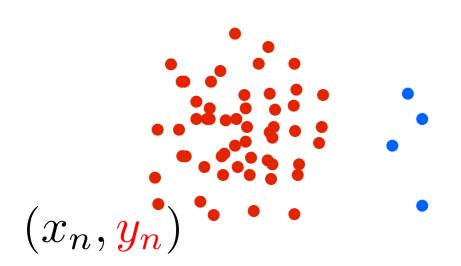


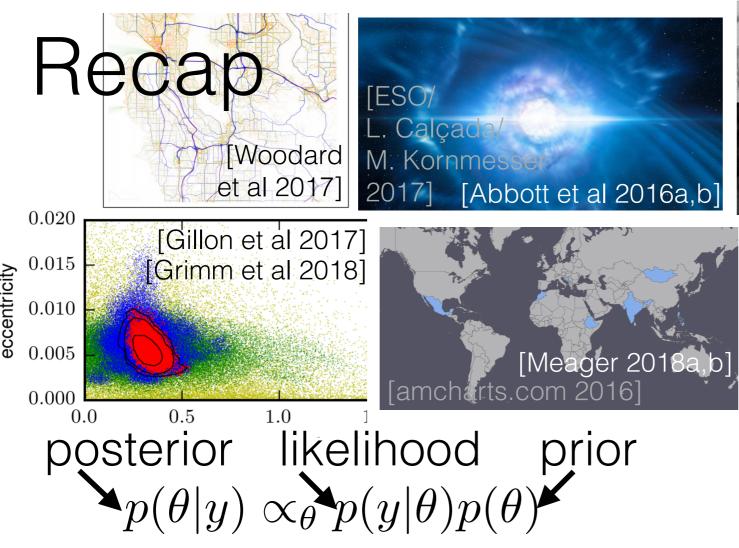








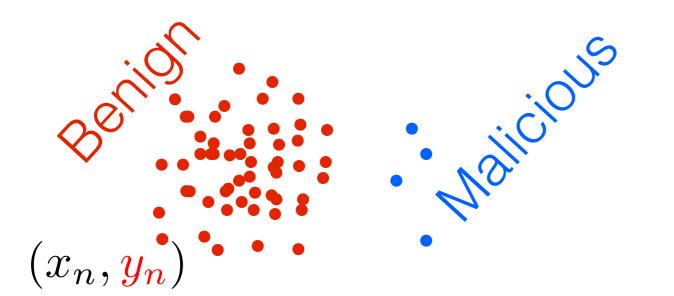


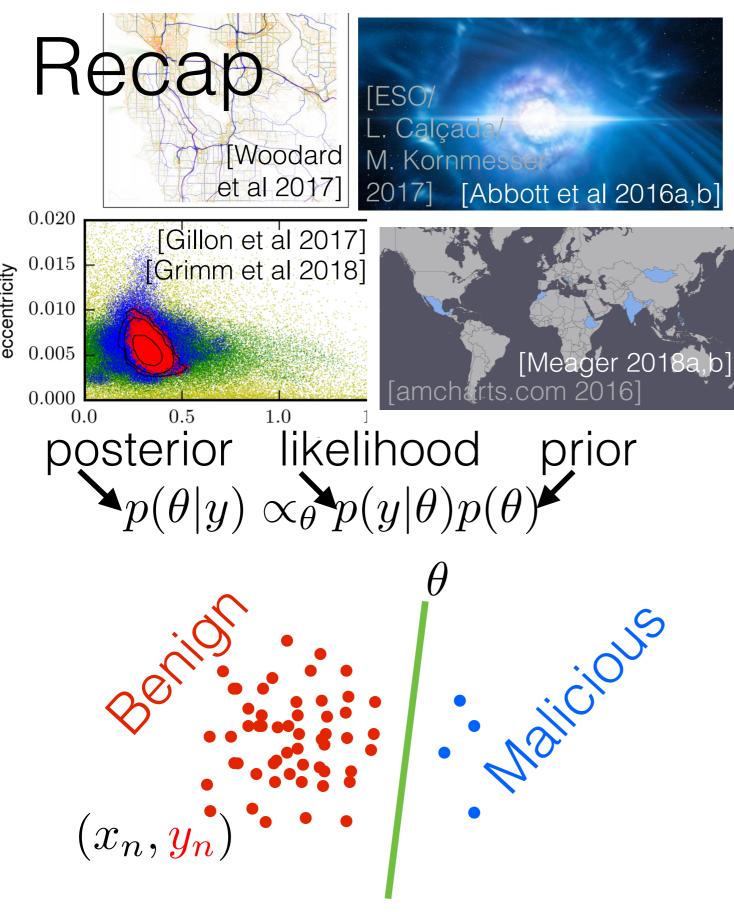


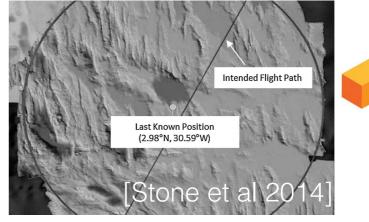








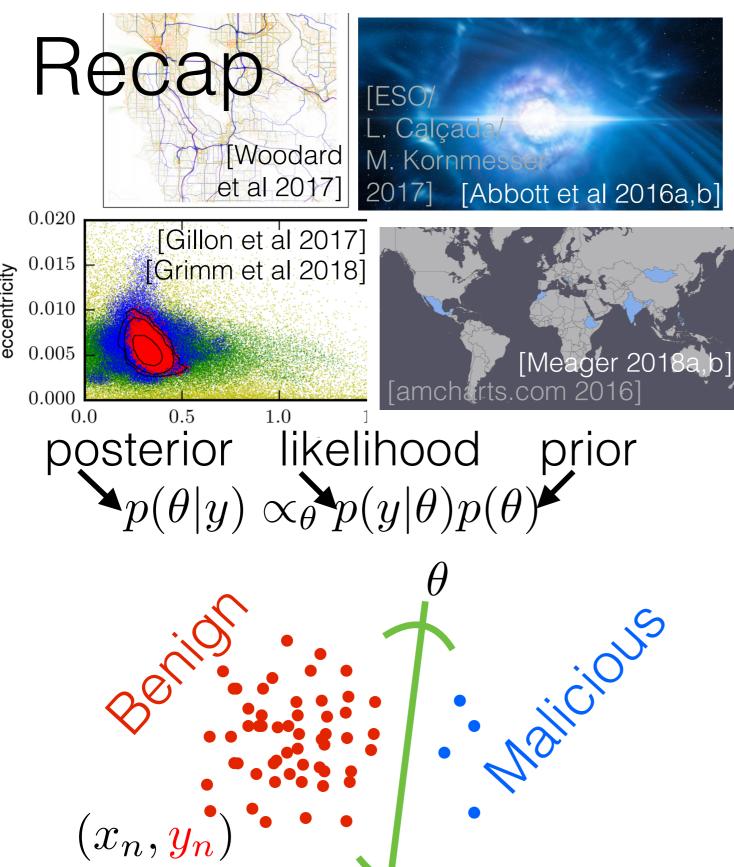


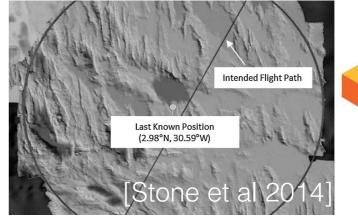








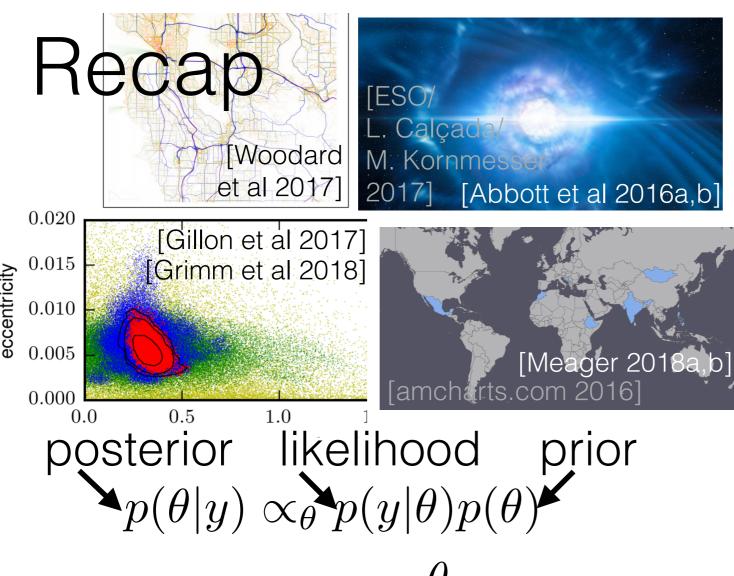


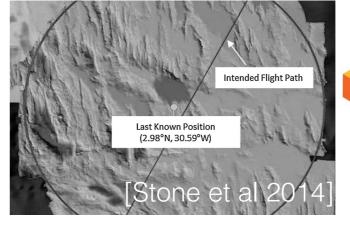








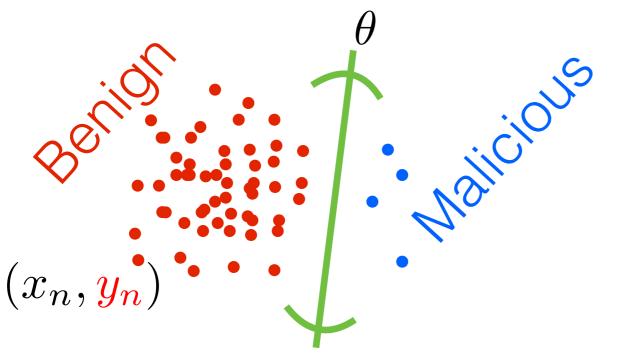


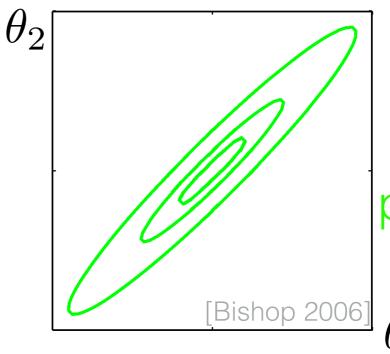






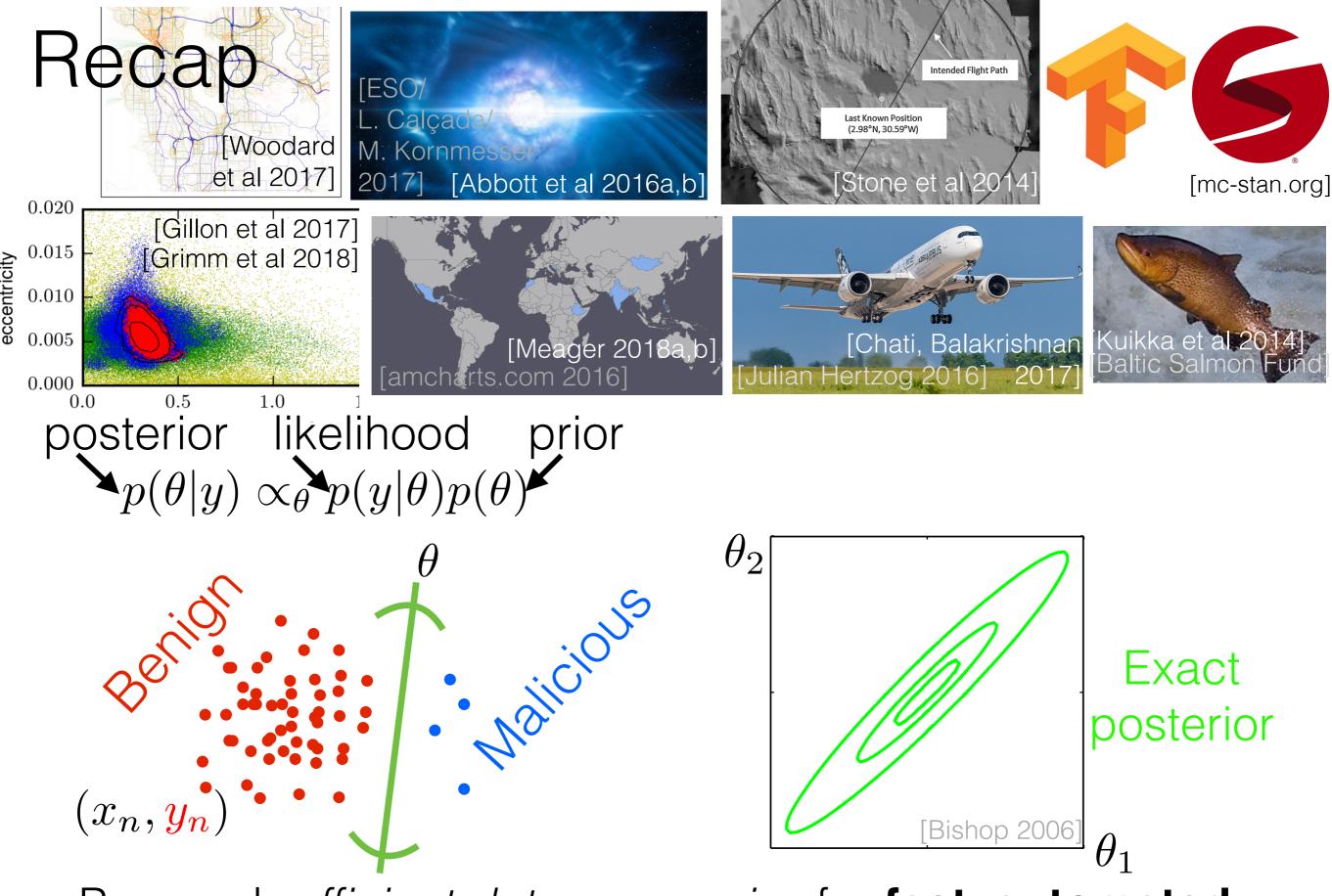






Exact posterior

 $\theta_1$ 



Proposal: efficient data summaries for fast, automated,
 approximations with error bounds for finite data

• The "core" of the data set

- The "core" of the data set
- Uniform data subsampling isn't enough

- The "core" of the data set
- Uniform data subsampling isn't enough
- Importance sampling for "coresets"

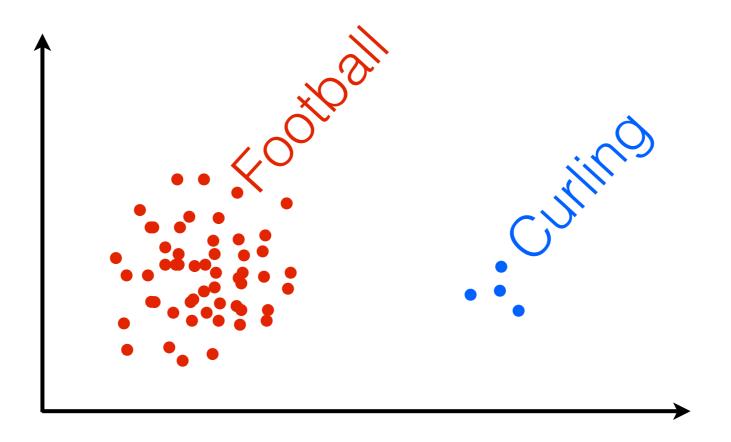
- The "core" of the data set
- Uniform data subsampling isn't enough
- Importance sampling for "coresets"
- Optimization for "coresets"

- The "core" of the data set
- Uniform data subsampling isn't enough
- Importance sampling for "coresets"
- Optimization for "coresets"
- Approximate sufficient statistics

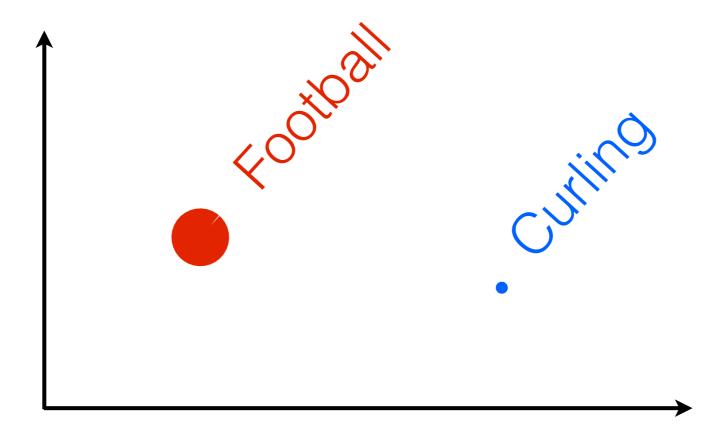
- The "core" of the data set
- Uniform data subsampling isn't enough
- Importance sampling for "coresets"
- Optimization for "coresets"
- Approximate sufficient statistics

• Observe: redundancies can exist even if data isn't "tall"

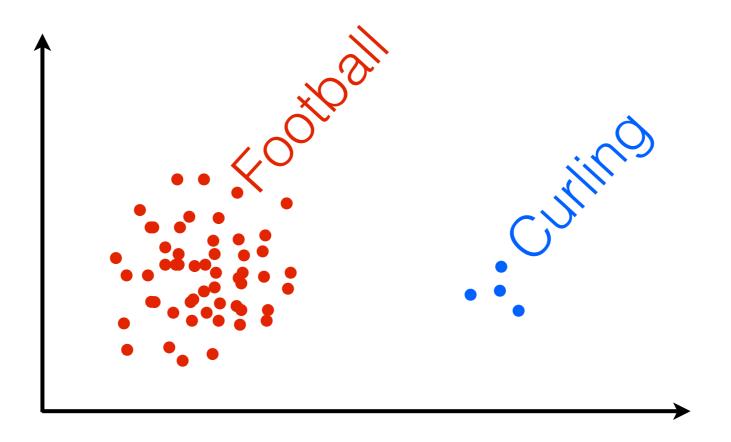
Observe: redundancies can exist even if data isn't "tall"



Observe: redundancies can exist even if data isn't "tall"

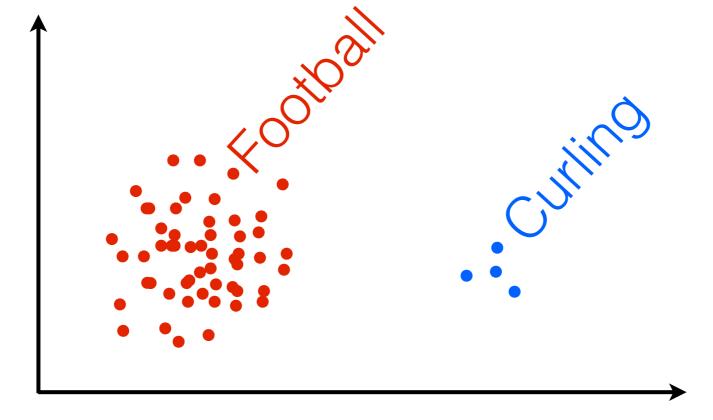


Observe: redundancies can exist even if data isn't "tall"



Observe: redundancies can exist even if data isn't "tall"

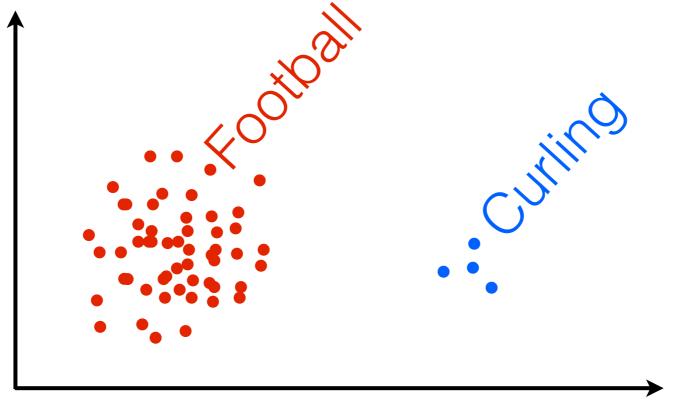
Coresets: pre-process data to get a smaller, weighted



Observe: redundancies can exist even if data isn't "tall"

Coresets: pre-process data to get a smaller, weighted

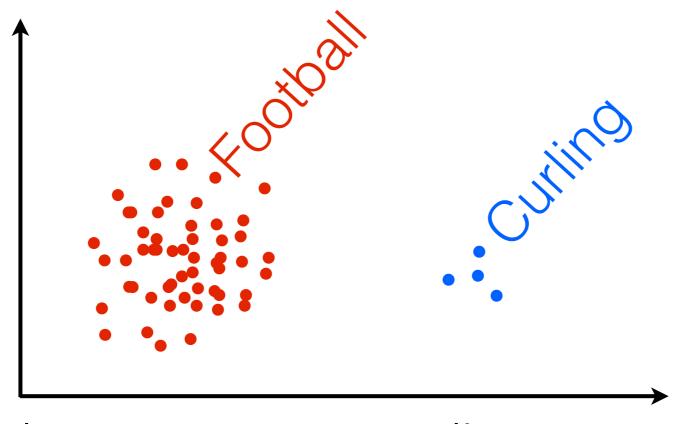
data set



Theoretical guarantees on quality

Observe: redundancies can exist even if data isn't "tall"

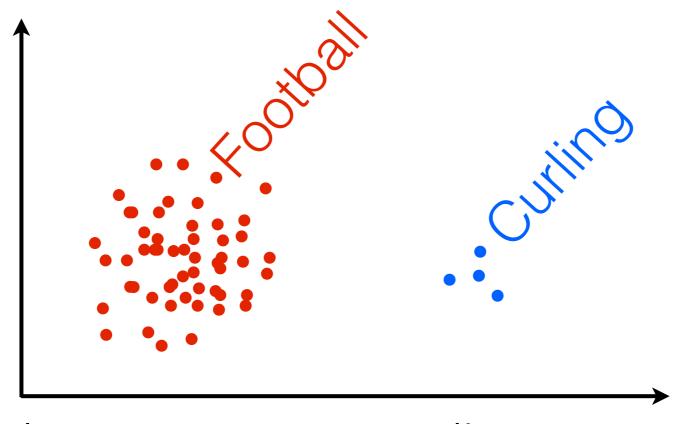
Coresets: pre-process data to get a smaller, weighted



- Theoretical guarantees on quality
- How to develop coresets for Bayes?

Observe: redundancies can exist even if data isn't "tall"

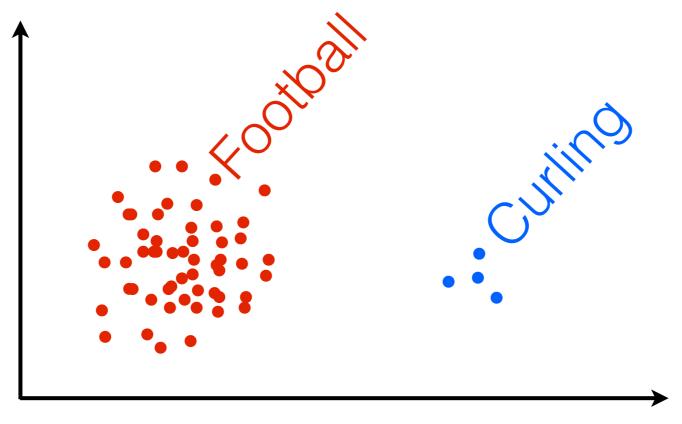
Coresets: pre-process data to get a smaller, weighted



- Theoretical guarantees on quality
- How to develop coresets for diverse tasks/geometries?

Observe: redundancies can exist even if data isn't "tall"

Coresets: pre-process data to get a smaller, weighted

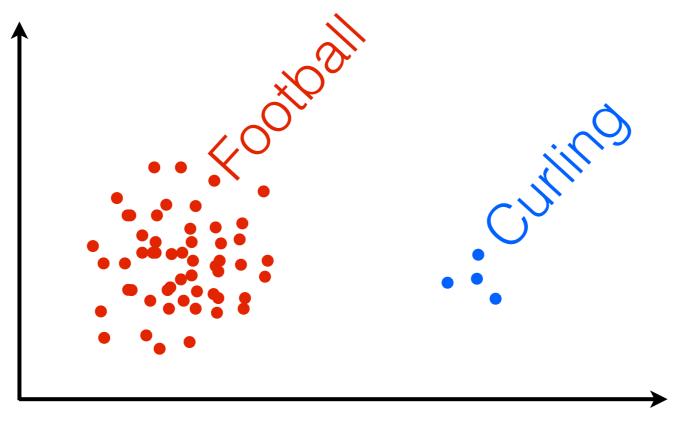


- Theoretical guarantees on quality
- How to develop coresets for diverse tasks/geometries?
- Previous heuristics: data squashing, big data GPs

Observe: redundancies can exist even if data isn't "tall"

Coresets: pre-process data to get a smaller, weighted

data set

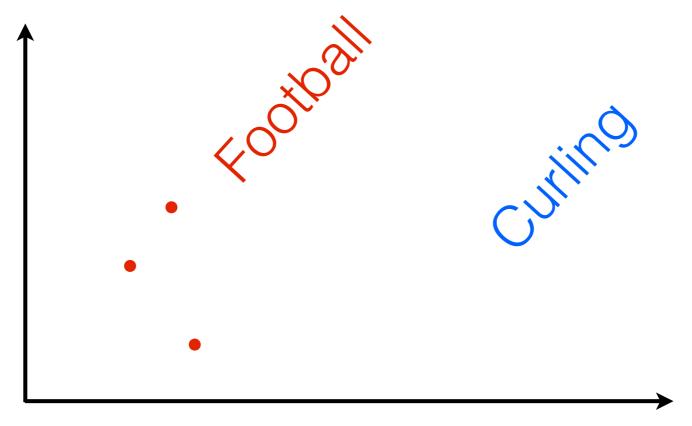


- Theoretical guarantees on quality
- How to develop coresets for diverse tasks/geometries?
- Previous heuristics: data squashing, big data GPs
- Compare to subsampling

Observe: redundancies can exist even if data isn't "tall"

Coresets: pre-process data to get a smaller, weighted

data set

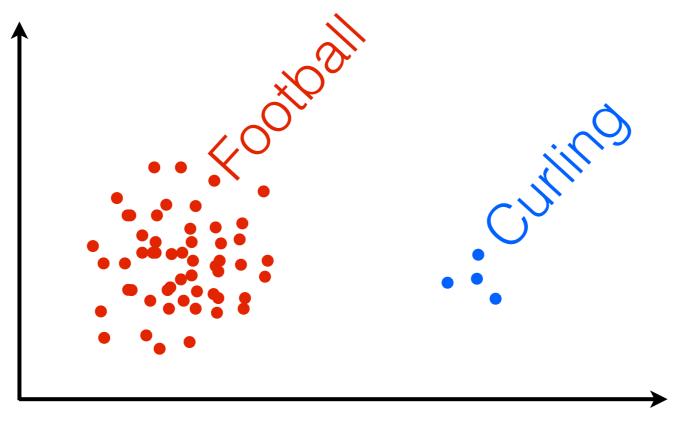


- Theoretical guarantees on quality
- How to develop coresets for diverse tasks/geometries?
- Previous heuristics: data squashing, big data GPs
- Compare to subsampling

Observe: redundancies can exist even if data isn't "tall"

Coresets: pre-process data to get a smaller, weighted

data set

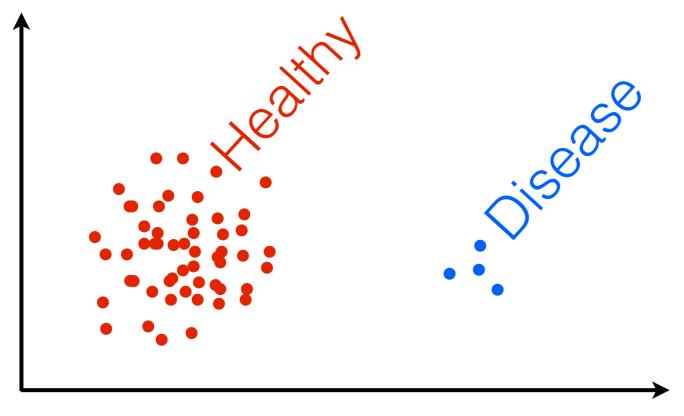


- Theoretical guarantees on quality
- How to develop coresets for diverse tasks/geometries?
- Previous heuristics: data squashing, big data GPs
- Compare to subsampling

Observe: redundancies can exist even if data isn't "tall"

Coresets: pre-process data to get a smaller, weighted

data set

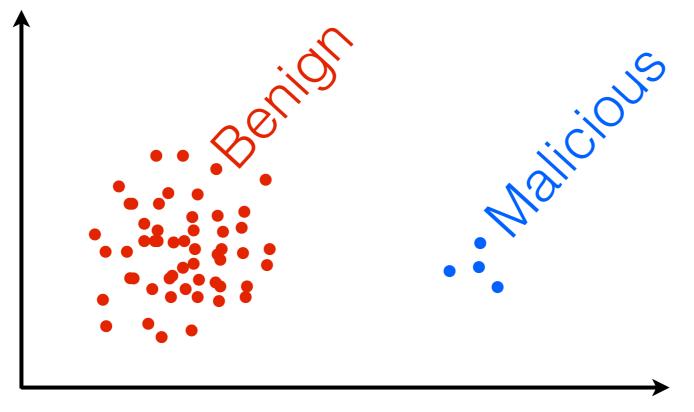


- Theoretical guarantees on quality
- How to develop coresets for diverse tasks/geometries?
- Previous heuristics: data squashing, big data GPs
- Compare to subsampling

Observe: redundancies can exist even if data isn't "tall"

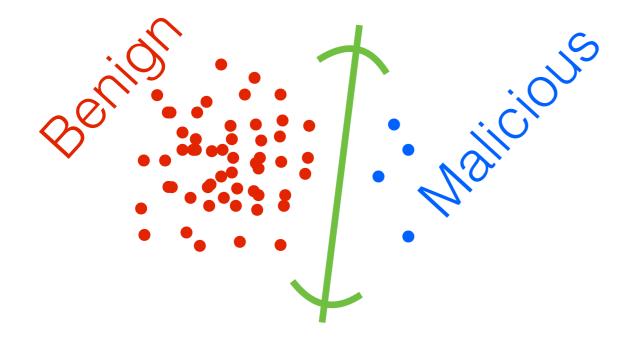
Coresets: pre-process data to get a smaller, weighted

data set

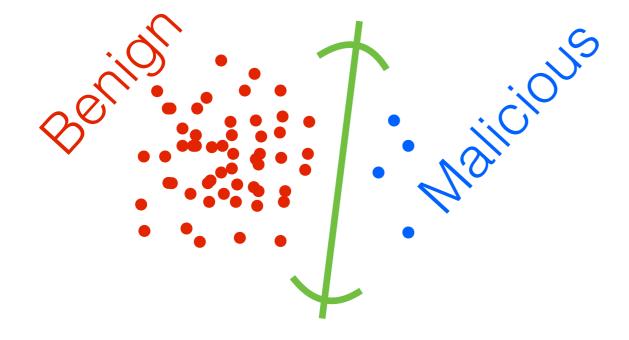


- Theoretical guarantees on quality
- How to develop coresets for diverse tasks/geometries?
- Previous heuristics: data squashing, big data GPs
- Compare to subsampling

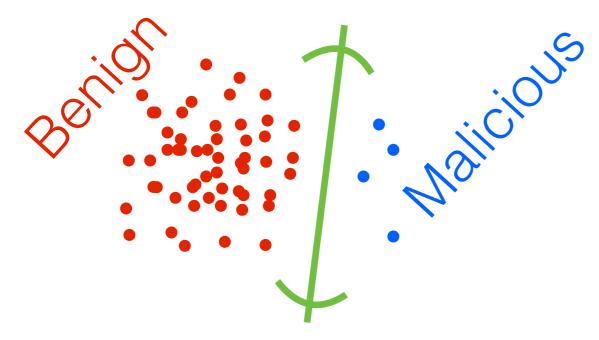
• Posterior  $p(\theta|y) \propto_{\theta} p(y|\theta)p(\theta)$ 



- Posterior  $p(\theta|y) \propto_{\theta} p(y|\theta)p(\theta)$
- Log likelihood  $\mathcal{L}_n(\theta) := \log p(y_n|\theta)$ ,  $\mathcal{L}(\theta) := \sum_{n=1}^{\infty} \mathcal{L}_n(\theta)$

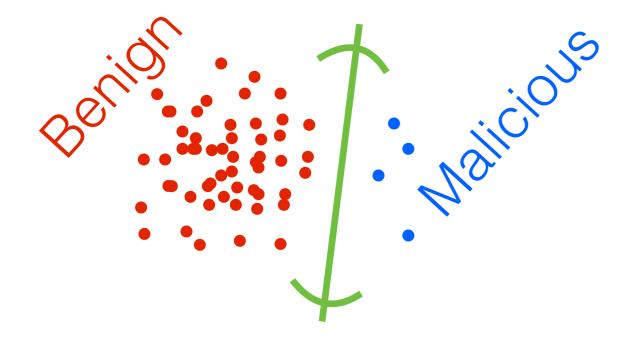


- Posterior  $p(\theta|y) \propto_{\theta} p(y|\theta)p(\theta)$
- Log likelihood  $\mathcal{L}_n(\theta) := \log p(y_n|\theta)$ ,  $\mathcal{L}(\theta) := \sum \mathcal{L}_n(\theta)$
- Coreset log likelihood

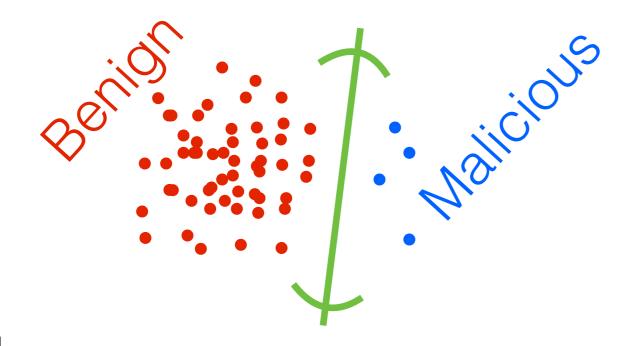


- Posterior  $p(\theta|y) \propto_{\theta} p(y|\theta)p(\theta)$
- Log likelihood  $\mathcal{L}_n(\theta) := \log p(y_n|\theta)$ ,  $\mathcal{L}(\theta) := \sum_{n=1}^{\infty} \mathcal{L}_n(\theta)$
- Coreset log likelihood

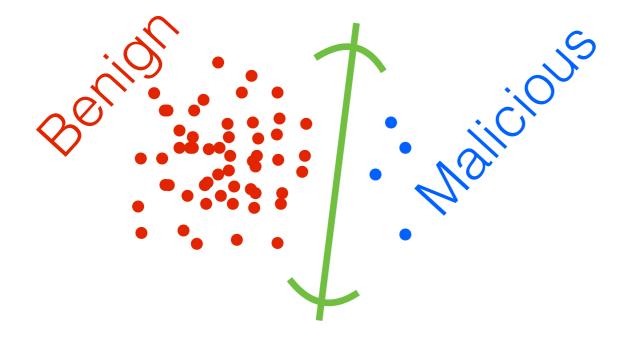
$$||w||_0 \ll N$$

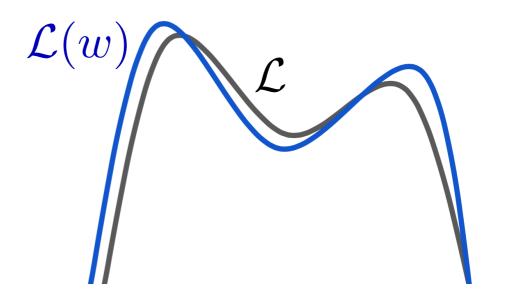


- Posterior  $p(\theta|y) \propto_{\theta} p(y|\theta)p(\theta)$
- Log likelihood  $\mathcal{L}_n(\theta) := \log p(y_n|\theta)$ ,  $\mathcal{L}(\theta) := \sum_{n=1}^N \mathcal{L}_n(\theta)$
- Coreset log likelihood  $\mathcal{L}(w,\theta) := \sum_{n=1}^N w_n \mathcal{L}_n(\theta)$  s.t.  $\|w\|_0 \ll N$

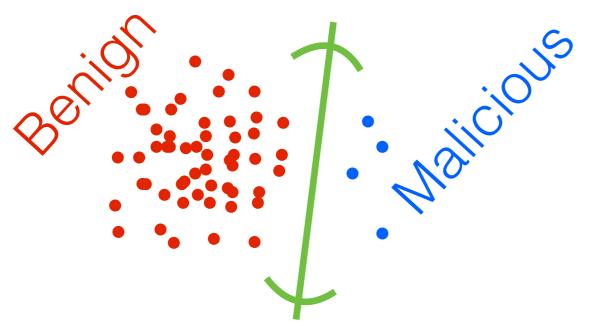


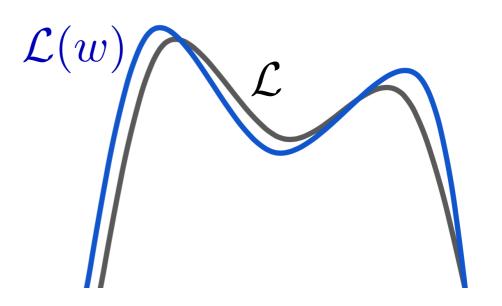
- Posterior  $p(\theta|y) \propto_{\theta} p(y|\theta)p(\theta)$
- Log likelihood  $\mathcal{L}_n(\theta) := \log p(y_n|\theta)$ ,  $\mathcal{L}(\theta) := \sum_{n=1}^N \mathcal{L}_n(\theta)$
- Coreset log likelihood  $\mathcal{L}(w,\theta) := \sum_{n=1}^N w_n \mathcal{L}_n(\theta)$  s.t.  $\|w\|_0 \ll N$



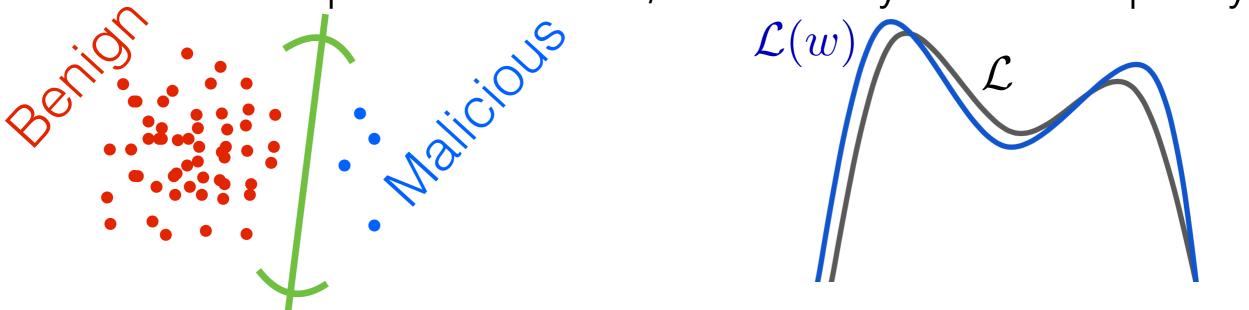


- Posterior  $p(\theta|y) \propto_{\theta} p(y|\theta)p(\theta)$
- Log likelihood  $\mathcal{L}_n(\theta) := \log p(y_n|\theta)$ ,  $\mathcal{L}(\theta) := \sum_{n=1}^N \mathcal{L}_n(\theta)$
- Coreset log likelihood  $\mathcal{L}(w,\theta) := \sum_{n=1}^{\infty} w_n \mathcal{L}_n(\theta)$  s.t.  $\|w\|_0 \ll N$
- $\varepsilon$ -coreset:  $\|\mathcal{L}(w) \mathcal{L}\| \leq \epsilon$





- Posterior  $p(\theta|y) \propto_{\theta} p(y|\theta)p(\theta)$
- Log likelihood  $\mathcal{L}_n(\theta) := \log p(y_n|\theta)$ ,  $\mathcal{L}(\theta) := \sum_{n=1}^N \mathcal{L}_n(\theta)$
- Coreset log likelihood  $\mathcal{L}(w,\theta) := \sum_{n=1}^{\infty} w_n \mathcal{L}_n(\theta)$  s.t.  $\|w\|_0 \ll N$
- $\varepsilon$ -coreset:  $\|\mathcal{L}(w) \mathcal{L}\| \leq \epsilon$ 
  - Bound on Wasserstein distance to exact posterior ->
    bound on posterior mean/uncertainty estimate quality

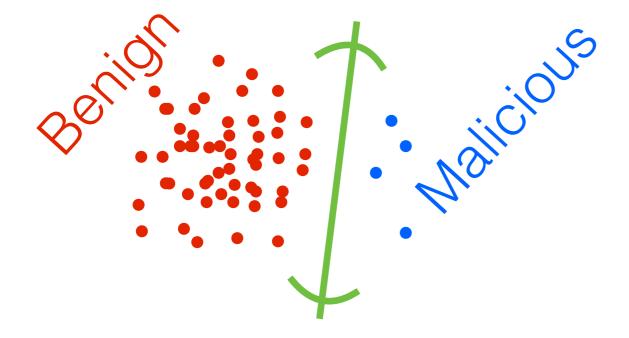


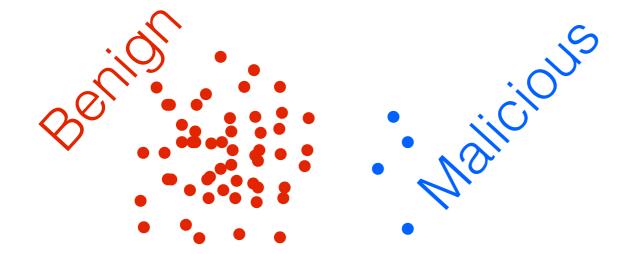
#### Roadmap

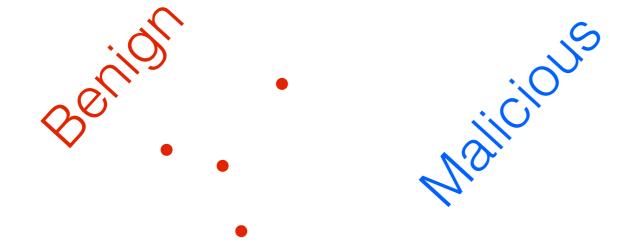
- The "core" of the data set
- Uniform data subsampling isn't enough
- Importance sampling for "coresets"
- Optimization for "coresets"
- Approximate sufficient statistics

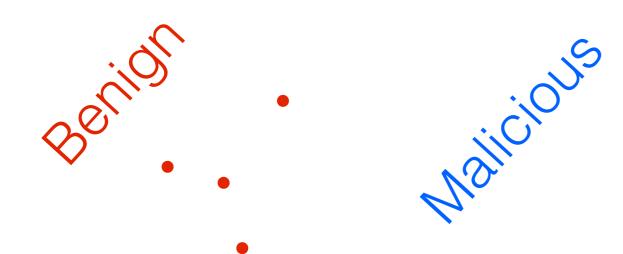
#### Roadmap

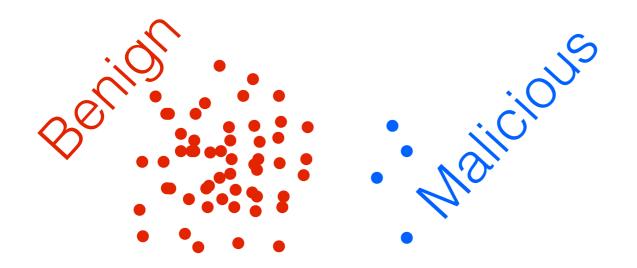
- The "core" of the data set
- Uniform data subsampling isn't enough
- Importance sampling for "coresets"
- Optimization for "coresets"
- Approximate sufficient statistics



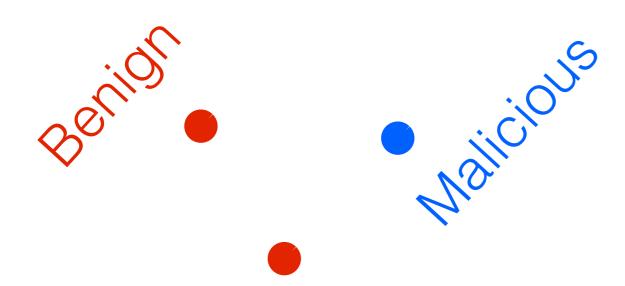


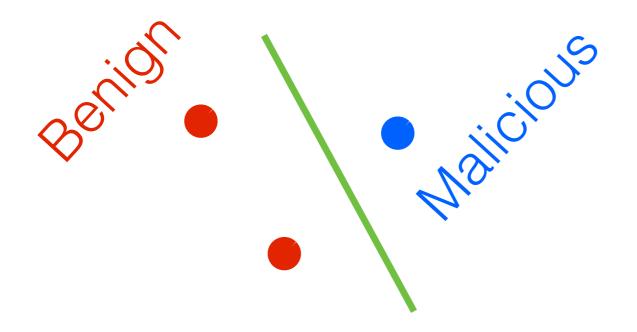


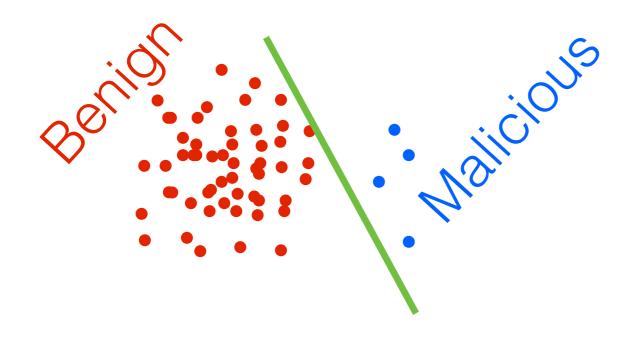


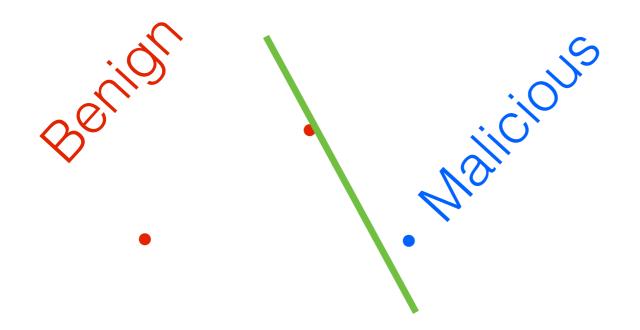


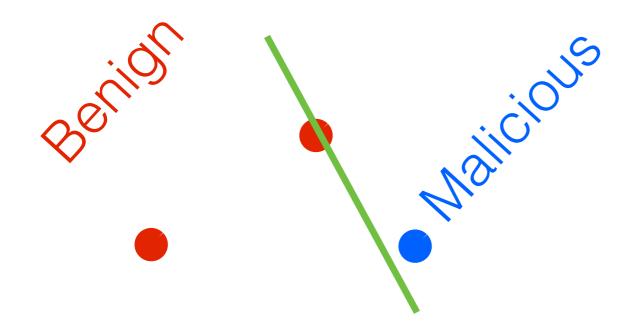


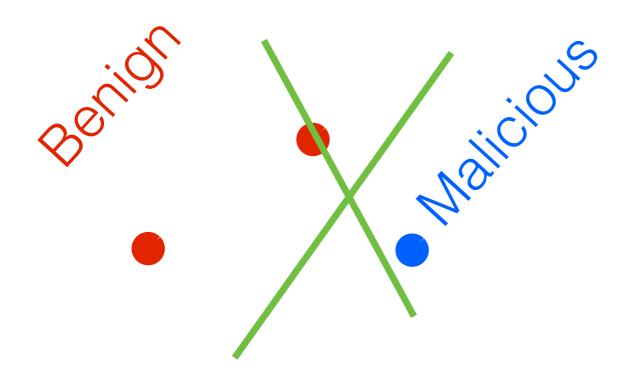


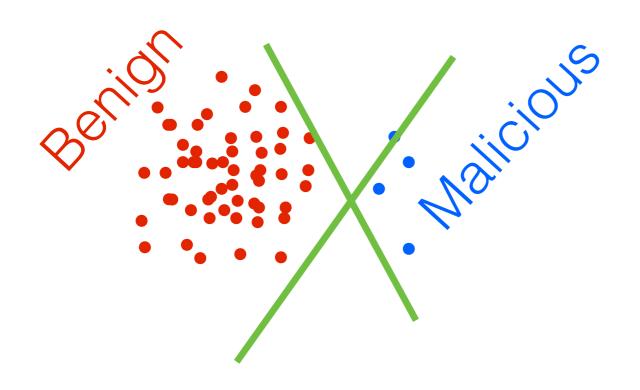


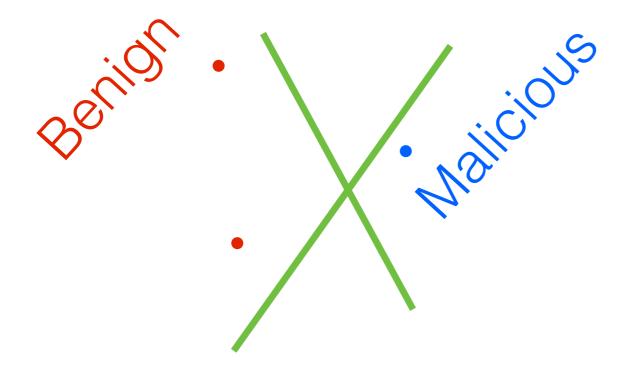


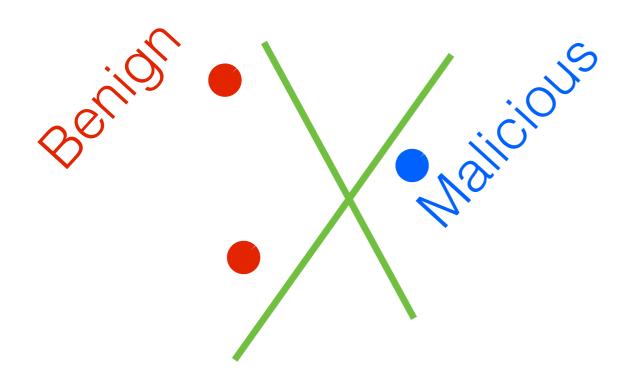


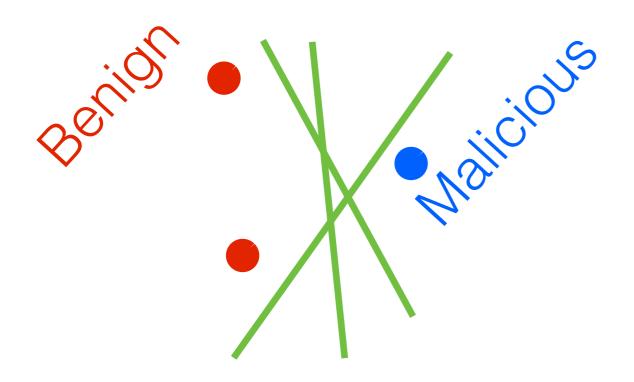


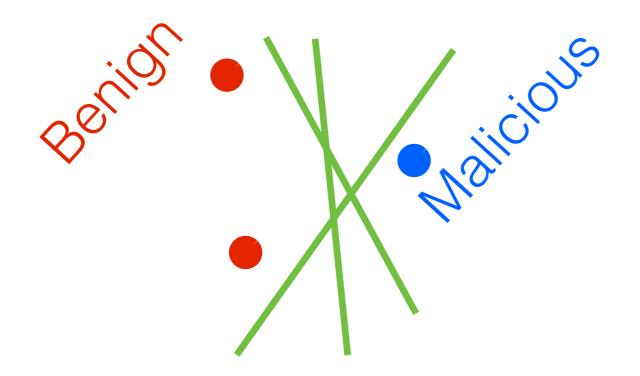




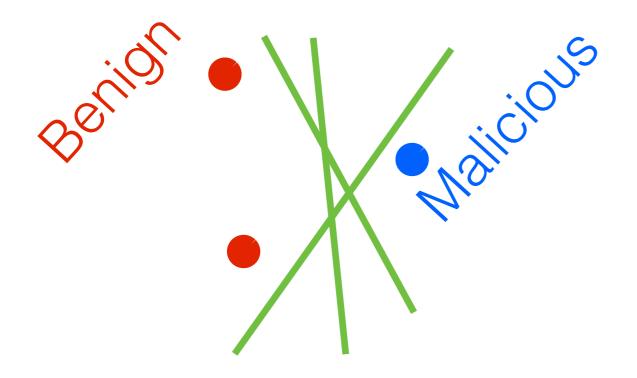




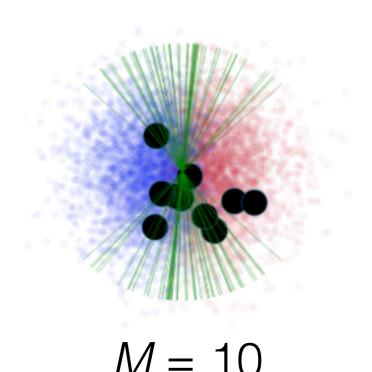


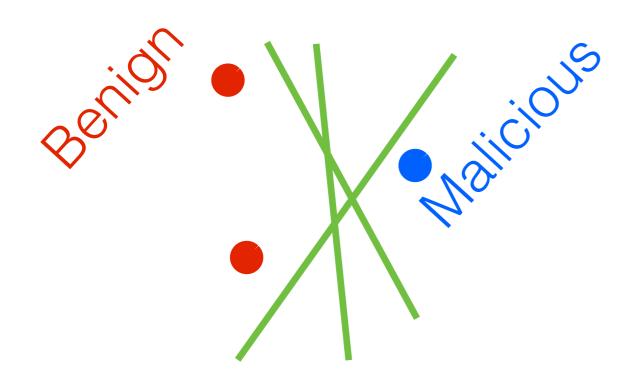


- Might miss important data
- Noisy estimates

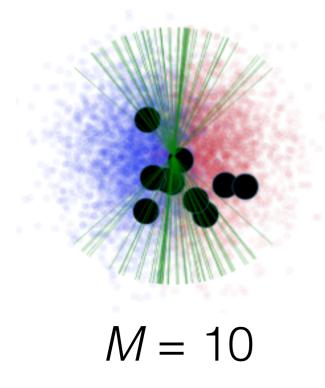


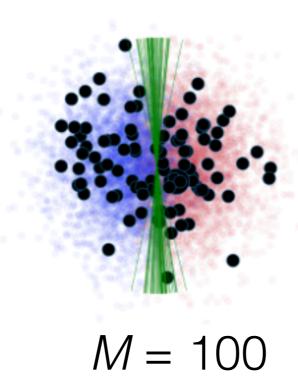
- Might miss important data
- Noisy estimates

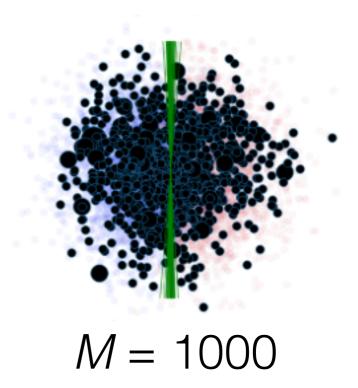




- Might miss important data
- Noisy estimates





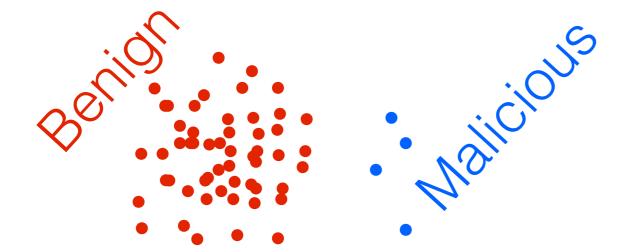


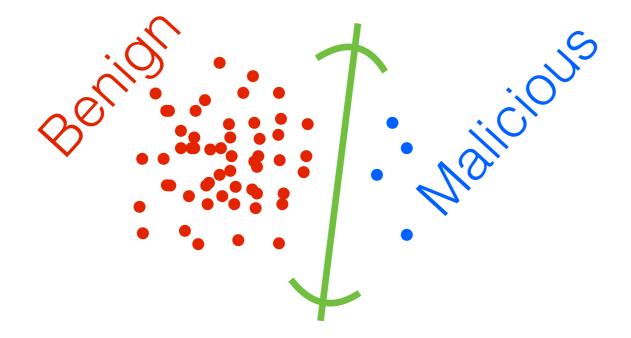
#### Roadmap

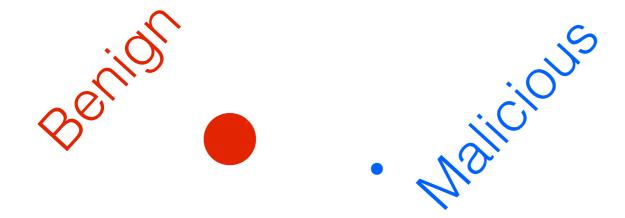
- The "core" of the data set
- Uniform data subsampling isn't enough
- Importance sampling for "coresets"
- Optimization for "coresets"
- Approximate sufficient statistics

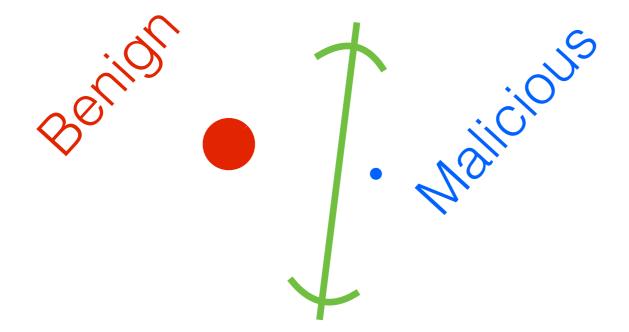
#### Roadmap

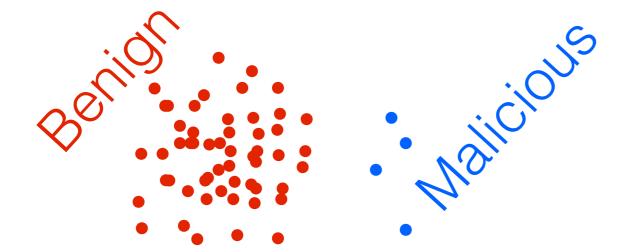
- The "core" of the data set
- Uniform data subsampling isn't enough
- Importance sampling for "coresets"
- Optimization for "coresets"
- Approximate sufficient statistics

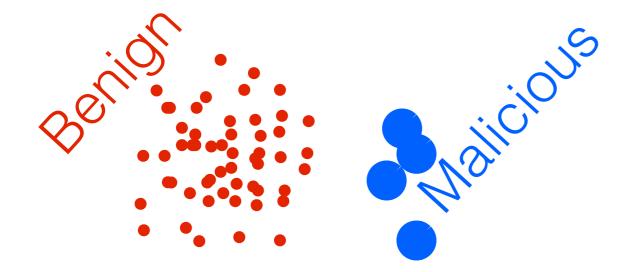


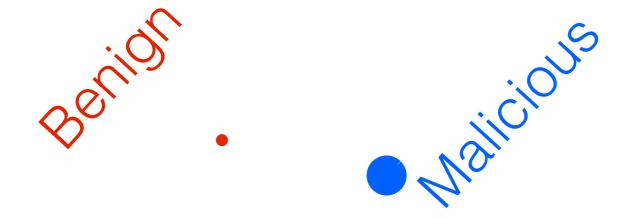


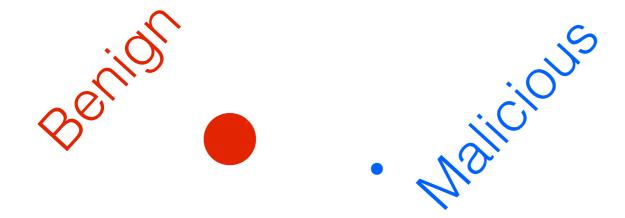


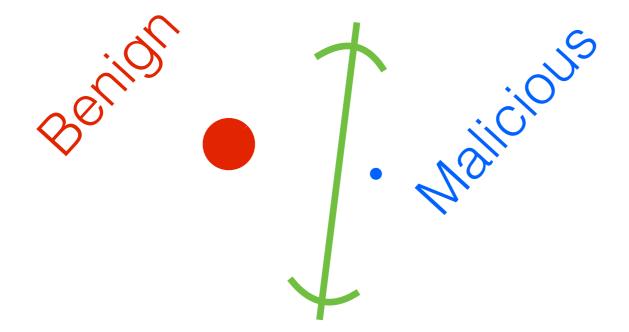


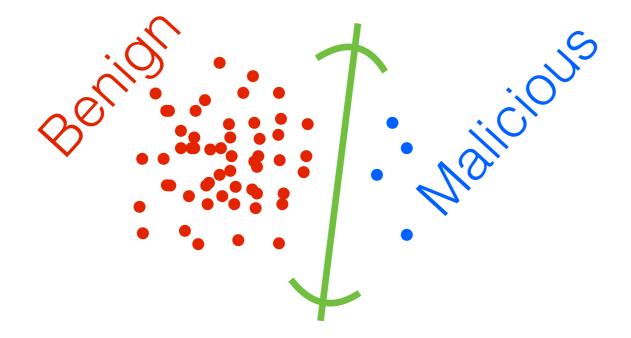


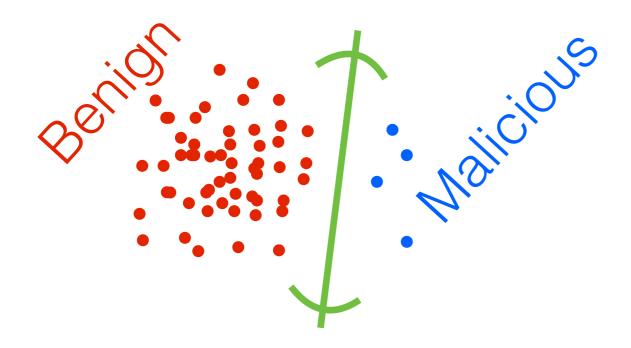




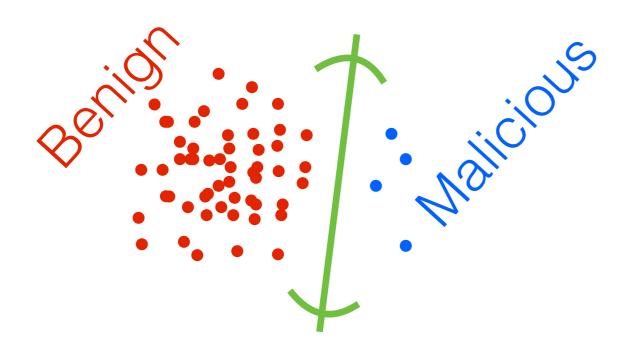






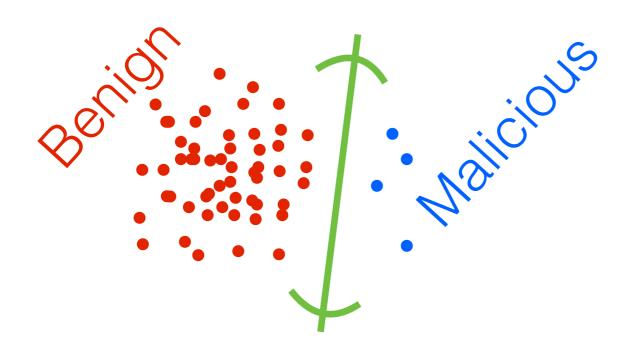


$$\sigma_n \propto \|\mathcal{L}_n\|$$



$$\sigma := \sum_{n=1}^{N} \|\mathcal{L}_n\|$$

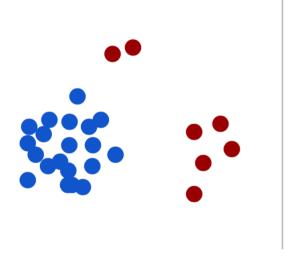
$$\sigma_n := \|\mathcal{L}_n\|/\sigma$$

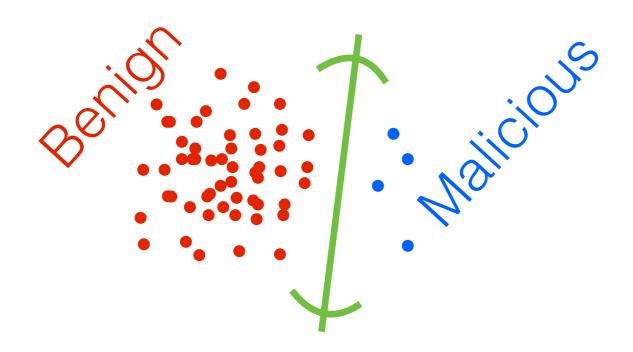


$$\sigma := \sum_{n=1}^{N} \|\mathcal{L}_n\|$$

$$\sigma_n := \|\mathcal{L}_n\|/\sigma$$

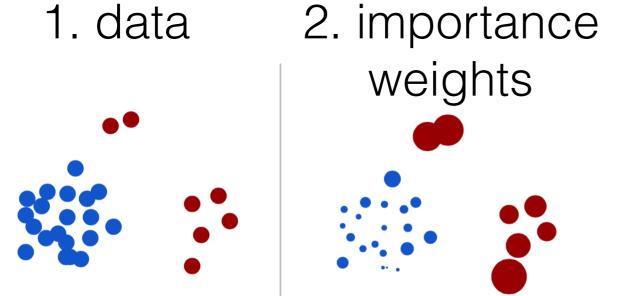
1. data

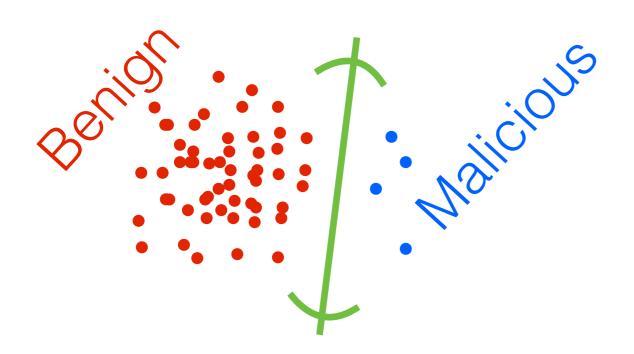




$$\sigma := \sum_{n=1}^{N} \|\mathcal{L}_n\|$$

$$\sigma_n := \|\mathcal{L}_n\|/\sigma$$

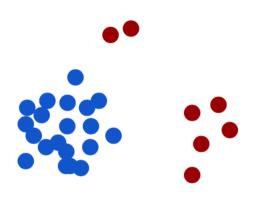




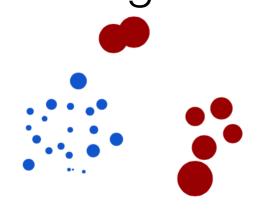
$$\sigma := \sum_{n=1}^{N} \|\mathcal{L}_n\|$$

$$\sigma_n := \|\mathcal{L}_n\|/\sigma$$

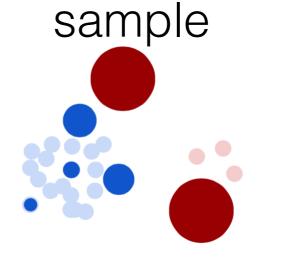
1. data

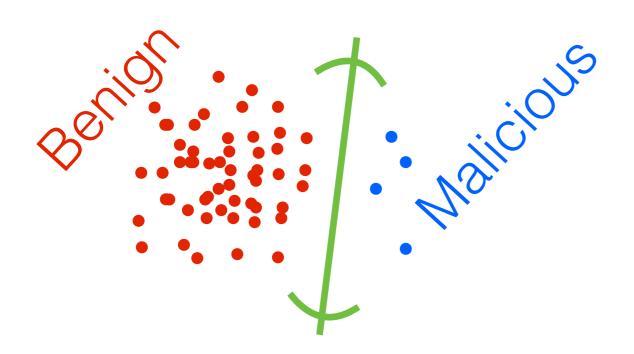


2. importance weights



3. importance

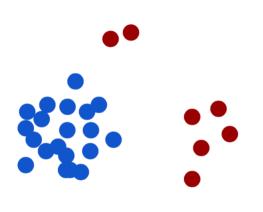




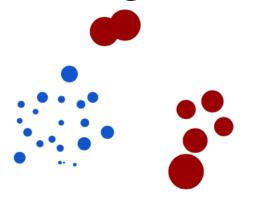
$$\sigma := \sum_{n=1}^{N} \|\mathcal{L}_n\|$$

$$\sigma_n := \|\mathcal{L}_n\|/\sigma$$

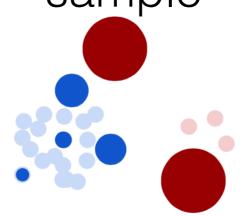
1. data



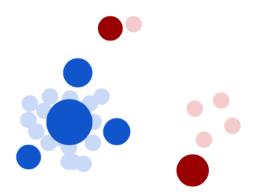
2. importance weights



3. importance sample



4. invert weights



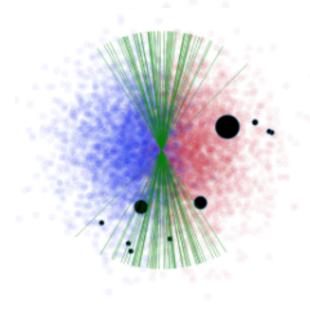
**Thm (CB)**.  $\delta \in (0,1)$ . With probability  $\geq 1 - \delta$ , after M iterations,

$$\|\mathcal{L}(w) - \mathcal{L}\| \le \frac{\sigma \bar{\eta}}{\sqrt{M}} \left( 1 + \sqrt{2 \log \frac{1}{\delta}} \right)$$

**Thm (CB)**.  $\delta \in (0,1)$ . With probability  $\geq 1 - \delta$ , after M iterations,

$$\|\mathcal{L}(w) - \mathcal{L}\| \le \frac{\sigma \bar{\eta}}{\sqrt{M}} \left( 1 + \sqrt{2 \log \frac{1}{\delta}} \right)$$

Still noisy estimates

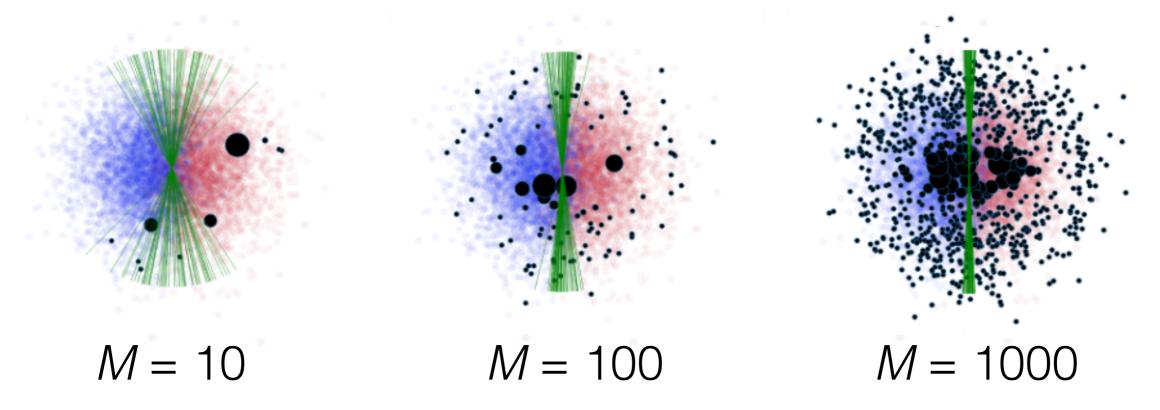


$$M = 10$$

**Thm (CB)**.  $\delta \in (0,1)$ . With probability  $\geq 1 - \delta$ , after M iterations,

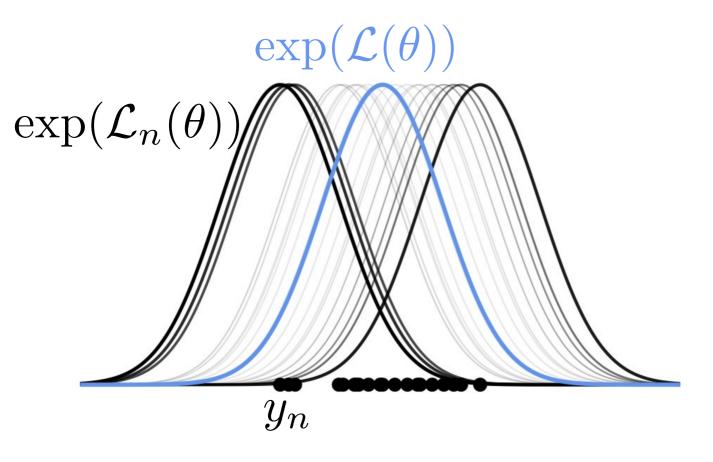
$$\|\mathcal{L}(w) - \mathcal{L}\| \le \frac{\sigma \bar{\eta}}{\sqrt{M}} \left( 1 + \sqrt{2 \log \frac{1}{\delta}} \right)$$

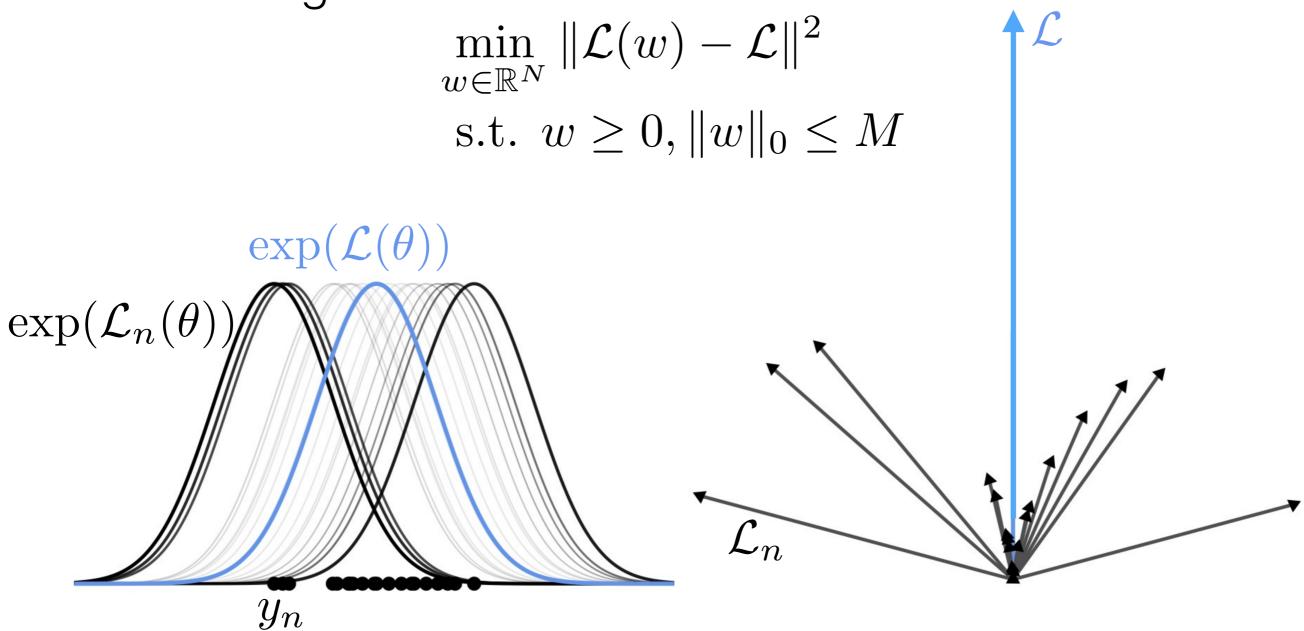
Still noisy estimates



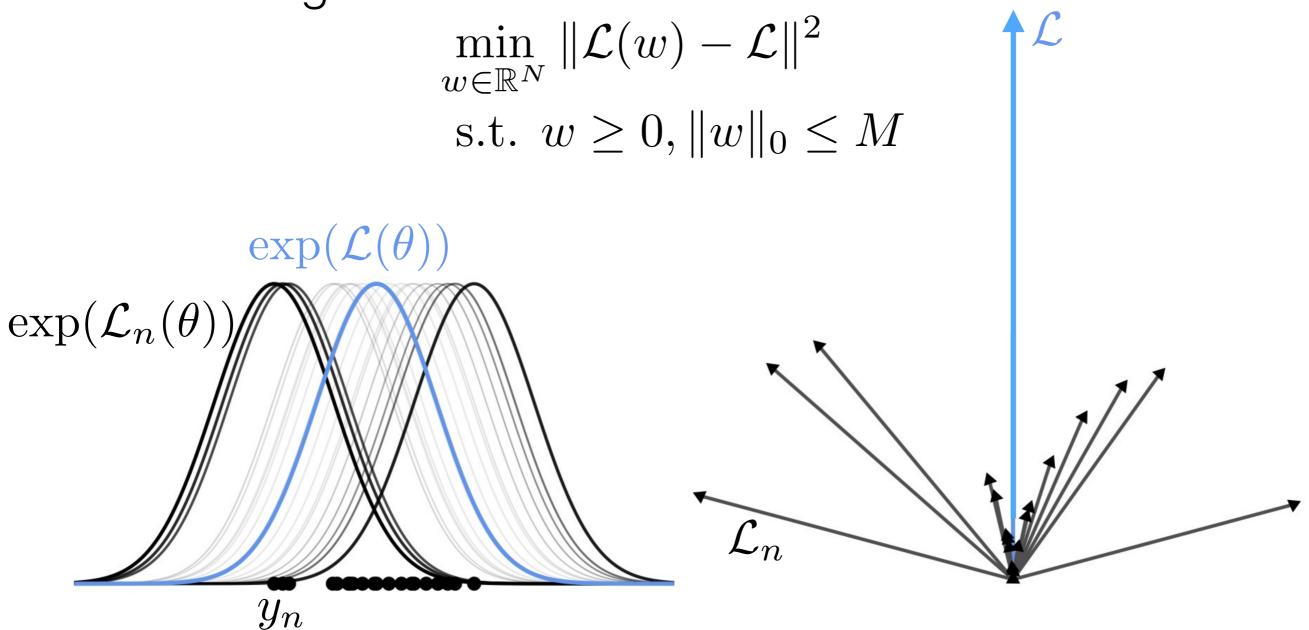
$$\min_{w \in \mathbb{R}^N} \|\mathcal{L}(w) - \mathcal{L}\|^2$$
  
s.t.  $w \ge 0, \|w\|_0 \le M$ 

$$\min_{w \in \mathbb{R}^N} \|\mathcal{L}(w) - \mathcal{L}\|^2$$
  
s.t.  $w \ge 0, \|w\|_0 \le M$ 

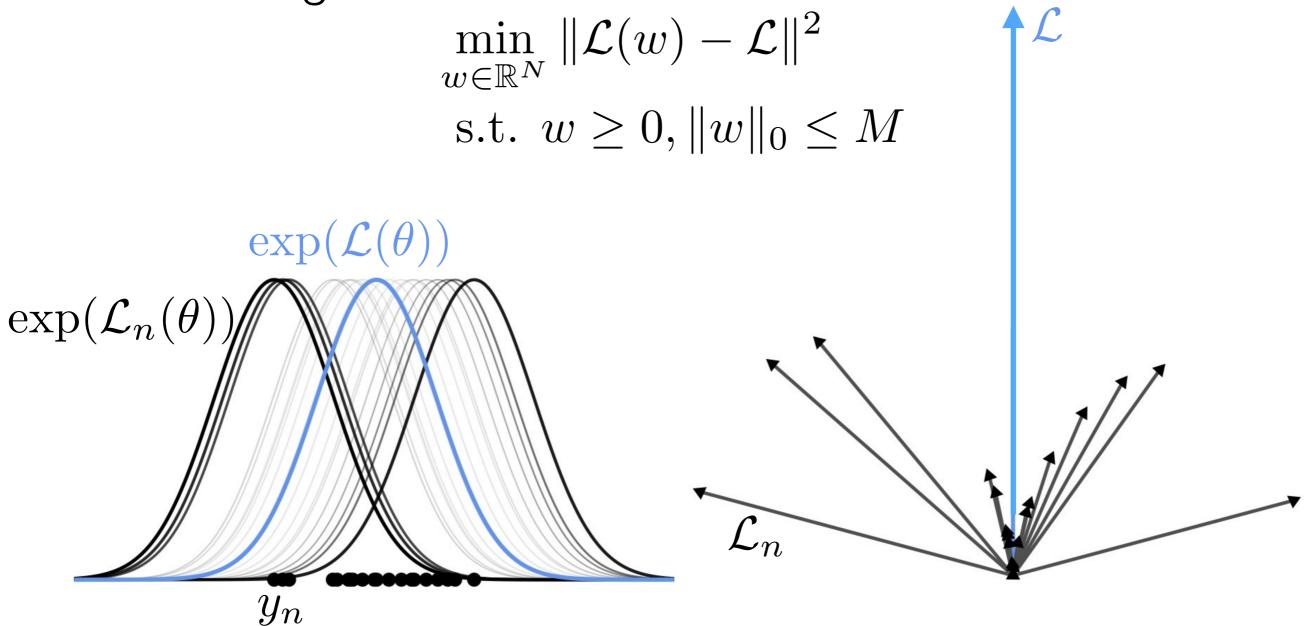




Want a good coreset:



need to consider (residual) error direction



- need to consider (residual) error direction
- sparse optimization

### Roadmap

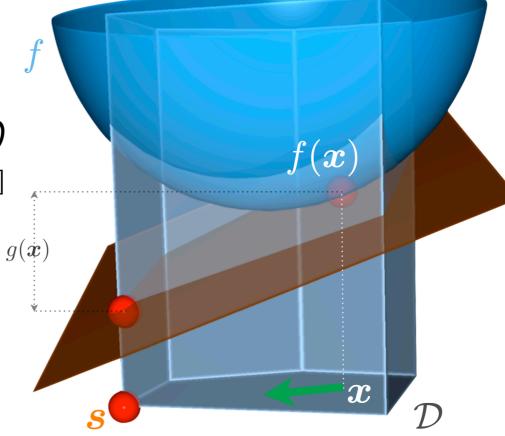
- The "core" of the data set
- Uniform data subsampling isn't enough
- Importance sampling for "coresets"
- Optimization for "coresets"
- Approximate sufficient statistics

### Roadmap

- The "core" of the data set
- Uniform data subsampling isn't enough
- Importance sampling for "coresets"
- Optimization for "coresets"
- Approximate sufficient statistics

Convex optimization on a polytope D

[Frank, Wolfe 1956]



[Jaggi 2013]

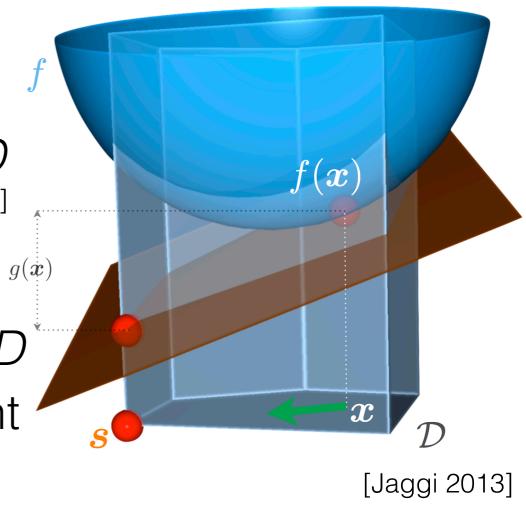
Convex optimization on a polytope D

Repeat:

1. Find gradient

2. Find argmin point on plane in D

3. Do line search between current point and argmin point



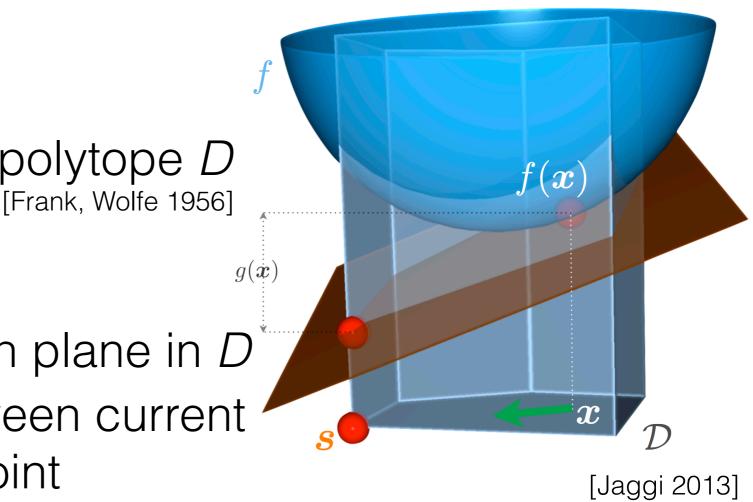
Convex optimization on a polytope D

Repeat:

1. Find gradient

2. Find argmin point on plane in D

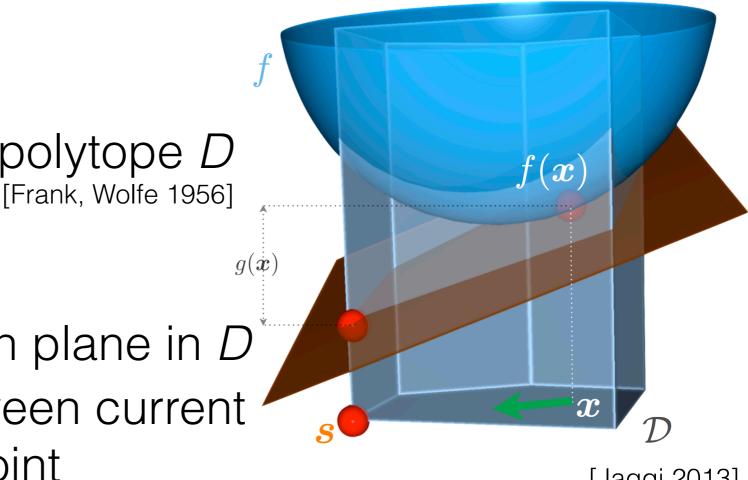
3. Do line search between current point and argmin point



Convex combination of M vertices after M -1 steps

Convex optimization on a polytope D

- Repeat:
  - 1. Find gradient
  - 2. Find argmin point on plane in D
  - 3. Do line search between current point and argmin point

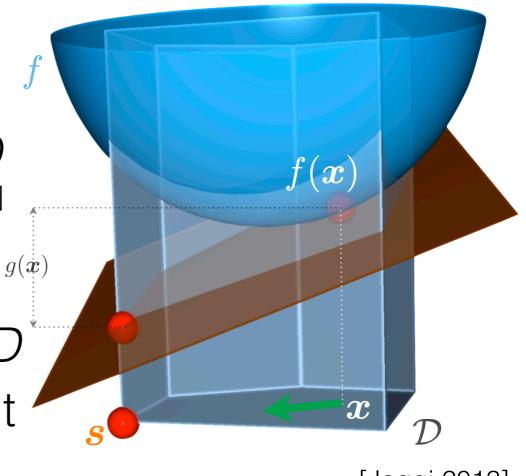


[Jaggi 2013]

- Convex combination of M vertices after M-1 steps
- Our problem:  $\min_{w \in \mathbb{R}^N} \|\mathcal{L}(w) \mathcal{L}\|$

Convex optimization on a polytope D
[Frank, Wolfe 1956]

- Repeat:
  - 1. Find gradient
  - 2. Find argmin point on plane in D
  - 3. Do line search between current point and argmin point

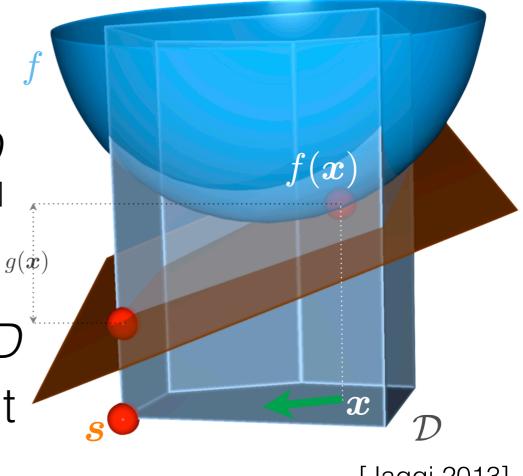


[Jaggi 2013]

- Convex combination of M vertices after M-1 steps
- Our problem:  $\min_{w \in \mathbb{R}^N} \|\mathcal{L}(w) \mathcal{L}\|^2$

Convex optimization on a polytope D

- Repeat:
  - 1. Find gradient
  - 2. Find argmin point on plane in D
  - 3. Do line search between current point and argmin point



[Jaggi 2013]

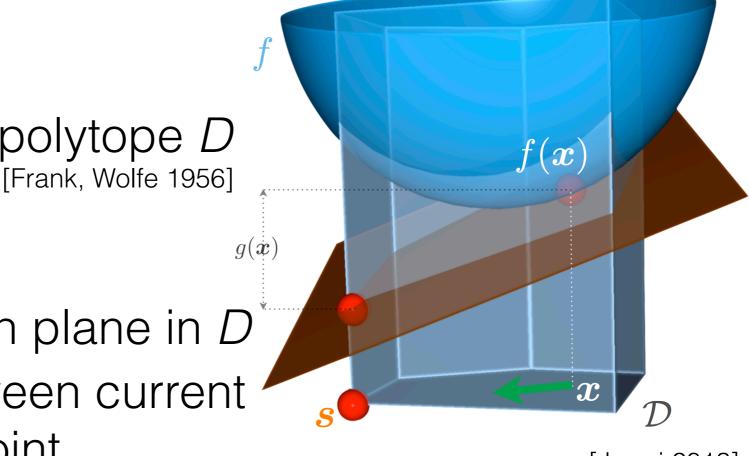
- Convex combination of M vertices after M -1 steps
- Our problem:  $\min_{w \in \mathbb{R}^N} \|\mathcal{L}(w) \mathcal{L}\|^2$

s.t. 
$$w \ge 0, ||w||_0 \le M$$

[Frank, Wolfe 1956]

Convex optimization on a polytope D

- Repeat:
  - 1. Find gradient
  - 2. Find argmin point on plane in D
  - 3. Do line search between current point and argmin point

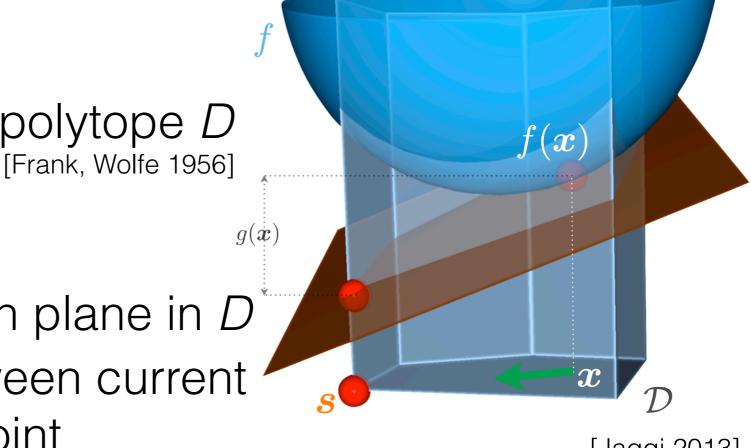


[Jaggi 2013]

- Convex combination of M vertices after M-1 steps
- Our problem:  $\min_{w \in \mathbb{R}^N} \|\mathcal{L}(w) \mathcal{L}\|^2$   $\Delta^{N-1} := \left\{ w \in \mathbb{R}^N : \sum_{n=1}^N \sigma_n w_n = \sigma, w \ge 0 \right\}$

Convex optimization on a polytope D

- Repeat:
  - 1. Find gradient
  - 2. Find argmin point on plane in D
  - 3. Do line search between current point and argmin point



[Jaggi 2013]

- Convex combination of M vertices after M-1 steps
- Our problem:  $\min_{w \in \mathbb{R}^N} \|\mathcal{L}(w) \mathcal{L}\|^2$   $\Delta^{N-1} := \left\{ w \in \mathbb{R}^N : \sum_{n=1}^N \sigma_n w_n = \sigma, w \ge 0 \right\}$

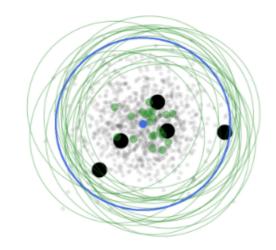
**Thm (CB)**. After *M* iterations,

$$\|\mathcal{L}(w) - \mathcal{L}\| \le \frac{c}{\sqrt{\alpha^{2M} + c'M}}$$

# Gaussian model (simulated)

• 1K pts; norms, inference: closed-form

Uniform subsampling

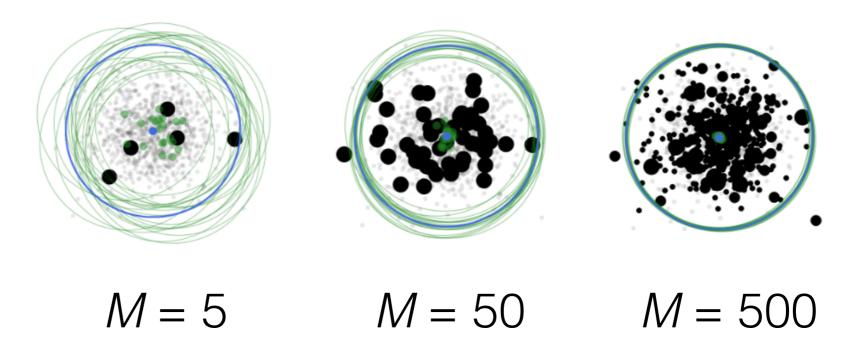


$$M = 5$$

# Gaussian model (simulated)

• 1K pts; norms, inference: closed-form

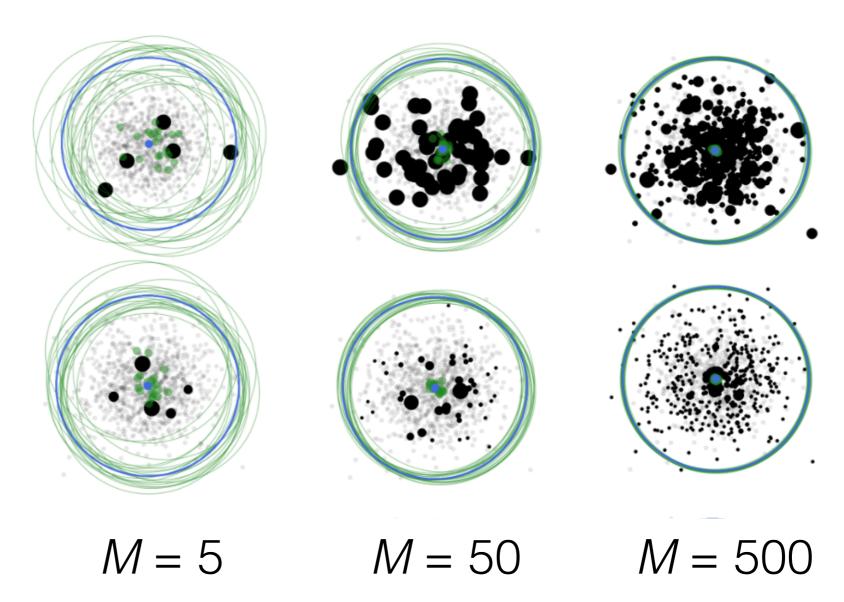
Uniform subsampling



# Gaussian model (simulated)

• 1K pts; norms, inference: closed-form

Uniform subsampling

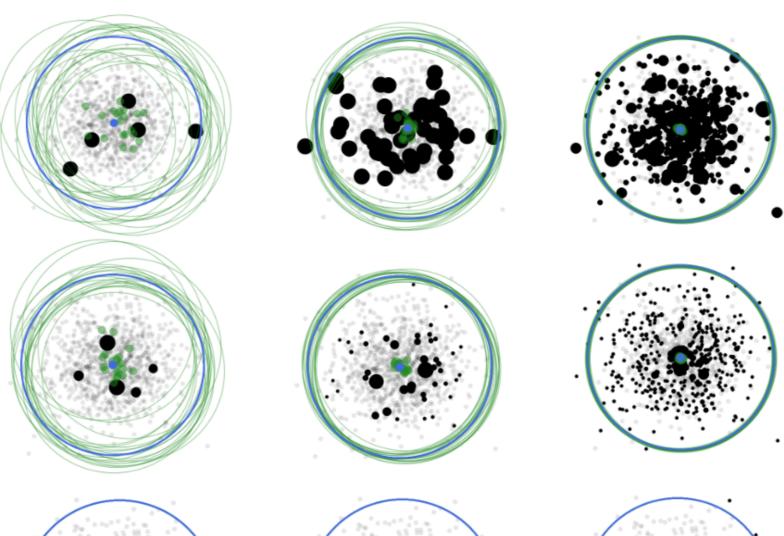


Importance sampling

# Gaussian model (simulated)

• 1K pts; norms, inference: closed-form

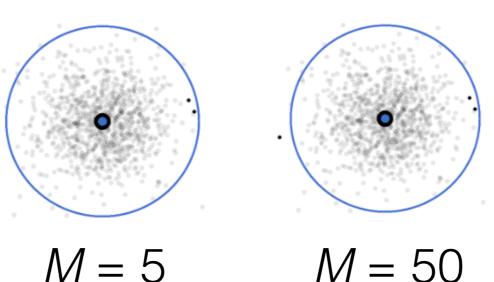
Uniform subsampling



M = 500

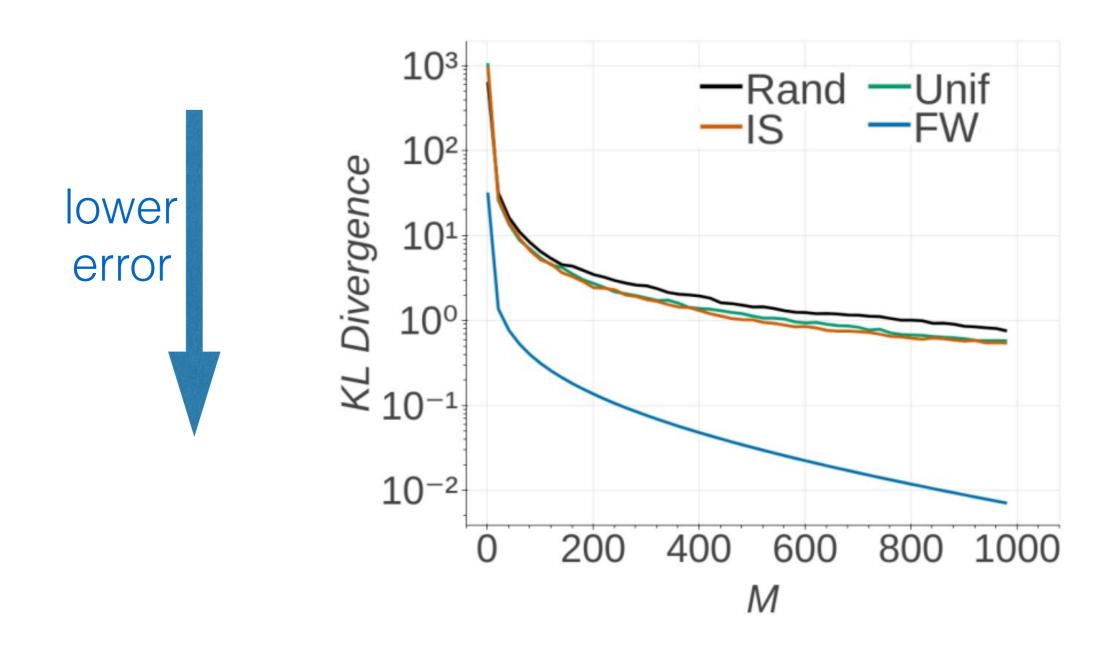
Importance sampling

Frank-Wolfe



# Gaussian model (simulated)

• 1K pts; norms, inference: closed-form



# Logistic regression (simulated)

• 10K pts; general inference

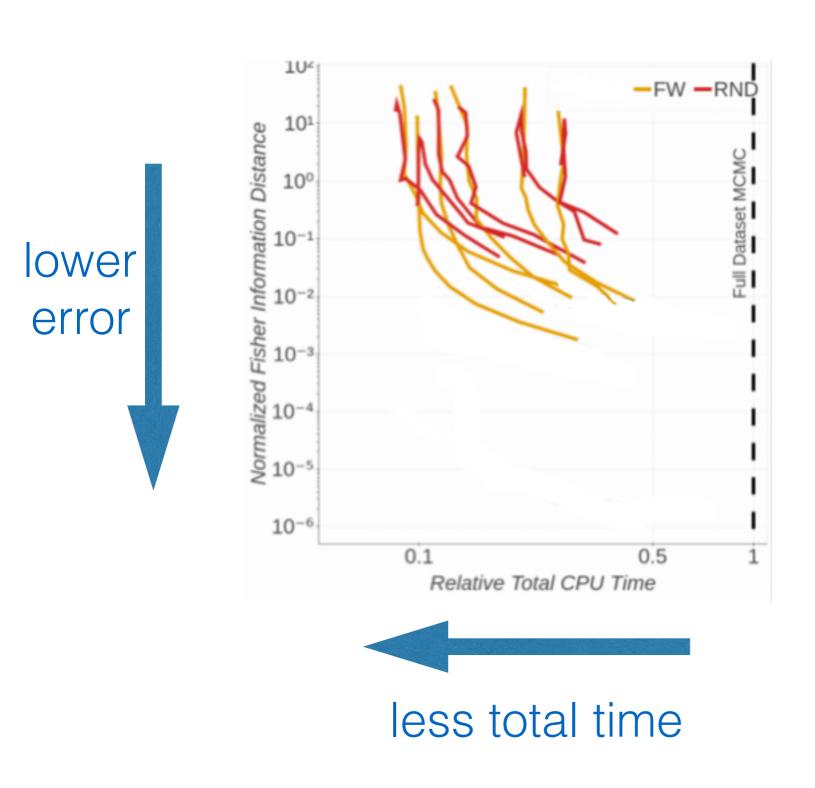
Uniform subsampling Importance sampling Frank-Wolfe M = 100M = 10M = 1000

# Poisson regression (simulated)

• 10K pts; general inference

Uniform subsampling Importance sampling Frank-Wolfe M = 100M = 10M = 1000

# Real data experiments

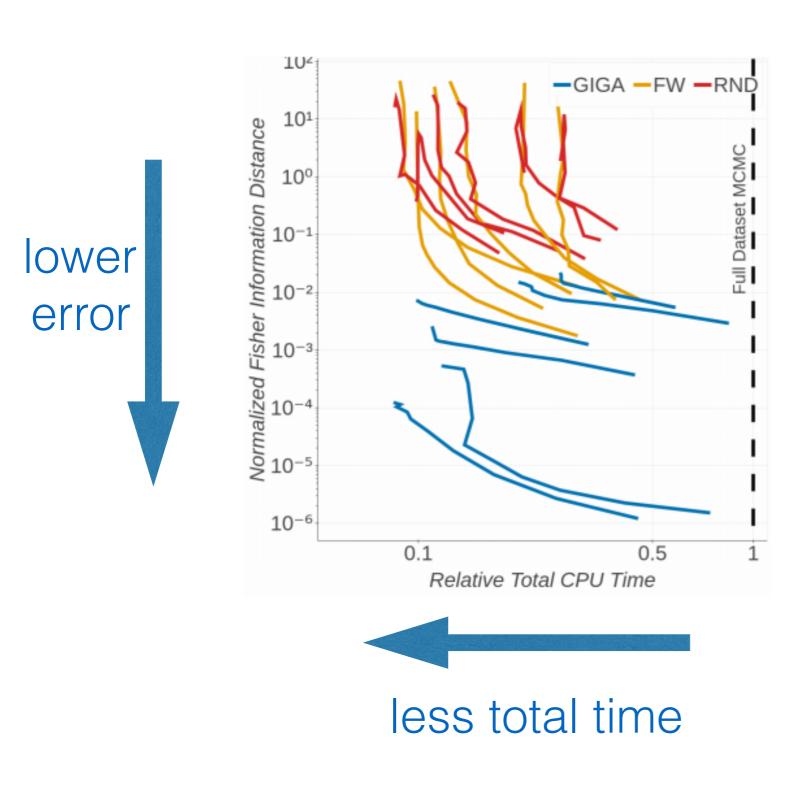


Uniform
subsampling
Frank Wolfe
coresets

Data sets include:

- Phishing
- Chemical reactivity
- Bicycle trips
- Airport delays

# Real data experiments





#### Data sets include:

- Phishing
- Chemical reactivity
- Bicycle trips
- Airport delays

# Roadmap

- The "core" of the data set
- Uniform data subsampling isn't enough
- Importance sampling for "coresets"
- Optimization for "coresets"
- Approximate sufficient statistics

# Roadmap

- The "core" of the data set
- Uniform data subsampling isn't enough
- Importance sampling for "coresets"
- Optimization for "coresets"
- Approximate sufficient statistics

Exponential family likelihood

Exponential family likelihood

$$p(y_{1:N}|x_{1:N},\theta) = \prod_{n=1}^{N} \exp \left[T(y_n,x_n) \cdot \eta(\theta)\right]$$

### **Sufficient statistics**

$$[T(y_n,x_n)\cdot\eta(\theta)]$$

Exponential family likelihood

### **Sufficient statistics**

$$p(y_{1:N}|x_{1:N},\theta) = \prod_{n=1}^{N} \exp \left[T(y_n,x_n) \cdot \eta(\theta)\right]$$

$$= \exp \left[ \left\{ \sum_{n=1}^{N} T(y_n, x_n) \right\} \cdot \eta(\theta) \right]$$

Exponential family likelihood

### **Sufficient statistics**

$$p(y_{1:N}|x_{1:N},\theta) = \prod_{n=1}^{N} \exp \left[T(y_n,x_n) \cdot \eta(\theta)\right]$$

$$= \exp \left[ \left\{ \sum_{n=1}^{N} T(y_n, x_n) \right\} \cdot \eta(\theta) \right]$$

 Scalable, single-pass, streaming, distributed, complementary to MCMC

Exponential family likelihood

# $p(y_{1:N}|x_{1:N}, \theta) = \prod_{1}^{N} \exp \left[ T(y_n, x_n) \cdot \eta(\theta) \right]$

$$= \exp \left[ \left\{ \sum_{n=1}^{N} T(y_n, x_n) \right\} \cdot \eta(\theta) \right]$$

- Scalable, single-pass, streaming, distributed, complementary to MCMC
- But: Often no simple sufficient statistics

Exponential family likelihood

### Sufficient statistics

$$p(y_{1:N}|x_{1:N},\theta) = \prod_{n=1}^{N} \exp \left[T(y_n,x_n) \cdot \eta(\theta)\right]$$

$$= \exp \left[ \left\{ \sum_{n=1}^{N} T(y_n, x_n) \right\} \cdot \eta(\theta) \right]$$

- Scalable, single-pass, streaming, distributed, complementary to MCMC
- But: Often no simple sufficient statistics
  - E.g. Bayesian logistic regression; GLMs; "deeper" models

• Likelihood 
$$p(y_{1:N}|x_{1:N}, \theta) = \prod_{n=1}^{N} \frac{1}{1 + \exp(-y_n x_n \cdot \theta)}$$

Exponential family likelihood

# $p(y_{1:N}|x_{1:N}, \theta) = \prod_{n=1}^{N} \exp \left[ T(y_n, x_n) \cdot \eta(\theta) \right]$

$$= \exp \left[ \left\{ \sum_{n=1}^{N} T(y_n, x_n) \right\} \cdot \eta(\theta) \right]$$

- Scalable, single-pass, streaming, distributed, complementary to MCMC
- But: Often no simple sufficient statistics
  - E.g. Bayesian logistic regression; GLMs; "deeper" models
    - Likelihood  $p(y_{1:N}|x_{1:N}, \theta) = \prod_{n=1}^{1} \frac{1}{1 + \exp(-y_n x_n \cdot \theta)}$

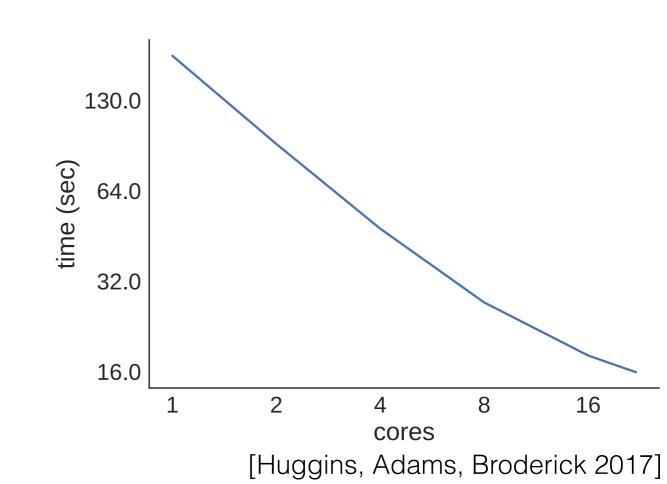
Our proposal: (polynomial) approximate sufficient statistics

Criteo Labs > Algorithms > Criteo Releases its New Dataset

### Criteo Releases its New Dataset

By: CriteoLabs / 31 Mar 2015

- 6M data points, 1000 features
- Streaming, distributed;
   minimal communication
- 22 cores, 16 sec
- Finite-data guarantees on Wasserstein distance to exact posterior



 Data summarization for scalable, automated approximate Bayes algorithms with error bounds on quality for finite data

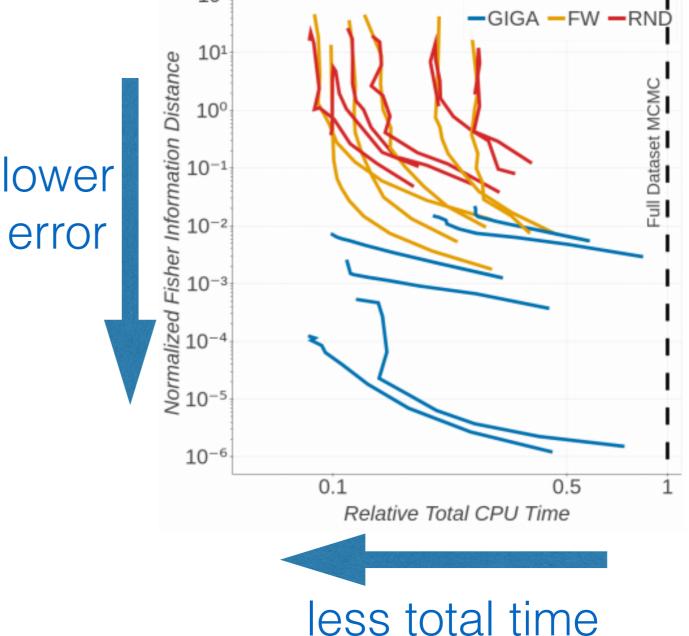
- Data summarization for scalable, automated approximate Bayes algorithms with error bounds on quality for finite data
  - Coresets
  - Approx. suff. statistics

Data summarization for scalable, automated approximate Bayes algorithms with error bounds on

quality for finite data

Coresets

Approx. suff. statistics



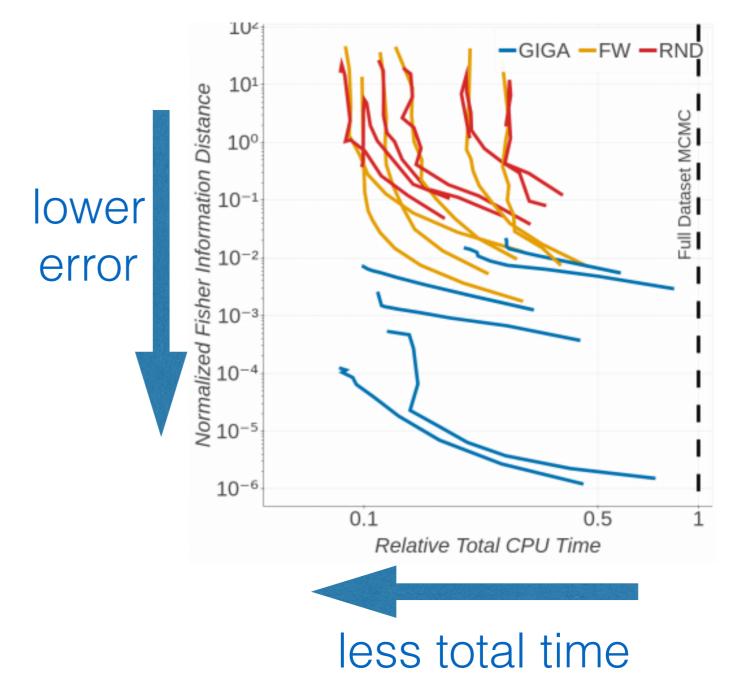
 Data summarization for scalable, automated approximate Bayes algorithms with error bounds on

quality for finite data

Coresets

Approx. suff. statistics

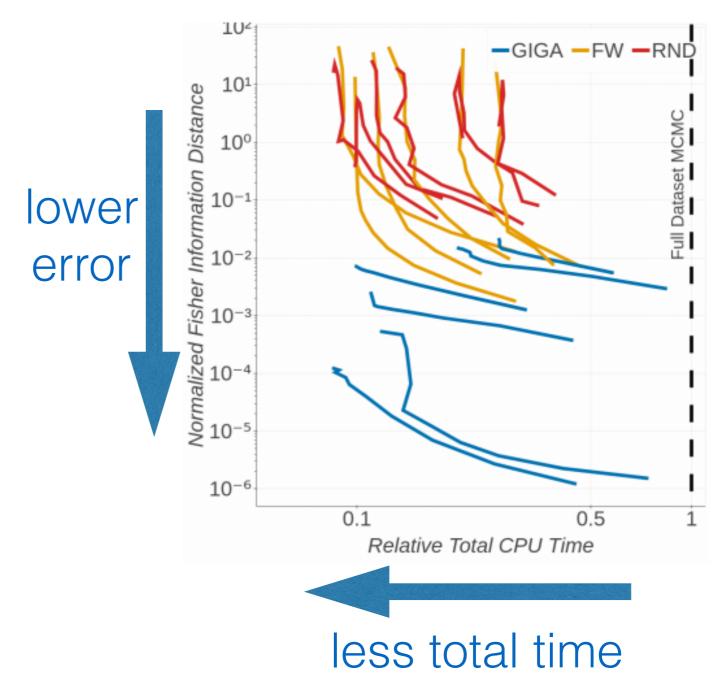
 More accurate with more computation investment



 Data summarization for scalable, automated approximate Bayes algorithms with error bounds on

quality for finite data

- Coresets
- Approx. suff. statistics
- More accurate with more computation investment
- A start
  - Lots of potential improvements/ directions



### References

# T Campbell and T Broderick. Automated scalable Bayesian inference via Hilbert coresets. *Journal of Machine Learning Research*, 2019.

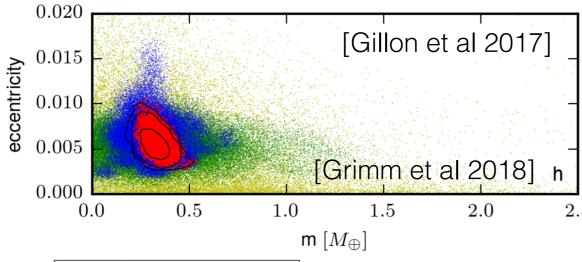
T Campbell and T Broderick. Bayesian coreset construction via greedy iterative geodesic ascent. *ICML* 2018.

JH Huggins, T Campbell, M Kasprzak, and T Broderick. Practical bounds on the error of Bayesian posterior approximations: A nonasymptotic approach. ArXiv:1809.09505.

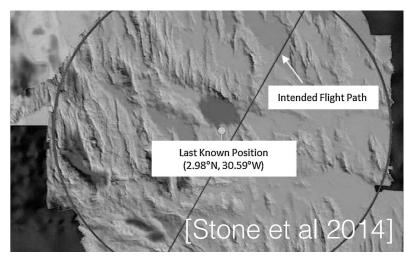
JH Huggins, T Campbell, and T Broderick. Coresets for scalable Bayesian logistic regression. *NeurIPS* 2016.

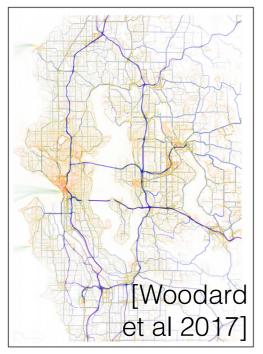
JH Huggins, RP Adams, and T Broderick. PASS-GLM: Polynomial approximate sufficient statistics for scalable Bayesian GLM inference. *NeurIPS* 2017.

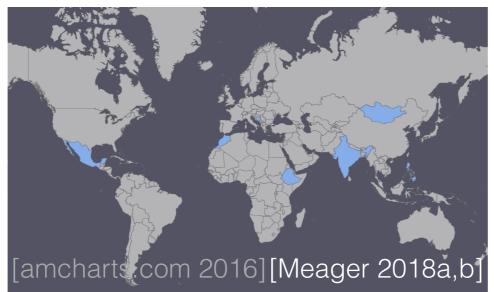
R Agrawal, T Campbell, JH Huggins, and T Broderick. Data-dependent compression of random features for large-scale kernel approximation. *AISTATS* 2019.





















Challenge: fast (compute, user), reliable inference



- Challenge: fast (compute, user), reliable inference
- Today: variational Bayes and beyond



- Challenge: fast (compute, user), reliable inference
- Today: variational Bayes and beyond
- Fundamental questions
  - What is achievable in speed and accuracy?

# References (1/6)

R Bardenet, A Doucet, and C Holmes. "On Markov chain Monte Carlo methods for tall data." Journal of Machine Learning Research 18.1 (2017): 1515-1557.

AG Baydin, BA Pearlmutter, AA Radul, and JM Siskind. "Automatic differentiation in machine learning: a survey." *Journal of Machine Learning Research*, 2018.

DM Blei, A Kucukelbir, and JD McAuliffe. "Variational inference: A review for statisticians." Journal of the American Statistical Association 112.518 (2017): 859-877.

T Broderick, N Boyd, A Wibisono, AC Wilson, and MI Jordan. Streaming variational Bayes. *NeurIPS* 2013.

CM Bishop. Pattern Recognition and Machine Learning. Springer-Verlag New York, 2006.

T Campbell and T Broderick. Automated scalable Bayesian inference via Hilbert coresets. Journal of Machine Learning Research, 2019.

T Campbell and T Broderick. Bayesian Coreset Construction via Greedy Iterative Geodesic Ascent. *ICML* 2018.

RJ Giordano, T Broderick, and MI Jordan. "Linear response methods for accurate covariance estimates from mean field variational Bayes." *NeurIPS* 2015.

R Giordano, T Broderick, R Meager, J Huggins, and MI Jordan. "Fast robustness quantification with variational Bayes." *ICML 2016 Workshop on #Data4Good: Machine Learning in Social Good Applications*, 2016.

RJ Giordano, T Broderick, and MI Jordan. Covariances, robustness, and variational Bayes. *Journal of Machine Learning Research*, 2018.

# References (2/6)

J Gorham and L Mackey. "Measuring sample quality with Stein's method." NeurIPS 2015.

J Gorham, and L Mackey. "Measuring sample quality with kernels." ArXiv:1703.01717 (2017).

PD Hoff. A first course in Bayesian statistical methods. Springer Science & Business Media, 2009.

MD Hoffman, DM Blei, C Wang, and J Paisley. "Stochastic variational inference." *The Journal of Machine Learning Research* 14.1 (2013): 1303-1347.

JH Huggins, T Campbell, and T Broderick. Coresets for scalable Bayesian logistic regression. NeurIPS 2016.

JH Huggins, RP Adams, and T Broderick. PASS-GLM: Polynomial approximate sufficient statistics for scalable Bayesian GLM inference. *NeurIPS* 2017.

J Huggins, T Campbell, M Kasprzak, T Broderick. Practical bounds on the error of Bayesian posterior approximations: A nonasymptotic approach, 2018. ArXiv:1809.09505.

A Kucukelbir, R Ranganath, A Gelman, and D Blei. Automatic variational inference in Stan. NeurIPS 2015.

A Kucukelbir, D Tran, R Ranganath, A Gelman, and DM Blei. Automatic differentiation variational inference. *Journal of Machine Learning Research* 18.1 (2017): 430-474.

DJC MacKay. *Information Theory, Inference, and Learning Algorithms*. Cambridge University Press, 2003.

Stan (open source software). http://mc-stan.org/ Accessed: 2018.

# References (3/6)

S Talts, M Betancourt, D Simpson, A Vehtari, and A Gelman. Validating Bayesian Inference Algorithms with Simulation-Based Calibration. ArXiv:1804.06788 (2018).

RE Turner and M Sahani. Two problems with variational expectation maximisation for timeseries models. In D Barber, AT Cemgil, and S Chiappa, editors, *Bayesian Time Series Models*, 2011.

Y Yao, A Vehtari, D Simpson, and A Gelman. Yes, but Did It Work?: Evaluating Variational Inference. *ICML* 2018.

# Application References (4/6)

Abbott, Benjamin P., et al. "Observation of gravitational waves from a binary black hole merger." *Physical Review Letters* 116.6 (2016): 061102.

Abbott, Benjamin P., et al. "The rate of binary black hole mergers inferred from advanced LIGO observations surrounding GW150914." *The Astrophysical Journal Letters* 833.1 (2016): L1.

Airoldi, Edoardo M., David M. Blei, Stephen E. Fienberg, and Eric P. Xing. "Mixed membership stochastic blockmodels." *Journal of Machine Learning Research* 9.Sep (2008): 1981-2014.

Blei, David M., Andrew Y. Ng, and Michael I. Jordan. "Latent Dirichlet allocation." *Journal of Machine Learning Research* 3.Jan (2003): 993-1022.

Chati, Yashovardhan Sushil, and Hamsa Balakrishnan. "A Gaussian process regression approach to model aircraft engine fuel flow rate." *Cyber-Physical Systems (ICCPS), 2017 ACM/IEEE 8th International Conference on.* IEEE, 2017.

Gershman, Samuel J., David M. Blei, Kenneth A. Norman, and Per B. Sederberg. "Decomposing spatiotemporal brain patterns into topographic latent sources." Neurolmage 98 (2014): 91-102.

Gillon, Michaël, et al. "Seven temperate terrestrial planets around the nearby ultracool dwarf star TRAPPIST-1." *Nature* 542.7642 (2017): 456.

Grimm, Simon L., et al. "The nature of the TRAPPIST-1 exoplanets." *Astronomy & Astrophysics* 613 (2018): A68.

# Application References (5/6)

Grogan Jr, William L., and Willis W. Wirth. "A new American genus of predaceous midges related to Palpomyia and Bezzia (Diptera: Ceratopogonidae). Un nuevo género Americano de purrujas depredadoras relacionadas con Palpomyia y Bezzia (Diptera: Ceratopogonidae)." *Proceedings of the Biological Society of Washington*. 94.4 (1981): 1279-1305.

Kuikka, Sakari, Jarno Vanhatalo, Henni Pulkkinen, Samu Mäntyniemi, and Jukka Corander. "Experiences in Bayesian inference in Baltic salmon management." *Statistical Science* 29.1 (2014): 42-49.

Meager, Rachael. "Understanding the impact of microcredit expansions: A Bayesian hierarchical analysis of 7 randomized experiments." *AEJ: Applied*, to appear, 2018a.

Meager, Rachael. "Aggregating Distributional Treatment Effects: A Bayesian Hierarchical Analysis of the Microcredit Literature." Working paper, 2018b.

Stegle, Oliver, Leopold Parts, Richard Durbin, and John Winn. "A Bayesian framework to account for complex non-genetic factors in gene expression levels greatly increases power in eQTL studies." PLoS computational biology 6.5 (2010): e1000770.

Stone, Lawrence D., Colleen M. Keller, Thomas M. Kratzke, and Johan P. Strumpfer. "Search for the wreckage of Air France Flight AF 447." *Statistical Science* (2014): 69-80.

Woodard, Dawn, Galina Nogin, Paul Koch, David Racz, Moises Goldszmidt, and Eric Horvitz. "Predicting travel time reliability using mobile phone GPS data." *Transportation Research Part C: Emerging Technologies* 75 (2017): 30-44.

Xing, Eric P., Wei Wu, Michael I. Jordan, and Richard M. Karp. "LOGOS: a modular Bayesian model for de novo motif detection." Journal of Bioinformatics and Computational Biology 2.01 (2004): 127-154.

# Additional image references (6/6)

amCharts. Visited Countries Map. https://www.amcharts.com/visited\_countries/ Accessed: 2016.

Baltic Salmon Fund. https://www.en.balticsalmonfund.org/about\_us Accessed: 2018.

ESO/L. Calçada/M. Kornmesser. 16 October 2017, 16:00:00. Obtained from: https://commons.wikimedia.org/wiki/

File:Artist%E2%80%99s\_impression\_of\_merging\_neutron\_stars.jpg || Source: https://www.eso.org/public/images/eso1733a/ (Creative Commons Attribution 4.0 International License)

J. Herzog. 3 June 2016, 17:17:30. Obtained from: https://commons.wikimedia.org/wiki/File:Airbus\_A350-941\_F-WWCF\_MSN002\_ILA\_Berlin\_2016\_17.jpg (Creative Commons Attribution 4.0 International License)

E. Xing. 2003. Slides "LOGOS: a modular Bayesian model for de novo motif detection." Obtained from: https://www.cs.cmu.edu/~epxing/papers/Old\_papers/slide\_CSB03/CSB1.pdf Accessed: 2018.